Directed Mixed Graphs for Latent Variables

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Stats 212 Graphical Models Lecture Notes

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- 2 Factorizations on ADMGs
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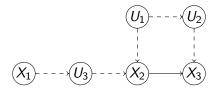
Latent projection of a DAG (Tian and Pearl 2002b):

Given a DAG with latent variables $\mathcal{G}(V \cup L)$, where V is observed and L latent, the *latent projection* $\mathcal{G}(V)$ is constructed as follows:

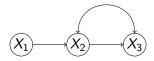
- **I** $\mathcal{G}(V)$ contains an edge $a \to b$ if there is a directed path $a \to \cdots \to b$ in $\mathcal{G}(V \cup L)$ with all intermediate vertices in L.
- **2** $\mathcal{G}(V)$ contains an edge $a \leftrightarrow b$ if there is a collider-free path $a \leftarrow \cdots \rightarrow b$ with all intermediate vertices in *L*.

Note: Step 1 adds all directed edges $a \rightarrow b$ in $\mathcal{G}(V \cup L)$ to $\mathcal{G}(V)$.

DAG $\mathcal{G}(V \cup L)$, $V = \{X_1, X_2, X_3\}$ and $L = \{U_1, U_2, U_3\}$:



Latent projection $\mathcal{G}(V)$ is an acyclic directed mixed graph (ADMG):



Definitions. Let $\mathcal{G} = (V, E)$ be a directed mixed graph, i.e. a graph with two types of edges: directed (\rightarrow) or bidirected (\leftrightarrow) .

- A path is a sequence of distinct adjacent edges, of any type or orientation, between distinct vertices. directed path: a → · · · → b. bidirected path: a ↔ · · · ↔ b.
- If $a \rightarrow b$, then a is a parent of b and b is a child of a.
- If there is a directed path from a to d or a = d, we say a is an ancestor of d and d is a descendant of a. Accordingly define non-descendant.
- If $a \leftrightarrow b$, then *a* is a sibling of *b*.
- notation: pa_G(a), ch_G(a), an_G(a), de_G(a), nd_G(a), and sib_G(a).

- A *directed cycle* is a path of the form $v \rightarrow \cdots \rightarrow w$ along with an edge $w \rightarrow v$.
- An acyclic directed mixed graph (ADMG) is a mixed graph containing no directed cycles.
- A topological sort of an ADMG is defined in the same way as for a DAG: a → b implies a ≺ b.

m-separation:

- A vertex z is a collider on a path if → z ←, ↔ z ↔, → z ↔, or ↔ z ←; otherwise, z is a non-collider.
- *m*-connection: A path between *a* and *b* is *m*-connecting given C if (i) every non-collider on the path is not in C and (ii) every collider on the path is an ancestor of C (an(C) := ∪_{a∈C}an(a)).
- *m*-separation: If there is no path *m*-connecting *a* and *b* given
 C, then *a* and *b* are *m*-separated given *C*.
- If \mathcal{G} is a DAG, *m*-separation is identical to *d*-separation.

m-separation:

Proposition 1 (Richardson et al. (2023))

Let $\mathcal{G}(V \cup L)$ be a DAG and $\mathcal{G}(V)$ be its latent projection. For disjoint subsets $A, B, C \subset V$, A and B are d-separated given C in $\mathcal{G}(V \cup L)$ if and only if A and B are m-separated given C in $\mathcal{G}(V)$.

- On every path between a, b ∈ V in G(V ∪ L), colliders (resp. non-colliders) in V are also colliders (resp. non-colliders) on a path in G(V).
- ADMG G(V) captures all conditional independence constraints among the observed variables V in the DAG G(V ∪ L) with latent variables.

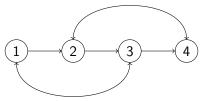
Districts in ADMG $\mathcal{G}(V)$:

- The *district* of vertex *v*, denoted dis_G(*v*), is the set of vertices that are connected to *v* by a bidirected path (including *v* itself).
- A district of G is a maximal bidirected-connected set of vertices.
- A district corresponds to a confounded component (c-component) (Tian and Pearl 2002b).
- Districts specify variable partitions that define terms in the factorization of $\mathbb{P}(V)$.

Denote districts by $\mathcal{D}(\mathcal{G}) = \{D : D \text{ is a district of } \mathcal{G}\}.$ Define $pa_{\mathcal{G}}(D) := (\cup_{a \in D} pa_{\mathcal{G}}(a)) \setminus D.$

Factorizations on ADMGs

District factorization:



Districts of G: $D_1 = \{1,3\}, D_2 = \{2,4\}.$ $pa_G(D_1) = \{2\},$ $pa_G(D_2) = \{1,3\}.$

Using $a \leftrightarrow b \Leftrightarrow a \leftarrow u \rightarrow b$:

$$p(x_1,...,x_4) = \left[\sum_{u_1} p(x_1 \mid u_1) p(x_3 \mid x_2, u_1) p(u_1)\right] \times \\ \left[\sum_{u_2} p(x_2 \mid x_1, u_2) p(x_4 \mid x_3, u_2) p(u_2)\right] \\ = q_{1,3}(x_1, x_3 \mid x_2) \times q_{2,4}(x_2, x_4 \mid x_1, x_3).$$

$$\begin{split} p(x_1,\ldots,x_4) &= q_{1,3}(x_1,x_3 \mid x_2) \times q_{2,4}(x_2,x_4 \mid x_1,x_3) \\ &= q_{D_1}(x_{D_1} \mid \mathsf{pa}_\mathcal{G}(D_1)) \times q_{D_2}(x_{D_2} \mid \mathsf{pa}_\mathcal{G}(D_2)). \end{split}$$

For general case, district factorization:

$$\mathbb{P}(V) = \prod_{D \in \mathcal{D}(\mathcal{G})} q_D(x_D \mid \mathsf{pa}_{\mathcal{G}}(D)). \tag{1}$$

Each factor q_Y(y | W) is called a *kernel*, i.e. a probability density of Y with W being a parameter: ∑_y q_Y(y | W = w) = 1, ∀w.
q_Y(y | W = w) = ℙ(Y = y | do(w)) and thus, in general q_Y(y | W) ≠ ℙ(Y = y | W = w).

Factorizations on ADMGs

Express $q_D(x_D \mid pa_{\mathcal{G}}(D))$ as $\prod_{i \in D} p(x_i \mid \cdots)$:

• The Markov blanket of $a \in V$ in ADMG \mathcal{G} is

$$\mathsf{mb}(a,\mathcal{G}) := \mathsf{pa}_{\mathcal{G}}(D) \cup (D \setminus \{a\}),$$

where $D = \operatorname{dis}_{\mathcal{G}}(a)$. We have $a \perp \operatorname{nd}(a) | \operatorname{mb}(a)$.

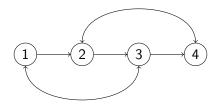
• Suppose that $1 \prec \cdots \prec p = |V|$ is a topological sort of \mathcal{G} . Let $V_i = \{1, \ldots, i\}$ and \mathcal{G}_i be the induced subgraph on V_i . Then $X_i \perp X_k \mid \mathsf{mb}(i, \mathcal{G}_i), k < i$:

$$q_D(x_D \mid \mathsf{pa}_{\mathcal{G}}(D)) = \prod_{i \in D} p(x_i \mid \mathsf{mb}(i, \mathcal{G}_i)). \tag{2}$$

Putting together into (1), we get

$$\mathbb{P}(V) = \prod_{i \in V} p(x_i \mid \mathsf{mb}(i, \mathcal{G}_i)). \tag{3}$$

Factorizations on ADMGs



 $\begin{array}{l} \text{Sort: } 1 \prec 2 \prec 3 \prec 4. \\ \text{mb}(1, \mathcal{G}_1) = \varnothing, \\ \text{mb}(2, \mathcal{G}_2) = \{1\}, \\ \text{mb}(3, \mathcal{G}_3) = \{1, 2\}, \\ \text{mb}(4, \mathcal{G}_4) = \{1, 2, 3\}. \end{array}$

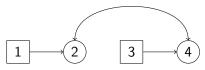
$$q_{1,3}(x_1, x_3 \mid x_2) = p(x_1)p(x_3 \mid x_1, x_2),$$
(4)

$$q_{2,4}(x_2, x_4 \mid x_1, x_3) = p(x_2 \mid x_1)p(x_4 \mid x_1, x_2, x_3).$$
(5)

$$\Rightarrow p(x) = p(x_1)p(x_2 \mid x_1)p(x_3 \mid x_1, x_2)p(x_4 \mid x_1, x_2, x_3).$$

This does NOT imply any conditional independence among X_1, \dots, X_4 . In particular, $X_1 \not\perp X_4 \mid S$ for any $S \subseteq \{X_2, X_3\}$ (*m*-connected) even though no edge between X_1 and X_4 . No edge between X_1 and X_4 encodes a generalized conditional independence a.k.a. Verma constraint (Verma and Pearl 1990).

Represent $q_{2,4}(x_2, x_4 | x_1, x_3) = p(x_2, x_4 | do(x_1, x_3))$ by a conditional ADMG (CADMG) with graph $\mathcal{G}^{|W|}(W = \{1, 3\})$ by cutting all edges with an arrow into W:



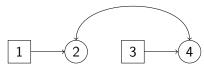
- Two types of vertices in a CADMG G(V, W):
 (i) Random V = {2,4}; (ii) Fixed W = {1,3}.
- Kernel q_V(x_V | x_W) is an (intervention) distribution for V after fixing W.
- We may further fix other random vertices if they are *fixable*.

Generalized CI constraints

Definition 1

The set of *fixable* vertices in a CADMAG $\mathcal{G}(V, W)$ is $F(\mathcal{G}) := \{v \in V : \operatorname{dis}_{\mathcal{G}}(v) \cap \operatorname{de}_{\mathcal{G}}(v) = \{v\}\}.$

v is fixable if none of its descendants is in the same district.



Fixable vertices = {2, 4}.

Fix vertex 2: (i) $\mathcal{G}(V = \{4\}, W = \{1, 2, 3\})$



(ii) New kernel district-factorized according to $\mathcal{G}(\{4\},\{1,2,3\})$:

 $q_4(x_4 \mid x_2, x_1, x_3) = f_4(x_4 \mid x_3)$. nested factorization (6)

The new kernel $q_4(x_4 | x_2, x_1, x_3)$ is defined by the fixing operator:

Definition 2

Given a kernel $q_V(x_V | W)$ associated with a CADMG $\mathcal{G} = \mathcal{G}(V, W)$, for any fixable vertex $r \in F(\mathcal{G})$, the fixing operator ϕ_r yields a new kernel

$$q_{V\setminus r}(x_{V\setminus r} \mid r, W) = \phi_r(q_V; \mathcal{G}) := \frac{q_V(x_V \mid W)}{q_V(x_r \mid \mathsf{mb}(r, \mathcal{G}), W)}.$$
 (7)

- $q_V(x_r \mid mb(r, \mathcal{G}), W)$ is a conditional distribution calculated from $q_V(x_V \mid W)$.
- If r is fixable, then r can be sorted as the last vertex in its district and its causal effect $\mathbb{P}(V \setminus r \mid do(r); \mathcal{G})$ on $V \setminus r$ can be calculated by (7).

Generalized CI constraints

Apply
$$\phi_2$$
 on $q_{2,4}(x_2, x_4 \mid x_1, x_3)$ (mb(2, G) = {1, 4, 3}):

$$q_{4}(x_{4} \mid x_{2}, x_{1}, x_{3}) = \phi_{2}(q_{2,4}; \mathcal{G}) = \frac{q_{2,4}(x_{2}, x_{4} \mid x_{1}, x_{3})}{q_{2,4}(x_{2} \mid x_{4}, x_{1}, x_{3})}$$
$$= q_{2,4}(x_{4} \mid x_{1}, x_{3})$$
$$= \sum_{x'_{2}} q_{2,4}(x'_{2}, x_{4} \mid x_{1}, x_{3})$$
$$= \sum_{x'_{2}} p(x'_{2} \mid x_{1})p(x_{4} \mid x_{1}, x'_{2}, x_{3}). \text{ by (5)}$$

By nested factorization (6):

$$\sum_{x'_2} p(x'_2 \mid x_1) p(x_4 \mid x_1, x'_2, x_3) = f_4(x_4 \mid x_3)$$

does not depend on x_1 , which is a GCI constraint.

Nested factorization:

- Suppose p(x) factorizes by a DAG G(V ∪ L) and G = G(V) is the ADMG defined by latent projection.
- For a valid fixing sequence w = (w₁,..., w_r), let φ_w(G) be the CADMG after fixing w sequentially and D_w = D(φ_w(G)) be the districts of (random vertices) in φ_w(G).

Theorem 1 (Richardson et al. (2023))

For any valid fixing sequence w,

$$\phi_w(p(x_V);\mathcal{G}) = \prod_{D \in \mathcal{D}_w} f_D^w(x_D \mid \mathsf{pa}_{\mathcal{G}}(D))$$

for some kernels $f_D^w(x_D \mid pa_G(D))$.

Algorithm to find systematically CI and GCI constraints implied by ADMG: Tian and Pearl (2002b).

Input: ADMG $\mathcal{G}(V)$; assume V is sorted, $1 \prec \ldots \prec p$. Output: CI and GCI constraints on $p(x_V)$ implied by $\mathcal{G}(V)$. For i = 1 to p, Part 1: CI constraints $X_i \perp X_k \mid mb(i, \mathcal{G}_i), k < i, k \notin mb(i, \mathcal{G}_i)$. Part 2: $S \leftarrow dis_{\mathcal{G}_i}(i)$ and $G \leftarrow \phi_{[i] \setminus S}(\mathcal{G}_i)$ ($[i] = \{1, \ldots, i\}$). For each descendent set $D \subset S$ s.t. $i \notin D$: Let $D' = S \setminus D$. 1 $\sum_{x_D} q_S = q_{D'}$ (fixing D); $G' = \phi_D(G)$.

- 2 If G' has 2 or more districts, $E \leftarrow \operatorname{dis}_{G'}(i)$ and $q_{D'} / \sum_{x_i} q_{D'}$ is a function of $\operatorname{mb}(i, G') = E \cup \operatorname{pa}_{G'}(E)$.
- **3** Repeat part 2 with $S \leftarrow E$ and $G \leftarrow \phi_{S \setminus E}(G)$.

Identification of causal effects given an ADMG $\mathcal{G}(V)$:

- Let $k \in V$ be a single variable and $S \subset V$.
- The causal effect of X_k on S is identifiable (from observational data) if P(S | do(X_k)) can be computed from the joint distribution P(V).

Theorem 2 (Tian and Pearl (2002a))

If there is no bidirected path connecting X_k to any of its children in $\mathcal{G}_{an(S)}$, then the causal effect of X_k on S is identifiable.

 Recent results: Theorem 48 in Richardson et al. (2023), Corollary 16 in Bhattacharya et al. (2022).

Identification of causal effects

Constructive proof of Theorem 2:

1 Let $V = \operatorname{an}(S)$, $\mathcal{G} = \mathcal{G}_{\operatorname{an}(S)}$ and $M = V \setminus \{S \cup k\}$. Then

$$p(x_{\mathcal{S}} \mid do(x_k)) = \sum_{x_{\mathcal{M}}} p(x_{V \setminus k} \mid do(x_k)).$$

2 Let
$$D = {\sf dis}_{\mathcal{G}}(k) \in \mathcal{D} = \mathcal{D}(\mathcal{G}).$$
 Since ${\sf ch}(k) \cap D = arnothing$,

$$p(x_{V\setminus k} \mid do(x_k)) = \sum_{x'_k} q_D(x_D \mid \mathsf{pa}_\mathcal{G}(D)) \prod_{D' \in \mathcal{D}} q_{D'}(x_{D'} \mid \mathsf{pa}_\mathcal{G}(D')).$$

If X_k is fixable, we may instead apply fixing operator:

$$p(x_{V\setminus k} \mid do(x_k)) = \phi_k(p(x); \mathcal{G}) = \frac{p(x_V)}{p(x_k \mid \mathsf{mb}(k, \mathcal{G}))}.$$

Identification of causal effects

The identify algorithm by Tian and Pearl (2002a) reformulated with fixing operators: Theorem 48 in Richardson et al. (2023).

Let
$$\mathcal{G} = \mathcal{G}(V)$$
. For $A, Y \subset V$, want to identify $\mathbb{P}(Y \mid do(A))$.
• Let $Y^* = \operatorname{an}_{\mathcal{G}_{V \setminus A}}(Y) \supseteq Y$: there is a directed path from every $v \in Y^*$ to Y not blocked by A.
Since $V \setminus (A \cup Y) = [V \setminus (A \cup Y^*)] \cup (Y^* \setminus Y)$,
 $\mathbb{P}(Y \mid do(A)) = \sum_{V \setminus (A \cup Y)} \mathbb{P}(V \setminus A \mid do(A))$
 $= \sum_{Y^* \setminus Y} \sum_{V \setminus (A \cup Y^*)} \mathbb{P}(V \setminus A \mid do(A))$
 $= \sum_{Y^* \setminus Y} \mathbb{P}(Y^* \mid do(A)), \quad (Y^* \text{ is ancestral}).$

• Let $\mathcal{D}^* = \mathcal{D}(\mathcal{G}_{Y^*})$. District factorization on \mathcal{G}_{Y_*} shows

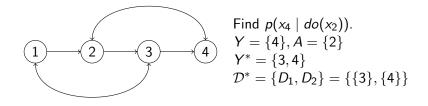
$$\mathbb{P}(Y^* \mid do(A)) = \prod_{D \in \mathcal{D}^*} \mathbb{P}[D \mid do(\operatorname{pa}_{\mathcal{G}}(D))].$$

If every *D* is intrinsic (i.e. $V \setminus D$ is fixable), then $\mathbb{P}[D \mid do(pa_{\mathcal{G}}(D))] = \phi_{V \setminus D}(\mathbb{P}(V); \mathcal{G})$, and

$$\therefore \quad \mathbb{P}(Y \mid do(A)) = \sum_{Y^* \setminus Y} \prod_{D \in \mathcal{D}^*} \phi_{V \setminus D}(\mathbb{P}(V); \mathcal{G}).$$
(8)

Otherwise, $\mathbb{P}[D \mid do(pa_{\mathcal{G}}(D))]$ is not identifiable for some D, and $\mathbb{P}(Y \mid do(A))$ is not identifiable.

Identification of causal effects



$$p(x_3 \mid do(x_2)) = \phi_{1,2,4}(p(x_V); \mathcal{G}) = \phi_1(q_{1,3}(x_1, x_3 \mid x_2); \mathcal{G}^{\mid 2,4})$$

= $\sum_{x_1} p(x_1)p(x_3 \mid x_1, x_2).$
$$p(x_4 \mid do(x_3)) = \phi_{2,1,3}(p(x_V); \mathcal{G}) = \phi_2(q_{2,4}(x_2, x_4 \mid x_1, x_3); \mathcal{G}^{\mid 1,3})$$

= $\sum_{x_2'} p(x_2' \mid x_1)p(x_4 \mid x_1, x_2', x_3).$

$$\therefore \quad p(x_4 \mid do(x_2)) = \sum_{x_3} p(x_3 \mid do(x_2)) p(x_4 \mid do(x_3)).$$

Given an ADMG G with directed edge set E_d and bidirected edge set E_b , define linear SEM

$$X_{j} = \sum_{i \in pa_{\mathcal{G}}(j)} \beta_{ij} X_{i} + \varepsilon_{j}, \quad j = 1, \dots, p.$$

$$(\varepsilon_{1}, \dots, \varepsilon_{p}) \sim \mathcal{N}_{p}(0, \Omega).$$

$$(9)$$

■
$$B \in \mathcal{B}(E_d) := \{ (\beta_{ij})_{p \times p} : \beta_{ij} = 0 \text{ if } i \to j \notin E_d \}.$$

■ $\Omega \in \mathcal{P}(E_b) := \{ (\omega_{ij})_{p \times p} : \omega_{ij} = 0 \text{ if } i \leftrightarrow j \notin E_b \}.$

The linear SEM (9) defines a family of multivariate Gaussian distributions $\mathcal{N}_{\rho}(0, \Sigma)$ with

$$\Sigma = \Sigma_{\mathcal{G}}(B, \Omega) := (\mathbf{I} - B)^{-\mathsf{T}} \Omega (\mathbf{I} - B)^{-1}$$

Definition 3 (Identifiability)

The linear SEM for an ADMG \mathcal{G} is said to be identifiable if $\Sigma_{\mathcal{G}}(B, \Omega)$ is an *injective* (one-to-one) map from $\mathcal{B}(E_d) \times \mathcal{P}(E_b)$ to the set of positive definite matrices.

Reachable closure (Shpitser et al. 2018).

Definition 4

For a CADMG $\mathcal{G}(V, W)$, a reachable subset $C \subseteq V$ is called a reachable closure for $S \subseteq C$ if the set of fixable vertices in $\phi_{V \setminus C}(\mathcal{G})$ is a subset of S.

- Reachable closure is unique for any $S \subseteq V$, denoted $\langle S \rangle$.
- $\langle S \rangle$ is the set of random vertices in $\phi_{\neg S}(\mathcal{G})$ (fixing as many vertices in $V \setminus S$ as possible).

Graphical criterion for identifiability:

Theorem 3 (Drton et al. (2011))

The linear SEM for an ADMG $\mathcal{G}(V)$ is identifiable if and only if $\langle v \rangle = \{v\}$ for all $v \in V$.

Identifiability means that given G(V) and Σ, there is a unique set of parameters (B, Ω) for the linear SEM. Thus, given G(V) and data, one may estimate (B, Ω).

• Example: $a \rightarrow s \leftarrow b$ and $b \leftrightarrow a \leftrightarrow s$.

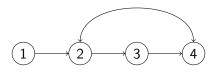
- $\langle s \rangle = \{a, b, s\}$ $(a, b \text{ are not fixable in } V \setminus s)$.
- *G*_{*a,b,s*} contains a sink node *s* and its parents *a*, *b* in the same district.
- Linear SEM is *not* identifiable.

Motivations.

- A class of ADMGs that represents conditional independences among V in a DAG $\mathcal{G}(V, L)$ with latent variables L.
- Retains the ancestral relationships and hence causal relations among V.
- Its equivalence class can be constructed from CI relations learned from observational data.
- Does not preserve all confounding structures in G(V, L), i.e. bidirected edges in the latent projection.
- Does not represent GCI (Verma) constraints: potential loss of efficiency.

Definitions. Let $\mathcal{G} = (V, E)$ be an ADMG.

- An almost directed cycle occurs when a ↔ b and a ∈ an_G(b) (removing the arrowhead at b results in a directed cycle).
- Let L ⊂ V. An inducing path relative to L is a path on which every intermediate vertex ∉ L is a collider and every collider is an ancestor of an endpoint. If L = Ø, call it an inducing path.



Almost directed cycle: (2,3,4,2). Inducing path: $1 \rightarrow 2 \leftrightarrow 4$ $\Rightarrow 1$ and 4 not *m*-separated by any subsets.

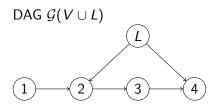
Definition 5 (MAG)

A mixed graph is a maximal ancestral graph (MAG) if

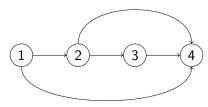
- (i) it does not contain any directed or almost directed cycles (ancestral);
- (ii) there is no inducing path between any two non-adjacent vertices (maximal).

Constructing MAG \mathcal{M} from DAG $\mathcal{G} = \mathcal{G}(V, L)$:

- **1** For each pair $a, b \in V$, a and b are adjacent in \mathcal{M} iff there is an inducing path between them relative to L in \mathcal{G} .
- 2 For each adjacent pair (a, b) in M, orient a → b in M if a ∈ an_G(b); orient b → a in M if b ∈ an_G(a); orient a ↔ b otherwise.



MAG



Every edge among V in a DAG (trivial inducing path) is an edge in MAG.

Inducing paths relative to *L*: $1 \rightarrow 2 \leftarrow L \rightarrow 4 \Rightarrow 1 \rightarrow 4$ in \mathcal{M} $2 \leftarrow L \rightarrow 4 \Rightarrow 2 \rightarrow 4$ in \mathcal{M} 1, 2 are ancestors of 4. Equivalence class of a MAG:

Two MAGs are Markov equivalent if they have the same set of *m*-separations.

Sufficient and necessary conditions: same skeleton and *v*-structures, and share some covered colliders (Proposition 2, Zhang (2008b)).

The equivalence class [M] of a MAG M is represented by a partial ancestral graph P:

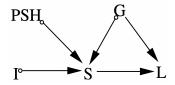
i \mathcal{P} has the same adjacencies (skeleton) as \mathcal{M} ;

ii A mark of arrowhead is in \mathcal{P} iff it is shared by all MAGs in $[\mathcal{M}]$;

iii A mark of tail is in \mathcal{P} iff it is shared by all MAGs in $[\mathcal{M}]$.

Edge marks in (ii) and (iii) are invariant across $[\mathcal{M}]$; other variable marks are represented by \circ in \mathcal{P} .

Example PAG (Zhang 2008a)



I: income, S: smoking, PSH: parent smoking habits, G: genotype, L: lung cancer

- $I \circ \rightarrow S = I \rightarrow S \text{ or } I \leftrightarrow S.$
- preserve the 3 v-structures at the collider S.
- no directed or almost directed cycles among G, S, L.

Constraint-based learning of MAGs by the FCI (fast causal inference) algorithm (Spirtes et al. 1999):

Use CI constraints learned from observational data to construct the equivalence class of a MAG represented by a PAG:

- skeleton;
- invariant marks (arrowheads and tails).

Algorithm outline

- 1: $E \leftarrow$ edge set of the complete undirected graph on V. Every edge is $\circ \circ$.
- 2: for $(i,j) \in E$ do
- 3: Search for a subset S_{ij} such that $X_i \perp X_j \mid S_{ij}$. If found, $E \leftarrow E \setminus \{(i,j), (j,i)\}$ and store S_{ij} .
- 4: end for
- 5: Orient edges in *v*-structures based on *E* and $\{S_{ij}\}$.
- 6: Apply orientation rules R1 to R4 (Zhang 2008b) until none of them applies.
- 7: Apply orientation rules R8 to R10 (Zhang 2008b) until none of them applies.

Suppose ${\mathcal M}$ is the true MAG, and assume we have CI oracle.

- Line 1 to 5: similar to the PC algorithm.
- After Line 4: correctly construct the skeleton $sk(\mathcal{M})$.
- After Line 6: identify all and only invariant arrowheads in [\mathcal{M}].
- After Line 7: identify all and only invariant tails in [\mathcal{M}].

Theorem 4 (Theorem 4, Zhang (2008b))

Given a perfect conditional independence oracle, the FCI algorithm returns the PAG for the true MAG $\mathcal{M}.$

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