## Structure Learning of DAGs

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Stats 212 Graphical Models Lecture Notes

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## Overview and assumptions

Structure learning: Let  $(\mathcal{G}, \mathbb{P})$  be a causal DAG model over  $X_1, \ldots, X_p$ . Given data  $x_i = (x_{i1}, \ldots, x_{ip}) \sim (\mathcal{G}, \mathbb{P}), i = 1, \ldots, n$ , how to estimate the DAG  $\mathcal{G}$ ?

- Constraint-based methods: Conditional independence tests against  $X_i \perp X_j \mid X_S$  for all i, j, S.
- Score-based methods: Optimizing a scoring function over graph space.
- Hybrid methods: First use constraint-based method to prune the search space, and then apply a score-based method to search for the optimal DAG.

See, e.g. Aragam et al. (2019) Section 1 for recent literature.

Data types:

- Observational data (no intervention)
- Experimental data (intervention available)

Main assumptions: (1) causal sufficiency; (2) faithfulness.

### Definition 1 (Causal sufficiency)

A set of variables V is causally sufficient if every common cause of any two or more variables in V is also in V.

- For *G*, this means that every common ancestor of two or more nodes is observed.
- In SEM  $X_i = f_i(PA_i, \varepsilon_i)$ ,  $i \in V$ , causal sufficiency implies  $\varepsilon_i$ 's are mutually independent.

### Definition 2 (Faithfulness)

For a graphical model  $(\mathcal{G}, \mathbb{P})$ , we say the distribution  $\mathbb{P}$  is faithful to the graph  $\mathcal{G}$  if for every triple of disjoint sets  $A, B, S \subset V$ ,

 $X_A \perp X_B \mid X_S \Leftrightarrow S$  separates (*d*-separates) A and B.

• Conditional independence (CI) in  $\mathbb{P} \Leftrightarrow d$ -separation in  $\mathcal{G}$ , i.e.

$$\mathcal{I}_{\mathbb{P}}(A, B|S) \Leftrightarrow \mathcal{D}_{\mathcal{G}}(A, B|S).$$

- Given *G*, almost all parameter values in the SEMs will define a faithful P.
- Structure learning: use CI relations learned from data to infer edges in *G*.

# Equivalence class and CPDAG

Suppose we only have observational data. What can be learned?

### Definition 3 (Markov equivalence)

Two DAGs  $\mathcal{G}$  and  $\mathcal{G}'$  on the same set of nodes V are Markov equivalent if  $\mathcal{D}_{\mathcal{G}}(X, Y|\mathbf{Z}) \Leftrightarrow \mathcal{D}_{\mathcal{G}'}(X, Y|\mathbf{Z})$  for any  $X, Y \in V$  and  $\mathbf{Z} \subseteq V \setminus \{X, Y\}$ .

- Two DAGs are Markov equivalent if and only if they have the same skeletons and the same v-structures.
- A v-structure is a triplet {i, j, k} ⊆ V of the form i → k ← j:
   i and j are nonadjacent; k is called an uncovered collider.
- Equivalent DAGs form an equivalence class.
- DAGs in the same equivalence class cannot be distinguished from observational data. Thus we can only learn the equivalence class of G from observational data.

How to represent an equivalence class? CPDAG (Completed partially DAG).

Two types of edges in a DAG  $\mathcal{G}$ :

- A directed edge i → j is compelled in G if for every DAG G' equivalent to G, the edge i → j exists in G'.
- If an edge is not compelled in  $\mathcal{G}$ , then it is *reversible*.

#### Definition 4 (CPDAG or essential graph)

The CPDAG of an equivalence class is the PDAG consisting of a directed edge for every compelled edge in the equivalence class, and an undirected edge for every reversible edge in the equivalence class.

Equivalence class  $[\mathcal{G}_1] = \{\mathcal{G}_1, \mathcal{G}_2, \mathcal{G}_3\}$  and CPDAG  $\mathcal{G}$ :



Red: compelled edges, same orientation in all equivalent DAGs. Black: reversible edges, either direction occurs in at least one equivalent DAG. Characterization of CPDAGs (or essential graphs):

Theorem 1 (Andersson et al. (1997))

A graph  ${\cal G}$  is a CPDAG for some DAG if and only if  ${\cal G}$  satisfies the following conditions:

- **1**  $\mathcal{G}$  is a chain graph.
- **2**  $\mathcal{G}_{\tau}$  is chordal for every chain component  $\tau$  of  $\mathcal{G}$ .
- 3 The configuration  $a \rightarrow b c$  does not occur as an induced subgraph of  $\mathcal{G}$ .
- **4** Every arrow  $a \rightarrow b$  in  $\mathcal{G}$  is strongly protected.

- Chordal graph: An undirected graph is chordal if every cycle of length n ≥ 4 possesses a chord, that is an edge between two nonconsecutive vertices on the cycle. (Triangulated graph)
- An arrow a → b is strongly protected in G if it occurs in at least one of the following configurations as an induced subgraph:

(a): 
$$a \longrightarrow b$$
 (b):  $a \longrightarrow b$  (c):  $a \longrightarrow b$  (d):  $a \longrightarrow c$ 

#### Theorem 2 (Spirtes et al. (1993))

Suppose  $(\mathcal{G}, \mathbb{P})$  satisfies the faithfulness assumption. Then there is no edge between a pair of nodes  $X, Y \in V$  if and only if there exists a subset  $\mathbf{Z} \subseteq V \setminus \{X, Y\}$  such that  $\mathcal{I}_P(X, Y | \mathbf{Z})$ .

Constraint-based methods:

- **1** Find the skeleton of  $\mathcal{G}$  by CI tests;
- 2 Identify v-structures;
- **3** Orient other edges.

Output: CPDAG (or PDAG)

Outline of PC algorithm (Spirtes and Glymour 1991):

- 1:  $E \leftarrow$  edge set of the complete undirected graph on V.
- 2: for  $(i,j) \in E$  do
- 3: Search for a subset  $S_{ij}$  of either  $N_i(E)$  or  $N_j(E)$  such that  $X_i \perp X_j \mid S_{ij}$ . If found,  $E \leftarrow E \setminus \{(i,j), (j,i)\}$  and store  $S_{ij}$ .
- 4: end for
- 5: Identify v-structures based on E and  $\{S_{ij}\}$ .
- 6: Orient as many edges in E as possible by Meek's rules.

Notes:

- **1** Line 3:  $N_i(E) = \{X_k : (i, k) \in E\}.$
- 2 For loop: implemented in ascending order of  $|S_{ij}| = \ell$  for  $\ell = 0, \ldots, \ell_{max}$ .
- **3** Line 1 to 4: Estimate skeleton  $sk(\widehat{\mathcal{G}})$  of  $\mathcal{G}$ .

Edge orientation steps:

- Identify v-structures (Line 5) given sk(G): For all nonadjacent pair (i, j) with a common neighbor k, orient i - k - j as i → k ← j if k ∉ S<sub>ij</sub>. Because otherwise, X<sub>i</sub> ⊥ X<sub>j</sub> | S<sub>ij</sub>, contradiction. After this step, we obtain a PDAG.
- 2 Meek's rules (Line 6): In the resulting PDAG, orient as many undirected edges as possible by repeated application of four rules (Meek 1995).
   Basic idea: If orienting an undirected edge *i* − *j* into *i* → *j* would result in additional *v*-structures or a directed cycle.

then orient it into  $i \leftarrow j$ .

### Meek's rules:



dashed line in R4: undirected or directed with either orientation

# Constraint-based learning

Conditional independence tests (H<sub>0</sub> : X ⊥ Y | S):
Gaussian data: partial correlation cor(X, Y | S) = 0.
Sample covariance matrix Σ̂ from data columns of (X, Y, S).
Ω̂ = (ω<sub>ij</sub>) ← Σ̂<sup>-1</sup> and ρ̂<sub>XY|S</sub> = -ω<sub>12</sub>/√ω<sub>11</sub>ω<sub>22</sub>.
Bisher z-transformation.

$$z(X, Y|S) = rac{1}{2} \log \left( rac{1 + \widehat{
ho}_{XY|S}}{1 - \widehat{
ho}_{XY|S}} 
ight)$$

and 
$$\sqrt{n-|S|-3} \cdot z(X,Y|S) \mid H_0 \sim \mathcal{N}(0,1).$$

Discrete data:  $G^2$  or  $\chi^2$  test for conditional independence.

$$G^{2}(X, Y; S = s) = 2 \sum_{x,y} O_{xys} \log(O_{xys}/E_{xys}),$$
  

$$G^{2}(X, Y; S) = \sum_{s} G^{2}(X, Y; S = s) \mid H_{0} \sim \chi^{2}_{(|X|-1)(|Y|-1)|S|},$$

 $E_{xys}$ : expected counts under  $H_0$ ;  $O_{xys}$ : observed counts.

Correctness and consistency:

Let  $\widehat{\mathcal{G}}_n$  be the estimated graph by PC from a sample of size *n* and  $\mathcal{C}$  be the CPDAG of  $\mathcal{G}$ . Suppose that  $\mathbb{P}$  is faithful to  $\mathcal{G}$ .

- CI oracles (Spirtes et al. 1993; Meek 1995): If all CI tests are perfect (CI oracles), then  $\widehat{\mathcal{G}}_n = \mathcal{C}$  and all found separating sets  $|S_{ij}| \leq \max\{|PA_i|, |PA_j|\}.$
- 2 Large-sample limit: When the sample size n → ∞, all CI tests involved will be perfect (no type I or II error) with high probability. Then the PC algorithm estimates the CPDAG of G consistently, i.e.

$$\lim_{n\to\infty}\mathbb{P}(\widehat{\mathcal{G}}_n=\mathcal{C})=1.$$

Score-based methods:

$$\widehat{\mathcal{G}} = \underset{\substack{G \in Space}}{\operatorname{argmax}} S(G, \mathbf{D}). \tag{1}$$

- $\mathbf{I} \ \mathbf{D} = (x_{ij})_{n \times p} = [X_1 \mid \ldots \mid X_p] \text{ i.i.d. data from } (\mathcal{G}, \mathbb{P}).$
- S(G, D) is a scoring function: log-likelihood of D given a graph G with a penalty term on model complexity (number of edges or number of free parameters). For example,

$$S_{\mathsf{BIC}}(G, \mathbf{D}) = \log p(\mathbf{D} \mid \widehat{\theta}, G) - \frac{d}{2} \log n, \qquad (2)$$

 $\widehat{\theta}$ : MLE of parameters under *G*, *d* = dimension of  $\theta$ .

Space of graphs: DAGs, equivalence class (CPDAGs) or topological sorts.

BIC score for Gaussian DAGs:

• Liner SEM for data columns  $X_j \in \mathbb{R}^n, j \in [p]$ :

$$X_j = \sum_{i \in PA_j} \beta_{ij} X_i + \varepsilon_j, \qquad \varepsilon_j \sim \mathcal{N}_n(0, \omega_j^2 I_n).$$

Decomposable:

$$S_{\text{BIC}}(G, \mathbf{D}) = \sum_{j=1}^{p} s(X_j, PA_j^G)$$

$$= \sum_j \log p(X_j \mid \widehat{\beta}_j, \widehat{\omega}_j^2, PA_j^G) - \frac{1}{2} |PA_j^G| \log n.$$
(3)

 $(\widehat{\beta}_j, \widehat{\omega}_j^2)$ : MLEs in Gaussian regression  $X_j \sim PA_j^G$ .

Bayesian Dirichlet score for discrete DAGs (Heckerman et al. 1995):

• Multinomial distribution:  $\theta_{ijk} = \mathbb{P}(X_i = k \mid PA_i = j)$ . Parameter for  $[X_i \mid PA_i]$  is a  $q_i \times r_i$  table:

$$\Theta_i = \left\{ heta_{ijk} : j \in [q_i], k \in [r_i], ext{such that} \sum_{k=1}^{r_i} heta_{ijk} = 1 
ight\}.$$

• Assume a conjugate prior over  $\Theta_i$  given G

$$\Theta_i \mid PA_i \sim \text{Product-Dirichlet}((\alpha_{ijk})_{q_i \times r_i}) \Leftrightarrow \\ \theta_{ij} = (\theta_{ij1}, \dots, \theta_{ijr_i}) \mid PA_i \sim_{ind} \text{Dirichlet}(\alpha_{ij1}, \dots, \alpha_{ijr_i}).$$

Choose  $\alpha_{ijk} = \alpha/(r_i \cdot q_i)$ .

Assume a prior over  $G: P(G) \propto \lambda^{d(G)}, \lambda \in (0, 1)$  and  $d(G) = \sum_{i=1}^{p} r_i q_i$  number of parameters.

Given  $(G, \mathbf{D})$ , how to compute the BD score:  $(PA_i \equiv PA_i^G)$ 

• Contingency tables:  $N_{ijk} = \#\{PA_i = j \& X_i = k\}$  in **D**. For each node, a  $q_i \times r_i$  table:  $N_i = \{N_{ijk} : j \in [q_i], k \in [r_i]\}$ .

■ Marginal likelihood of N<sub>ij</sub> (one row) given PA<sub>i</sub>:

$$P(N_{ij} \mid PA_i) = \int P(N_{ij} \mid \theta_{ij}) \pi(\theta_{ij} \mid PA_i) d\theta_{ij}$$
  
=  $\frac{\Gamma(\alpha/q_i)}{\Gamma(N_{ij\bullet} + \alpha/q_i)} \prod_{k=1}^{r_i} \frac{\Gamma(N_{ijk} + \alpha/(q_i r_i))}{\Gamma(\alpha/(q_i r_i))},$ 

where  $N_{ij\bullet} = \sum_k N_{ijk}$  (row sum).

Marginal likelihood of N<sub>i</sub> (the whole table):

$$P(N_i \mid PA_i) = \prod_{j=1}^{q_i} P(N_{ij} \mid PA_i).$$

• Marginal likelihood of **D** (all *p* tables, one for each node):

$$P(\mathbf{D} \mid G) = \prod_{i=1}^{p} P(N_i \mid PA_i).$$

Posterior distribution

$$P(G \mid \mathbf{D}) \propto P(G)P(\mathbf{D} \mid G)$$
  
=  $\prod_{i=1}^{p} \lambda^{q_i r_i} \prod_{j=1}^{q_i} \frac{\Gamma(\alpha/q_i)}{\Gamma(N_{ij\bullet} + \alpha/q_i)} \prod_{k=1}^{r_i} \frac{\Gamma(N_{ijk} + \alpha/(q_i r_i))}{\Gamma(\alpha/(q_i r_i))}.$ 

BD score is decomposable:

$$S_{BD}(G, \mathbf{D}) := \log P(G) + \log P(\mathbf{D} \mid G) = \sum_{i=1}^{p} s(N_i, PA_i).$$
 (4)

Properties of the scoring functions (3) and (4):

- Score-equivalent: For any two Markov equivalent DAGs  $G_1$  and  $G_2$ , we have  $S(G_1, \mathbf{D}) = S(G_2, \mathbf{D})$ .
- Consistent (Chickering 2002): A scoring function S(G, •) is consistent if the following two properties hold for D<sub>n</sub> ~<sub>iid</sub> P:
  - 1 If  $\mathbb{P} \in G \setminus H$ , then  $\lim_{n} \mathbb{P}\{S(G, \mathbf{D}_{n}) > S(H, \mathbf{D}_{n})\} = 1$ .
  - 2 If  $\mathbb{P} \in G \cap H$  and d(G) < d(H), i.e. G has fewer parameters, then  $\lim_{n} \mathbb{P}\{S(G, \mathbf{D}_{n}) > S(H, \mathbf{D}_{n})\} = 1$ .

Haughton (1988) established:

**1**  $S_{\text{BIC}}(G, \bullet)$  (2) is consistent for exponential family.

**2**  $S_{BD}(G, \mathbf{D}_n) = S_{BIC}(G, \mathbf{D}_n) + O_{\rho}(1) = O_{\rho}(n) + O_{\rho}(1).$ 

Thus, both (3) and (4) are consistent scoring functions.

Consistency of score-based learning:

#### Theorem 3

Suppose  $\mathbb{P}$  is faithful to  $\mathcal{G}$  and  $\mathbf{D}_n \sim_{iid} \mathbb{P}$ . If  $S(G, \bullet)$  is consistent and score-equivalent, then

$$\lim_{n\to\infty} \mathbb{P}\left\{ \operatorname*{argmax}_{G} S(G,\mathbf{D}_n) = \mathcal{C} \right\} = 1,$$

where  $\mathcal{C} = [\mathcal{G}] := \{ G : G \simeq \mathcal{G} \}$  is the Markov equivalence class of  $\mathcal{G}$ .

Space and search:

- DAG space: greedy hill climbing (Heckerman et al. 1995; Gámez et al. 2011), stochastic search (e.g. Zhou (2011)).
- Topological sorts: Larranaga et al. (1996); Teyssier and Koller (2005).

Define score for a sort  $\pi \in \mathcal{P}$  (space of permutations): Then search for  $\widehat{\pi} = \operatorname{argmax}_{\pi \in \mathcal{P}} S(\pi, \mathbf{D})$ .

 Equivalence classes: Greedy Equivalence Search (GES) (Chickering 2002). Search over topological sorts:

• Define score for a sort  $\pi \in \mathcal{P}$  (space of permutations):

$$S(\pi, \mathbf{D}) := \max_{G \in \mathcal{D}(\pi)} S(G, \mathbf{D}),$$

where  $\mathcal{D}(\pi)$  is the set of DAGs that can be sorted by  $\pi$ .

- $S(\pi, \mathbf{D})$  can be calculated by dynamic programming when  $|PA_i| \leq d$  (small) for all *i*.
- Then search for  $\hat{\pi} = \operatorname{argmax}_{\pi \in \mathcal{P}} S(\pi, \mathbf{D})$  by optimization over permutation space.

GES (Greedy Equivalence Search):

Define score for an equivalence class  $\mathcal{E}$ :

$$S(\mathcal{E},\mathbf{D}):=S(G,\mathbf{D}), \quad \forall G\in \mathcal{E}.$$

 $S(\mathcal{E}, \mathbf{D})$  is well-defined if  $S(G, \mathbf{D})$  is score-equivalent.

- Neighbors: *E'* ∈ *N*<sup>+</sup>(*E*) iff there is *G* ∈ *E* to which a single edge addition results in a *G'* ∈ *E'*. Similarly define *N*<sup>-</sup>(*E*) via single edge deletion.
- Two phases of greedy search from an initial empty graph: Phase 1:  $\mathcal{E}^{t+1} \leftarrow \operatorname{argmax} \{ S(\mathcal{E}, \mathbf{D}) : \mathcal{E} \in \mathcal{N}^+(\mathcal{E}^t) \}$ . Phase 2:  $\mathcal{E}^{t+1} \leftarrow \operatorname{argmax} \{ S(\mathcal{E}, \mathbf{D}) : \mathcal{E} \in \mathcal{N}^-(\mathcal{E}^t) \}$ .
- In the large sample limit n→∞, *Ê* found by GES with the BIC or the BD score is the true equivalence class (pr → 1).

Continuous relaxation of the scoring function:

- Consider Gaussian DAGs for simplicity. The BIC score
   S<sub>BIC</sub>(G, D) (3) is over a discrete space and hard to optimize.
- B = (β<sub>ij</sub>) = [β<sub>1</sub> | ··· | β<sub>p</sub>] and Ω = diag(ω<sub>j</sub><sup>2</sup>).
   Maximum regularized likelihood (Fu and Zhou 2013; Aragam and Zhou 2015):

$$(\widehat{B},\widehat{\Omega}) = \operatorname*{argmax}_{B \in \mathcal{B},\Omega} \sum_{j=1}^{p} \log p(X_j \mid X\beta_j, \omega_j^2) - \lambda_n \rho(\beta_j).$$
(5)

Compare regularizers:  $\ell_1$ , concave, and  $\ell_0$ .



Black:  $\ell_0$  penalty; Teal:  $\ell_1$  penalty; Blue: MCP; Red, dashed: Capped- $\ell_1$  penalty.

Maximizing regularized log-likelihood (5)

- Apply continuous optimization, such as block-wise coordinate descent, subject to acyclicity constraint (supp(B) defines a DAG), e.g. Fu and Zhou (2013); Aragam and Zhou (2015).
- Considering maximizing over topological sorts:

$$S(\pi, \mathbf{D}) := \max_{B \in \mathcal{B}(\pi), \Omega} \sum_{j=1}^{p} \log p(X_j \mid X \beta_j, \omega_j^2) - \lambda_n \rho(\beta_j).$$

 $\mathcal{B}(\pi)$ : weighted adjacency matrices compatible with  $\pi$ . Computed via *p* regularized regression problems (lasso or MCP) (Ye et al. 2021).

Reformulation of acyclicity constraint (Zheng et al. 2018):  $B \in \mathcal{B}$  if and only if h(B) = 0, where  $h(\cdot)$  is differentiable. Score-based learning with experimental data:

- If  $X_i$  is under intervention, i.e.  $do(X_i = x^*)$ : delete edges  $X_k \to X_i$  for all  $k \in PA_i$ .
- Let O<sub>i</sub> be the row indices of the data matrix **D** for which node X<sub>i</sub> is *not* under intervention (i.e. observational). Replace p(X<sub>i</sub> | PA<sub>i</sub>) by p(X<sub>O<sub>i</sub>i</sub> | PA<sub>O<sub>i</sub>i</sub>).

**1** Gaussian data: log-likelihood in (3) and (5) replaced by

$$\ell(B,\Omega;\mathbf{D}) = \sum_{j=1}^{p} \log p(X_{\mathcal{O}_j j} \mid X_{\mathcal{O}_j} \beta_j, \omega_j^2).$$
(6)

2 Multinomial data: Replace N<sub>ijk</sub> by

$$N_{ijk}(\mathcal{O}_i) = \#\{rows \in \mathcal{O}_i : PA_i = j \& X_i = k\}.$$

# Learning with experimental data

Identifiability of causal DAGs:

Assumptions:

(A1) The true parameter  $\Theta^*$  is faithful to  $\mathcal{G}$ .

- (A2) The parameter for  $[X_j | PA_j]$  is identifiable.
- (A3) Each node  $X_j$  is under intervention for  $n_j \gg \sqrt{n}$  data points.

### Theorem 4 (Gu et al. (2019))

Assume (A1), (A2) and (A3). Denote by  $\ell(\Theta; \mathbf{D}_n)$  the log-likelihood of the data  $\mathbf{D}_n$ . For any  $\Theta \neq \Theta^*$ ,

$$\lim_{n\to\infty} \mathbb{P}\{\ell(\Theta^*;\mathbf{D}_n) > \ell(\Theta;\mathbf{D}_n)\} = 1.$$

- **1** Gaussian data,  $\ell(\Theta; \mathbf{D}_n) = (6)$ .
- 2 Discrete data,  $\ell(\Theta; \mathbf{D}_n) = \sum_{i=1}^{p} \sum_{j,k} N_{ijk}(\mathcal{O}_i) \log \theta_{ijk}$ .

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