

UCLA STAT XL 10 Introduction to Statistical Reasoning

● **Instructor: Ivo Dinov,**
Asst. Prof. In Statistics and Neurology

University of California, Los Angeles, Spring 2002
<http://www.stat.ucla.edu/~dinov/>

STAT XL10, UCLA, Ivo Dinov Slide 1

Part IV: Chances, Variability, Probabilities and Proportions

- Chances and chance variability
- Where do probabilities come from?
- Simple probability models
- Probability rules
- Conditional probability
- Statistical independence

STAT XL10, UCLA, Ivo Dinov Slide 2

Let's Make a Deal Paradox – aka, Monty Hall 3-door problem

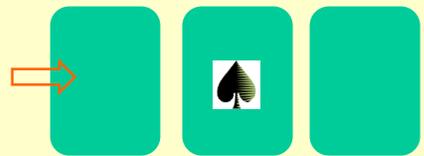
- This paradox is related to a popular television show in the 1970's. In the show, a contestant was given a choice of **three doors/cards** of which one contained a prize (**diamond**). The other two doors contained gag gifts like a chicken or a donkey (clubs).



Slide 3 STAT XL 10, UCLA, Ivo Dinov

Let's Make a Deal Paradox.

- After the contestant chose an initial door, the host of the show then revealed an empty door among the two unchosen doors, and asks the contestant if he or she would like to switch to the other unchosen door. The question is should the contestant switch. Do the odds of winning increase by switching to the remaining door?



1. Pick One card

2. Show one Club Card

3. Change 1st pick?

Slide 4 STAT XL 10, UCLA, Ivo Dinov

Let's Make a Deal Paradox.

- The *intuition* of most people tells them that each of the doors, the chosen door and the unchosen door, are equally likely to contain the prize so that there is a **50-50 chance** of winning with either selection? This, however, is **not the case**.
- The **probability of winning by using the switching technique is 2/3**, while the odds of winning by not switching is 1/3. The easiest way to explain this is as follows:

Slide 5 STAT XL 10, UCLA, Ivo Dinov

Let's Make a Deal Paradox

- The probability of picking the wrong door in the initial stage of the game is 2/3.
- If the contestant picks the wrong door initially, the host must reveal the remaining empty door in the second stage of the game. Thus, if the contestant switches after picking the wrong door initially, the contestant will win the prize.
- The probability of winning by switching then reduces to the probability of picking the wrong door in the initial stage which is clearly 2/3.
- StatGames.exe (Make a Deal Paradox)

Slide 6 STAT XL 10, UCLA, Ivo Dinov

Chance

- The **chance** of something happening gives the percentage of time it is expected to happen, when the basic process is repeatedly performed.
- E.g., What is the chance of getting an ace (1) if we roll a regular 6-face (hexagonal) die? 
- Chances are always between 0% - 100%.
- The chance of an event is equal to 100% - the chance of the opposite (complementary) event.
- E.g., $\text{Chance}(\text{getting } 1) = 100 - \text{Chance}(\text{2 or 3 or 4 or 5 or 6 turns up})$.

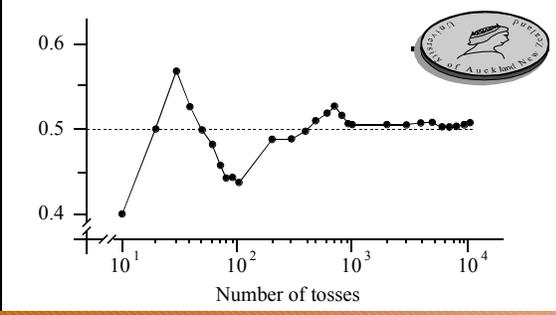
Slide 7 STAT 10, UCLA, Ivo Dinov

Coin toss experiments (Head vs. Tail)

- The **law of averages** about the behavior of coin tosses – the **relative proportion (relative frequency)** of heads-to-tails in a coin toss experiment becomes more and **more stable** as the **number of tosses increases**. The **law of averages** applies to **relative frequencies not absolute counts** of #H and #T.
- Two widely held **misconceptions** about what the **law of averages** about coin tosses:
 - Differences between the actual numbers of heads & tails becomes more and more variable with increase of the number of tosses – a seq. of 10 heads doesn't increase the chance of a tail on the next trial.
 - Coin toss results are **fair**, but behavior is still **unpredictable**.

Slide 8 STAT 10, UCLA, Ivo Dinov

Long run behavior of coin tossing – proportion of heads vs. number of tosses



Slide 9 STAT 10, UCLA, Ivo Dinov

Coin Toss Models

- Is the **coin tossing model** adequate for describing the **sex order** of children in families?
 - This is a **rough model** which is not exact. In most countries rates of B/G is different; from 48% ... to 52%, usually. Birth rates of boys in some places are higher than girls, however, female population seems to be about 51%.
 - **Independence**, if a second child is born the chance it has the same gender (as the first child) is slightly bigger.

Slide 10 STAT 10, UCLA, Ivo Dinov

Two die throw example

- What is the chance that the sum of the numbers, turning up when 2 dice are rolled, is equal to 8?
- Do the HTML Java-applet: UCLA_Classes/Applets.dir/DiceApplet.htm 

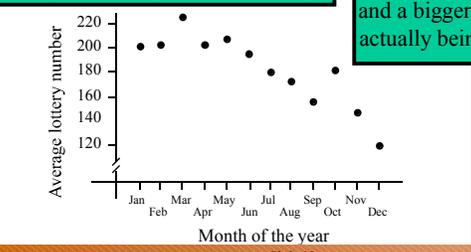
	1	2	3	4	5	6
1	2	3	4	5	6	7
2	3	4	5	6	7	8
3	4	5	6	7	8	9
4	5	6	7	8	9	10
5	6	7	8	9	10	11
6	7	8	9	10	11	12

Slide 11 STAT 10, UCLA, Ivo Dinov

Data from a "random" draw

366 cylinders (for each day in the year) for the US Vietnam war draft. The N-th drawn number, corr. to one B-day, indicating order of drafting.

So, people born later in the year tend to have lower lottery numbers and a bigger chance of actually being drafted.



Slide 12 STAT 10, UCLA, Ivo Dinov

Types of Probability

- Probability models have two essential components (*sample space*, the space of all possible outcomes from an experiment; and a list of *probabilities* for each event in the sample space). Where do the *outcomes* and the *probabilities* come from?
- Probabilities from models – say mathematical/physical description of the sample space and the chance of each event. Construct a fair die tossing game.
- Probabilities from data – data observations determine our probability distribution. Say we toss a coin 100 times and the observed Head/Tail counts are used as probabilities.
- Subjective Probabilities – combining data and psychological factors to design a reasonable probability table (e.g., gambling, stock market).

Slide 13 STAT.XL10, UCLA, Ivo Dinov

CA State Lottery – Supper Lotto Plus

- California Lotto, chose 5 out of 47 and choose one Mega from [1 : 27], fee \$1, your odds are 1 in 41,416,353! Why?
- 47-choose-5 = $[47!]/[(47-5)!(5!)] \rightarrow$
 $47\text{-choose-5} \times 27 = 1,533,939 \times 27 = 41,416,353$

Slide 14 STAT.XL10, UCLA, Ivo Dinov

Sample spaces and events

- A *sample space*, S , for a random experiment is the set of **all possible outcomes** of the experiment.
 - Examples?
- An *event* is a *collection of outcomes*.
 - Examples?
- An event *occurs* if **any outcome** making up that event **occurs**.
 - Examples?

Slide 15 STAT.XL10, UCLA, Ivo Dinov

The complement of an event

- The **complement** of an event A , denoted \bar{A} , occurs *if and only if* A does not occur.



(a) Sample space containing event A



(b) Event A shaded



(c) \bar{A} shaded

Slide 16 STAT.XL10, UCLA, Ivo Dinov

Combining events – all statisticians agree on

- “ A or B ” contains all outcomes in A or B (or both).
- “ A and B ” contains all outcomes which are **in both** A and B .



(a) Events A and B



(b) “ A or B ” shaded



(c) “ A and B ” shaded



(d) Mutually exclusive events

Mutually exclusive events cannot occur at the same time.

Slide 17 STAT.XL10, UCLA, Ivo Dinov

Probability distributions

- Probabilities always lie **between 0 and 1** and they **sum up to 1** (across all simple events).
- $pr(A)$ can be obtained by adding up the probabilities of all the outcomes in A .

$$pr(A) = \sum_{\substack{E \text{ outcome} \\ \text{in event } A}} pr(E)$$

Slide 18 STAT.XL10, UCLA, Ivo Dinov

Job losses in the US in \$1,000, 1987-1991

	Reason for Job Loss			Total
	Workplace moved/closed	Slack work	Position abolished	
Male	1,703	1,196	548	3,447
Female	1,210	564	363	2,137
Total	2,913	1,760	911	5,584

Slide 19 STAT XL 10, UCLA, Ivo Dinov

Job losses raw-data vs. proportions

	Workplace		Position	Total
	moved/closed	Slack work	abolished	
Male	1,703	1,196	548	3,447
Female	1,210	564	363	2,137
Total	2,913	1,760	911	5,584

	Reason for Job Loss			Row totals
	Workplace moved/closed	Slack work	Position abolished	
Male	.305	.214	.098	.617
Female	.217	.101	.065	.383
Column totals	.552	.315	.163	1.000

Slide 20 STAT XL 10, UCLA, Ivo Dinov

Review

- What is a **sample space**? What are the **two essential criteria** that must be satisfied by a possible sample space? (**completeness** – every outcome is represented; and **uniqueness** – no outcome is represented more than once.)
- What is an **event**? (collection of outcomes)
- If A is an event, what do we mean by its **complement**, \bar{A} ? When does \bar{A} occur?
- If A and B are events, when does **A or B** occur? When does **A and B** occur?

Slide 21 STAT XL 10, UCLA, Ivo Dinov

Example of probability distributions

- Tossing a coin twice. **Sample space** $S = \{HH, HT, TH, TT\}$, for a fair coin each outcome is equally likely, so the probabilities of the 4 possible outcomes should be identical, p . Since, $p(HH) = p(HT) = p(TH) = p(TT) = p$ and

$$p_k \geq 0; \sum_k p_k = 1$$

- $p = 1/4 = 0.25$.

Slide 22 STAT XL 10, UCLA, Ivo Dinov

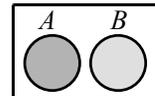
Proportion vs. Probability

- How do the concepts of a **proportion** and a **probability** differ? A **proportion** is a **partial description** of a real population. The **probabilities** give us the **chance** of something happening in a random experiment. Sometimes, **proportions** are **identical** to **probabilities** (e.g., in a real population under the experiment **choose-a-unit-at-random**).
- See the **two-way table of counts** (**contingency table**) E.g., **choose-a-person-at-random** from the ones laid off, and compute the chance that the person would be a **male**, laid off due to **position-closing**. We can apply the same rules for manipulating probabilities to proportions, in the case where these two are identical.

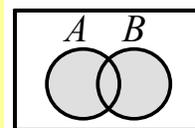
Slide 23 STAT XL 10, UCLA, Ivo Dinov

Rules for manipulating Probabilities

For mutually exclusive events,
 $\text{pr}(A \text{ or } B) = \text{pr}(A) + \text{pr}(B)$



$\text{pr}(A \text{ or } B) = \text{pr}(A) + \text{pr}(B) - \text{pr}(A \& B)$



Slide 24 STAT XL 10, UCLA, Ivo Dinov

Paradox of the Cavalier De Mere

- Betting on the event $E = \{\text{In 4 die rolls at least 1 ace turns up } \square\}$. $B = \{\text{In 24 rolls of a pair of dice, at least one double-ace shows up } \square\square\}$.
- **Claim:** $P(E) = P(B)!!?$
- **Reasoning E:** 1 roll gives a chance $1/6$ for an ace! So, in 4 rolls we have $4 \times 1/6 = 2/3$ to get at least 1 ace!
B: In one roll of a pair of dice, chance of a double-ace is $1/36$. So in 24 rolls we have $24 \times 1/36 = 2/3$ chance.
- **Experience** showed $P(E) > P(B)!!!$
- What's wrong? Well, extrapolating these arguments we get that the chance of getting 1 ace in 6 rolls is $6 \times 1/6 = 1$? Obviously, **incorrect!**

Slide 25 STAT XL 10, UCLA, Ivo Dinov

Paradox of the Cavalier De Mere

- The chance of **winning** (getting at least one ace) is hard to compute, but can we calculate the chance of **losing** – the complement event?!?
- Than *chance-of-winning* = $1 - \text{chance-of-losing}$.
- E^c , **complement** of E , = {none of 4 rolls shows an ace}.
- In one roll, chance of losing is $5/6$, no ace turns up.
- 2 die rolls are independent, hence we can use the multiplication rule, Chance of no ace in two rolls is $(5/6)^2$. Similarly, chance of 4 rolls with no ace, the probability $P(E^c) = (5/6)^4 \sim 0.482$.
- Game 2: Pair-of-dice: Chance of no-ace in 1 roll is $35/36$. Hence, $P(\{\text{no-Ace in 24 rolls}\}) = (35/36)^{24} \sim 0.509$.
- $P(\{\text{at-least-1-ace-in-4-rolls}\}) = 1 - 0.482 = 0.518 >>>>$
 $P(\{\text{at-least-1-double-ace-in-24-rolls}\}) = 1 - 0.509 = 0.491$.

Slide 26 STAT XL 10, UCLA, Ivo Dinov

Paradox of the Cavalier De Mere

- The chance of winning (getting at least one ace) is hard to compute, but can we calculate the chance of losing – the complement event?!?
- Than *chance-of-winning* = $1 - \text{chance-of-losing}$.
- E^c , **complement** of E , = {none of 4 rolls shows an ace}.
- In one roll, chance of losing is $5/6$, no ace turns up.
- 2 die rolls are independent, hence we can use the multiplication rule, Chance of no ace in two rolls is $(5/6)^2$. Similarly, chance of 4 rolls with no ace, the probability $P(E^c) = (5/6)^4 \sim 0.482$.
- Game 2: Pair-of-dice: Chance of no-ace in 1 roll is $35/36$. Hence, $P(\{\text{no-Ace in 24 rolls}\}) = (35/36)^{24} \sim 0.509$.
- $P(\{\text{at-least-1-ace-in-4-rolls}\}) = 1 - 0.482 = 0.518 >>>>$
 $P(\{\text{at-least-1-double-ace-in-24-rolls}\}) = 1 - 0.509 = 0.491$.

Slide 27 STAT XL 10, UCLA, Ivo Dinov

Conditional Probability

The **conditional probability** of A occurring **given** that B occurs is given by

$$\text{pr}(A | B) = \frac{\text{pr}(A \text{ and } B)}{\text{pr}(B)}$$

Slide 38 STAT XL 10, UCLA, Ivo Dinov

Multiplication rule- what's the percentage of Israelis that are **poor and Arabic**?

$\text{pr}(P \text{ and } A) = \text{pr}(P | A)\text{pr}(A) = \text{pr}(A | P)\text{pr}(P)$

Slide 39 STAT XL 10, UCLA, Ivo Dinov

Basic Rules for computing Probabilities

$\text{pr}(A \text{ and } B) = \text{pr}(A | B)\text{pr}(B) = \text{pr}(B | A)\text{pr}(A)$
 $\text{pr}(A) = 1 - \text{pr}(\bar{A})$

Properties of probabilities.

$\{p_k\}_{k=1}^N$ define probabilities $\Leftrightarrow p_k \geq 0; \sum_k p_k = 1$

Slide 40 STAT XL 10, UCLA, Ivo Dinov

A tree diagram for computing conditional probabilities

$$\text{pr}(A \text{ and } B) = \text{pr}(A | B)\text{pr}(B) = \text{pr}(B | A)\text{pr}(A)$$

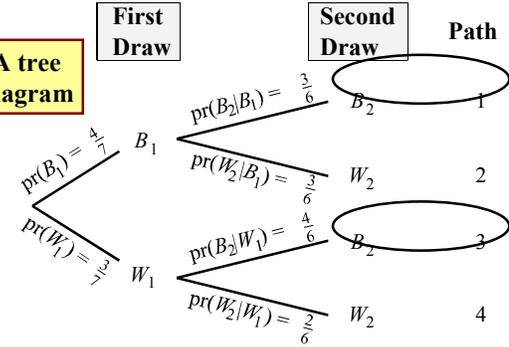
Suppose we draw 2 balls at random one at a time *without replacement* from an urn containing **4 black** and **3 white** balls, otherwise identical. What is the probability that the second ball is black? Sample Spc?

$$P(\{2\text{-nd ball is black}\}) = \text{Mutually exclusive}$$

$$P(\{2\text{-nd is black}\} \& \{1\text{-st is black}\}) + P(\{2\text{-nd is black}\} \& \{1\text{-st is white}\}) = 3/6 \times 4/7 + 4/6 \times 3/7 = 4/7.$$

Slide 41 STAT.XL 10, UCLA, Ivo Dinov

A tree diagram



Slide 42 STAT.XL 10, UCLA, Ivo Dinov

Conditional probabilities and 2-way tables

- Many problems involving conditional probabilities can be solved by constructing two-way tables
- This includes *reversing the order of conditioning*

$$P(A \& B) = P(A | B) \times P(B) = P(B | A) \times P(A)$$

Slide 43 STAT.XL 10, UCLA, Ivo Dinov

Number of Individuals Having a Given Mean Absorbance Ratio (MAR) in the ELISA for HIV Antibodies

MAR	Healthy Donor	HIV patients
<2	202	0
2 - 2.99	73	2

3 - 3.99	15	7
4 - 4.99	3	7
5 - 5.99	2	15
6 - 11.99	2	36
12+	0	21
Total	297	88

Slide 44 STAT.XL 10, UCLA, Ivo Dinov

HIV cont.

Disease status	Test result		Total
	Positive	Negative	
HIV	.98 × .01	?	.01 ← pr(HIV) = .01
Not HIV	?	.93 × .99	.99 ← pr(Not HIV) = .99
Total	?	?	1.00

$$86/88 \sim 0.98 \quad 275/297 \sim 0.93$$

Slide 45 STAT.XL 10, UCLA, Ivo Dinov

HIV – reconstructing the contingency table

$$\text{pr}(HIV \text{ and Positive}) = \text{pr}(Positive|HIV) \times \text{pr}(HIV) = .98 \times .01 = 98\% \text{ of } 1\%$$

$$\text{pr}(Not HIV \text{ and Negative}) = \text{pr}(Negative|Not HIV) \times \text{pr}(Not HIV) = .93 \times .99 = 93\% \text{ of } 99\%$$

Disease status	Test result		Total
	Positive	Negative	
HIV	.98 × .01	?	.01 ← pr(HIV) = .01
Not HIV	?	.93 × .99	.99 ← pr(Not HIV) = .99
Total	?	?	1.00

Disease Status	Test Result		Total
	Positive	Negative	
HIV	.0098	.0002	.01
Not HIV	.0693	.9207	.99
Total	.0791	.9209	1.00

Slide 46 STAT.XL 10, UCLA, Ivo Dinov

Statistical independence

- Events A and B are *statistically independent* if knowing whether B has occurred gives no new information about the chances of A occurring,

i.e. if $\text{pr}(A | B) = \text{pr}(A)$
- Similarly, $P(B | A) = P(B)$, since

$$P(B|A)=P(B \ \& \ A)/P(A) = P(A \ \& \ B)/P(A) = P(B)$$
- If A and B are *statistically independent*, then

$$\text{pr}(A \text{ and } B) = \text{pr}(A) \times \text{pr}(B)$$

Slide 47 STAT.XL.10.UCLA.Too.Dinner

People vs. Collins

Yellow car	$\frac{1}{10}$	Girl with blond hair	$\frac{1}{3}$
Frequencies assumed by the prosecution:			
Man with mustache	$\frac{1}{4}$	Black man with beard	$\frac{1}{10}$
Girl with ponytail	$\frac{1}{10}$	Interracial couple in car	$\frac{1}{1000}$

The first occasion where a conviction was made in an American court of law, largely on statistical evidence, 1964. A woman was mugged and the offender was described as a wearing **dark clothes**, with **blond hair** in a **pony tail** who got into a **yellow car driven by a black male** accomplice with **mustache** and **beard**. The suspect brought to trial were picked out in a line-up and fit all of the descriptions. Using the *product rule for probabilities* an expert witness computed the chance that a random couple meets these characteristics, as 1:12,000,000.

Slide 48 STAT.XL.10.UCLA.Too.Dinner

Examples

- Two coins are given. One is fair ($P(H)=0.5$) and the other is biased with $P(H)=2/3$. One of the coins is tossed once, resulting in H. The other is tossed three times, resulting in two heads. Which coin is more likely to be the biased one?
- We won't look for the probability of the first or the second coin being the biased one, rather we look for the probability of the given outcomes in two different cases: the first coin being the fair one, and the second--the biased one, and vice versa.
- If we assume that the first coin is fair, then the probability of the heads is $1/2$. The second coin must be the biased one, and the probability of it coming up with 2 heads and 1 tail in three tosses is $3 \times 2/3 \times 2/3 \times 1/3 = 4/9$. Note that there are three ways to get 2 heads: HHT, HTH, THH, the probability of each being $4/27$. Thus, the probability of both coins coming up with the given results is $2/9$.
- If, on the other hand, the first coin is the biased one, and the second coin is fair the probability of them resulting in the combination given in the problem is $(2/3) \times (3 \times 1/2 \times 1/2 \times 1/2) = 1/4$, or $2/8 > 2/9$. Therefore, it is more probable that the first coin is the biased one.

Slide 49 STAT.XL.10.UCLA.Too.Dinner

Examples

- Let X be the sum of spots from rolling 4 fair dice. Determine the expected value, the variance, and the standard deviation of the random variable X .
- Let $X = X_1 + X_2 + X_3 + X_4$, where X_1, X_2, X_3, X_4 be the random variables for the spots showing on the 1st, 2nd, 3rd, and 4th dice, respectively. [$E(X) = \text{Sum}(X_value \text{ times } P(x))$.]

$$E(Y) = \sum y \times P(Y = y) = \frac{\text{Sum_Of_All_Obs's}}{\text{Total_Number}}$$

- For each die,

$$E(X_1) = 1 \times 1/6 + 2 \times 1/6 + 3 \times 1/6 + 4 \times 1/6 + 5 \times 1/6 + 6 \times 1/6 = 3.5$$
- $E(X) = E(X_1) + \dots + E(X_4) = 3.5 + 3.5 + 3.5 + 3.5 = 14$

Slide 50 STAT.XL.10.UCLA.Too.Dinner

Examples

- Let X be the sum of spots from rolling 4 fair dice. Determine the expected value, the variance, and the standard deviation of the random variable X . $X = X_1 + X_2 + X_3 + X_4$.

$$E(Y) = \sum y \times P(Y = y) \quad ; \quad \text{Var}(Y) = \frac{1}{N-1} \sum (y - E(Y))^2$$

- Because the rolls are independent:

$$\text{Var}(X) = \text{Var}(X_1) + \text{Var}(X_2) + \text{Var}(X_3) + \text{Var}(X_4)$$
- The variance for any single roll is: $(1/5) \times (1-3.5)^2 + (1/5) \times (2-3.5)^2 + (1/5) \times (3-3.5)^2 + (1/5) \times (4-3.5)^2 + (1/5) \times (5-3.5)^2 + (1/5) \times (6-3.5)^2 = 3.5$.
- So, $\text{Var}(X) = 4 \times 3.5 = 14$. $\text{SD}(X) = \text{sqrt}(\text{Var}(X)) = 3.74$.
- So, from 4 dice, the expected value (Sum) is 14, with a SE of 3.74

Slide 51 STAT.XL.10.UCLA.Too.Dinner

Examples – Birthday Paradox

- The Birthday Paradox:** In a random group of N people, what is the change that at least two people have the same birthday?
- E.x., if $N=23$, $P > 0.5$. Main confusion arises from the fact that in real life we rarely meet people having the same birthday as us, and we meet more than 23 people.
- The reason for such high probability is that any of the 23 people can compare their birthday with any other one, not just you comparing your birthday to anybody else's.
- There are $N\text{-Choose-}2 = 20 \times 19/2$ ways to select a pair or people. Assume there are 365 days in a year, $P(\text{one-particular-pair-same-B-day}) = 1/365$, and
- $P(\text{one-particular-pair-failure}) = 1 - 1/365 \sim 0.99726$.
- For $N=20$, $20\text{-Choose-}2 = 190$. $E = \{\text{No 2 people have the same birthday is the event all 190 pairs fail (have different birthdays)}\}$, then $P(E) = P(\text{failure})^{190} = 0.99726^{190} = 0.59$.
- Hence, $P(\text{at-least-one-success}) = 1 - 0.59 = 0.41$, quite high.
- Note: for $N=42 \rightarrow P > 0.9 \dots$

Slide 52 STAT.XL.10.UCLA.Too.Dinner

Summary

- What does it mean for two events A and B to be *statistically independent*?
- Why is the working rule under independence, $P(A \text{ and } B) = P(A) P(B)$, just a special case of the multiplication rule $P(A \& B) = P(A | B) P(B)$?
- *Mutual independence* of events $A_1, A_2, A_3, \dots, A_n$ if and only if $P(A_1 \& A_2 \& \dots \& A_n) = P(A_1)P(A_2)\dots P(A_n)$

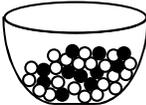
Slide 53 STAT.XL.10, UCLA, Ivo Dinov

Binomial Distribution

Slide 54 STAT.XL.10, UCLA, Ivo Dinov

The two-color urn model

N balls in an urn, of which there are
 M black balls
 $N - M$ white balls



Sample n balls and count $X = \#$ black balls in sample

We will compute the probability distribution of the R.V. X

Slide 55 STAT.XL.10, UCLA, Ivo Dinov

The biased-coin tossing model



toss 1 $\text{pr}(H) = p$ toss 2 $\text{pr}(H) = p$... toss n $\text{pr}(H) = p$

Perform n tosses and count $X = \#$ heads

We also want to compute the probability distribution of this R.V. X !
 Are the two-color urn and the biased-coin models related? How do we present the models in mathematical terms?

Slide 56 STAT.XL.10, UCLA, Ivo Dinov

The answer is: Binomial distribution

- The distribution of the number of heads in n tosses of a biased coin is called the *Binomial distribution*.

Slide 57 STAT.XL.10, UCLA, Ivo Dinov

Binomial(N, p) – the probability distribution of the number of Heads in an N -toss coin experiment, where the probability for Head occurring in each trial is p .
 E.g., Binomial(**6, 0.7**)

	x	0	1	2	3	4	5	6
Individual	$\text{pr}(X=x)$	0.001	0.010	0.060	0.185	0.324	0.303	0.118
Cumulative	$\text{pr}(X \leq x)$	0.001	0.011	0.070	0.256	0.580	0.882	1.000

For example $P(X=0) = P(\text{all 6 tosses are Tails}) =$
 $(1 - 0.7)^6 = 0.3^6 = 0.001$

Slide 58 STAT.XL.10, UCLA, Ivo Dinov

Binary random process

The *biased-coin tossing model* is a physical model for situations which can be characterized as a series of trials where:

- each trial has only **two outcomes**: *success* or *failure*;
- $p = P(\text{success})$ is the same for every trial; and
- trials are **independent**.

- The distribution of $X =$ number of successes (heads) in N such trials is

$$\text{Binomial}(N, p)$$

Slide 59 STAT XL 10, UCLA, Ivo Dinov

Sampling from a finite population – Binomial Approximation

If we take a sample of size n

- from a much larger population (of size N)
- in which a proportion p have a characteristic of interest, then the distribution of X , the number in the sample with that characteristic,
- is *approximately* Binomial(n, p).
 - (Operating Rule: Approximation is adequate if $n / N < 0.1$.)
- Example, polling the US population to see what proportion is/has-been married.

Slide 60 STAT XL 10, UCLA, Ivo Dinov

Review

- For what types of situation is the urn-sampling model useful? For modeling binary random processes. When sampling **with replacement**, Binomial distribution is exact, where as, in sampling **without replacement** Binomial distribution is an approximation.
- For what types of situation is the biased-coin sampling model useful? Defective parts. Approval poll of cloning for medicinal purposes. Number of Boys in 151 presidential children (90).
- Give the three essential conditions for its applicability. (**two outcomes**; **same p** for every trial; **independence**)

Slide 61 STAT XL 10, UCLA, Ivo Dinov

The Expected value

- The expected value:

$$E(X) = \sum_{\text{all } x} x P(x)$$

- = Sum of (**value times probability of value**)
- = Sum of all Obs's / Total Number

Slide 62 STAT XL 10, UCLA, Ivo Dinov

Example

A couple wants to have children, but they insist on stopping when they have at **least one of each gender** or at **most 3** children example, where $X = \{\text{number of Girls}\}$ we have:

X	0	1	2	3
$\text{pr}(x)$	$\frac{1}{8}$	$\frac{5}{8}$	$\frac{1}{8}$	$\frac{1}{8}$

$$E(X) = \sum_x x P(x)$$

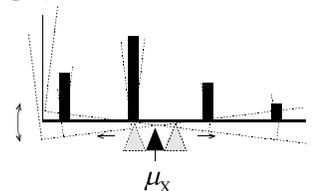
$$= 0 \times \frac{1}{8} + 1 \times \frac{5}{8} + 2 \times \frac{1}{8} + 3 \times \frac{1}{8}$$

$$= 1.25$$

Slide 63 STAT XL 10, UCLA, Ivo Dinov

The expected value as the point of balance

$\text{pr}(x)$



The mean μ_x is the balance point.

Slide 64 STAT XL 10, UCLA, Ivo Dinov

Expected values

- The game of chance: cost to play: \$1.50; Prices {\$1, \$2, \$3}, probabilities of winning each price are {0.6, 0.3, 0.1}, respectively.
- Should we play the game? What are our chances of winning/loosing?

Prize (\$)	x	1	2	3	
Probability	pr(x)	0.6	0.3	0.1	

What we would "expect" from 100 games

Number of games won	0.6 × 100	0.3 × 100	0.1 × 100	add across row
\$ won	1 × 0.6 × 100	2 × 0.3 × 100	3 × 0.1 × 100	Sum

Total prize money = Sum; Average prize money = Sum/100 = 1 × 0.6 + 2 × 0.3 + 3 × 0.1 = 1.5

Theoretically Fair Game: price to play EQ the expected return!

Slide 65 STAT XL 10, UCLA, Ivo Dinov

Binomial Probabilities - ?!?!?!?

- Suppose $X \sim \text{Binomial}(n, p)$, then the probability

$$P(X = x) = \binom{n}{x} p^x (1-p)^{(n-x)}, \quad 0 \leq x \leq n$$
- Where the binomial coefficients are defined by

$$\binom{n}{x} = \frac{n!}{(n-x)! x!}, \quad n! = 1 \times 2 \times 3 \times \dots \times (n-1) \times n$$

n-factorial

Slide 66 STAT XL 10, UCLA, Ivo Dinov

Binomial Formula - examples

- Rolling a pair of dice 4 times. What's the chance of getting at least 2 but less than 4 aces (1-1)?

$$\sum_x P(X = x) = \sum_x \binom{n}{x} p^x (1-p)^{(n-x)}$$

- Let $X = \{\text{Event an ace turns up}\}$, $p = P(\text{success}) = 1/36$

$$P(2 \leq X < 4) = \sum_{x=2}^3 \binom{4}{x} p^x (1-p)^{(4-x)} =$$

$$\binom{4}{2} p^2 (1-p)^2 + \binom{4}{3} p^3 (1-p)^1 =$$

$$6 \times \left(\frac{1}{36}\right)^2 \times \left(\frac{35}{36}\right)^2 + 4 \times \left(\frac{1}{36}\right)^3 \times \left(\frac{35}{36}\right)^1 = 0.005$$

Slide 67 STAT XL 10, UCLA, Ivo Dinov

For the Binomial distribution . . . Mean/SD

$E(X) = np$

$SD(X) = \sqrt{np(1-p)}$

Slide 68 STAT XL 10, UCLA, Ivo Dinov

Example

- Suppose we have a box containing the following numbers in it:

0 0 1 0 1 1 0 1 1

4 0's

5 1's

- What is the probability that if we randomly draw 3 numbers with replacement at least two numbers are 1's?
- $P(\text{getting } 1) = 5/9$. Let $X = \# \text{ 1's in a sample of } 3$
 $\rightarrow X \sim \text{Bin}(3, 0.555)$, $P(X \geq 2) = P(X=2 \text{ or } X=3) = 0.582$.
- What's the probability that we get at most one 1?
 $P(X \leq 1) = P(X=0 \text{ or } X=1) = 0.418$ (Online Res.)

Slide 69 STAT XL 10, UCLA, Ivo Dinov

Example

- According to the data base of a large multinational airline company, 35% of all its 475 pilots are over 40 years of age. The company is about to purchase the latest model Boeing and is planning to select a random sample of 25 pilots to receive training in flying this new plane.

Let X be the number of pilots over 40 years of age in this sample.

- Discuss the validity of using the Binomial distribution in this situation. (Hint: each trial has only two outcomes: $p = P(\text{success})$ is the same for every trial; and trials are independent. If $n/N < 0.1$ binomial approximation is justified! $25/475 \approx 0.05$).

Slide 70 STAT XL 10, UCLA, Ivo Dinov

Examples

- 35% of all its 475 pilots are over 40. The company is to select a random sample of 25 pilots. $X = \#$ pilots over 40 years in sample.

- State the value of the parameter(s) of this distribution.

- Binomial(25, 0.35)

- Assuming that the Binomial distribution you have described above is an appropriate model for X , Find the probability that:

(i) more than 7 of the pilots selected are over 40 years of age.

$P(X > 7) = 1 - P(X \leq 7) = 0.173$, From the online table, but need to know how to compute by hand

(ii) 5 or 6 of the pilots selected are over 40 years of age.

$$P(X=5 \text{ or } X=6) = \binom{25}{5} 0.35^5 \times 0.65^{20} + \binom{25}{6} 0.35^6 \times 0.65^{19}$$

Binomial(25, 0.35)

$$= 0.051 + 0.091 = 0.142$$

(iii) between 13 and 18 (inclusive) of the pilots selected are over 40 years of age. $P(13 \leq X \leq 18) = 0.06$, From the online table.

Slide 71

STAT XL 10, UCLA, Ivo Dinov

Examples

- 35% of all its 475 pilots are over 40 years of age. The company is to select a random sample of 25 pilots. $X = \#$ pilots over 40 years in this sample.

- How many of the pilots selected would you expect to be over 40 years of age? What is the standard deviation of X ?

- $E(X) = n \times p = 25 \times 0.35 = 8.75$

- $SD^2(X) = \text{Var}(X) = n \times p \times (1-p) = 5.6875$

- $SD(X) = 2.385$

- Why would the airline company be interested in the variable "pilots over 40 years of age"? Suggest another variable the company may be interested in measuring? Briefly justify your answer. (GO TO Ch. 21, CI's)

Slide 72

STAT XL 10, UCLA, Ivo Dinov