

## Stat13 Homework 6

[http://www.stat.ucla.edu/~dinov/courses\\_students.html](http://www.stat.ucla.edu/~dinov/courses_students.html)

### Suggested Solutions

7.8:

$$SE_1 = \frac{0.400}{\sqrt{9}} = 0.133, SE_2 = \frac{0.220}{\sqrt{6}} = 0.09$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 0.16$$

7.9

$$SE = \sqrt{SE_1^2 + SE_2^2} = 10.2$$

7.11

$$SE = \sqrt{SE_1^2 + SE_2^2} = 9.192, df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 6$$

$$95\% \text{ CI: } \bar{y}_1 - \bar{y}_2 \pm t_{0.025} \cdot SE(\bar{y}_1 - \bar{y}_2) = (-45.5, -0.5)$$

$$90\% \text{ CI: } \bar{y}_1 - \bar{y}_2 \pm t_{0.05} \cdot SE(\bar{y}_1 - \bar{y}_2) = (-40.9, -5.1)$$

7.12

$$a) SE = \sqrt{SE_1^2 + SE_2^2} = 1.867, df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 190$$

$$95\% \text{ CI: } \bar{y}_1 - \bar{y}_2 \pm t_{0.025} \cdot SE(\bar{y}_1 - \bar{y}_2) = (6.1, 13.5), \text{ (use } df = 140)$$

b) We are 95% confident that the population mean reduction in systolic blood pressure for those who receive training for eight weeks is larger than that for others by an amount that might be as small as 6.1 mmHg or as large as 13.5 mmHg.

7.13

No. The CI found in 7.11 is valid even if the distribution is not normal, because the sample sizes are large.

7.14

$$a) SE = \sqrt{SE_1^2 + SE_2^2} = 4.050, df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 17.2$$

$$90\% \text{ CI: } \bar{y}_1 - \bar{y}_2 \pm t_{0.05} \cdot SE(\bar{y}_1 - \bar{y}_2) = (-5.0, 9.0), \text{ (use } df = 17)$$

b) We are 90% confident that the population mean prothrombin time for rats treated with an antibiotic is smaller than that for control rats by an amount that might be as much as 5 seconds or is larger than that for control rats by an amount that might be as much as 9 seconds.

7.16

$$a) SE = \sqrt{SE_1^2 + SE_2^2} = 0.08763, df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 35.7$$

$$95\% \text{ CI: } \bar{y}_1 - \bar{y}_2 \pm t_{0.025} \cdot SE(\bar{y}_1 - \bar{y}_2) = (-0.12, 0.24), \text{ (use } df = 30)$$

b) We are 95% confident that the population mean head width of all females who mate successfully is smaller than that for rejected females by an amount that might be as much

as 0.12 mm or is larger than that for rejected females by an amount that might be as much as 0.24 mm.

7.17

We are 97.5% confident that the population mean drop in systolic blood pressure of adults placed on a diet rich in fruits and vegetables for eight weeks is larger than that for adults placed on a standard diet by an amount that might be as small as 0.9 mmHg or as large as 4.7 mmHg.

7.19

$$SE = \sqrt{SE_1^2 + SE_2^2} = 5.02, \quad df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 17.3$$

$$90\% \text{ CI: } \bar{y}_1 - \bar{y}_2 \pm t_{0.05} \cdot SE(\bar{y}_1 - \bar{y}_2) = (-7.33, 10.13), \quad (\text{use } df = 17)$$

7.21

$$SE = \sqrt{SE_1^2 + SE_2^2} = 0.509, \quad df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 38$$

$$95\% \text{ CI: } \bar{y}_1 - \bar{y}_2 \pm t_{0.025} \cdot SE(\bar{y}_1 - \bar{y}_2) = (-1.62, 0.46), \quad (\text{use } df = 30)$$

7.27

$$H_0 : \mu_1 = \mu_2$$

$$H_1 : \mu_1 \neq \mu_2$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 4.09, \quad t_{stat} = \frac{\hat{\theta} - \theta_0}{SE(\hat{\theta})} = 1.07, \quad df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 71.9$$

$0.2 < p\text{-value} < 0.4$ . Because  $p\text{-value} > \alpha$ , we do not reject  $H_0$ .

7.30

$$H_0 : \mu_1 = \mu_2$$

$$a) \quad H_1 : \mu_1 \neq \mu_2$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 0.62056, \quad t_{stat} = \frac{\hat{\theta} - \theta_0}{SE(\hat{\theta})} = -3.26,$$

$$df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 94.3$$

$0.001 < p\text{-value} < 0.01$ . Because  $p\text{-value} < \alpha$ , we reject  $H_0$ .

b) There is sufficient evidence ( $0.001 < p\text{-value} < 0.01$ ) to conclude that mean tibia length is larger in females than in males.

c) Judging from the means and SDs, the two distributions overlap substantially, so tibia length would be a poor predictor of sex.

$$d) \quad H_0 : \mu_1 = \mu_2$$

$$H_1 : \mu_1 \neq \mu_2$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 1.962, \quad t_{stat} = \frac{\hat{\theta} - \theta_0}{SE(\hat{\theta})} = -1.03,$$

$$df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 7.8$$

$0.2 < p\text{-value} < 0.4$ . Because  $p\text{-value} > \alpha$ , we do not reject  $H_0$ .

7.31

$$H_0 : \mu_1 = \mu_2$$

$$H_1 : \mu_1 \neq \mu_2$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 5.06, t_{stat} = \frac{\hat{\theta} - \theta_0}{SE(\hat{\theta})} = 0.49, df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 7.7$$

$p$ -value  $> 0.4$ . Because  $p$ -value  $> \alpha$ , we do not reject  $H_0$ .

b) According to the  $p$ -value found in part a), the fact that  $\mu_1$  is larger than  $\mu_2$  could easily be attributed to chance.

7.33

$$a) H_0 : \mu_1 = \mu_2$$

$$H_1 : \mu_1 \neq \mu_2$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 67.08, t_{stat} = \frac{\hat{\theta} - \theta_0}{SE(\hat{\theta})} = 3.73,$$

$$df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 33.6$$

$p$ -value  $< 0.001$ . Because  $p$ -value  $< \alpha$ , we reject  $H_0$ .

b) There is sufficient evidence ( $p$ -value  $< 0.001$ ) to conclude that albumin is more effective than polygelatin as a plasma expander.

7.34

$$H_0 : \mu_1 = \mu_2$$

$$H_1 : \mu_1 \neq \mu_2$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 13.53, t_{stat} = \frac{\hat{\theta} - \theta_0}{SE(\hat{\theta})} = -0.14,$$

$$df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 17.9$$

$p$ -value  $> 0.4$ . Because  $p$ -value  $> \alpha$ , we do not reject  $H_0$ . There is insufficient evidence ( $p$ -value  $> 0.4$ ) to conclude that two diets differ in their effects on cholesterol.

7.35

a) True. We would reject  $H_0$  because  $p$ -value is less than  $\alpha$ .

b) True. We would reject  $H_0$  because  $p$ -value is less than  $\alpha$ .

c) True. This follows directly from the definition of a  $p$ -value.

7.37

$$a) H_0 : \mu_1 = \mu_2$$

$$H_1 : \mu_1 \neq \mu_2$$

$$SE = \sqrt{SE_1^2 + SE_2^2} = 10.21, t_{stat} = \frac{\hat{\theta} - \theta_0}{SE(\hat{\theta})} = 0.92,$$

$$df = \frac{(SE_1^2 + SE_2^2)^2}{SE_1^4/(n_1 - 1) + SE_2^4/(n_2 - 1)} = 10.4$$

$0.2 < p\text{-value} < 0.4$ . Because  $p\text{-value} > \alpha$ , we do not reject  $H_0$ .

b) There is insufficient evidence ( $0.2 < p\text{-value} < 0.4$ ) to conclude that the mean number of colonies differs for control and soap.

7.43

A type II error may be made.

7.45.

Yes. Because 0 is outside of the CI, we know that the p-value is less than 0.05, so the p-value is less than 0.1. Thus we reject the null hypothesis:  $\mu_1 - \mu_2 = 0$ .