

## Stat 13

[http://www.stat.ucla.edu/~dinov/courses\\_students.html](http://www.stat.ucla.edu/~dinov/courses_students.html)

### Chapter 12 Problems/Solutions

All Problems are from: Myra L. Samuels and Jeffrey A. Witmer,  
Statistics for the Life Sciences, 3rd edition, Prentice-Hall (2003)

**12.5.** SOCR (<http://www.socr.ucla.edu>) output is given below.

- (a)  $\text{cob-wt} = 316 - 0.721 \text{ plant-density}$
  - (b) scatterplot (not shown) shows a strong negative linear association between cob-weight (gm grain/cob) and plant density (# plants / pot).
  - (c) as plant density increases by 1 plant per plot, cob weight decreases by 0.72 gm of grain per cob, on average.
  - (d)  $s_Y = \sqrt{11831.8/19} = 25 \text{ gm}$  and  $s_{Y/X} = \sqrt{1337.3/18} = 8.6 \text{ gm}$
  - (e) Predictions of cob weight based on the regression model tend to be off by 8.6 gm on average.
- Equivalently, the data points deviate above or below the regression line by 8.6 gm on average.

#### Regression Analysis

The regression equation is

$$\text{cob-wt} = 316 - 0.721 \text{ plant-density}$$

Predictor	Coef	StDev	T	P
Constant	316.376	8.000	39.55	0.000
plant-de	-0.72063	0.06063	-11.89	0.000

S = 8.619    R-Sq = 88.7%    R-Sq(adj) = 88.1%

#### Analysis of Variance

Source	DF	SS	MS	F	P
Regression	1	10495	10495	141.26	0.000
Residual Error	18	1337	74		
Total	19	11832			

#### 12.6

(a) The slope and intercept of the regression line are

$$b_1 = -927.75/1303 = -.7120;$$

$$b_0 = \bar{y} - b_1\bar{x} = 23.64 - (-.7120)(11.5) = 31.83$$

The fitted regression line is  $\hat{y} = 31.83 - .7120X$

(c)  $s_{Y|X} = \sqrt{SS(\text{resid})/df} = \sqrt{16.7812/(12-2)} = 1.3$

**12.8a:**  $b_1 = 161.40/50667 = .0003186$ ;  $b_0 = .210 - (.0003186)(433.3) = .0720$ .

The fitted regression line is  $Y = .072 - .0003186X$ .

**12.8c:** As altitude of origin goes up by 1 m, respiration rate goes up by .0003186 mul/hr-mg, on average.

**12.8d:**  $s_{\{Y|X\}} = \sqrt{.013986/10} = .0374$

## 12.12

The intercept of the regression line  $b_0$  is based on all 12 data points, not just on the two point for which  $X = 0$ . If there is a linear relationship between  $X$  and  $Y$  (a scatter plot of the data strongly suggest that there is), then the best estimate of the average for  $Y$  at any given value of  $X$  is given by the regression line, taking into account all of the data. In contrast, the average  $(33.3 + 31.0)/2 = 32.15$  ignores most of the data.

**12.14a:** (See Exercise 12.5 for  $b_0$  and  $b_1$ .)

(i) plugging in plant-density = 100 plants gives a predicted cob-wt of  $316 - 0.721(100) = 244.34$  gm

(ii) plugging in plant-density = 120 plants gives a predicted cob-wt of  $316 - 0.721(120) = 229.93$  gm

## 12.14b:

(i)  $(244.34)(100) = 24434$  gm = 24.43 kg

(ii)  $(229.928)(12) = 27591$  gm = 27.6 kg

## 12.15

Using the fitted regression line found in Exercise 12.6 above, we substitute  $X = 15$ . This yields  $\hat{y} = 31.83 - (.7120)(15) = 21.1$ .

Thus, we estimate that the mean fungus growth would be 21.1 mm at a laetiseric acid concentration of 15 microg/ml.

According to the linear model, the standard deviation of fungus growth does not depend on  $X$ . Our estimate of this standard deviation from the regression line is the Residual Standard Deviation  $\sigma_{\{Y|X\}} = \sqrt{[SS(\text{resid})/(n-2)]} = \sqrt{16.7812/10} = 1.3$  mm.

Thus we estimate that the standard deviation of fungus growth would be 1.3 mm at a laetiseric acid concentration of 15 microg/ml.

For  $X = 15$ , we have  $\hat{y} = 21.1 \pm 1.3$  mm.

**12.19a:**  $b_1 = 81.90/2800 = 0.02925$  ng/min (the rate of incorporation)

$b_0 = 0.83 - (0.02925)(30) = -0.05$

$$s_{\{y|x\}} = \sqrt{SS(\text{resid})/(n-2)} = \sqrt{0.035225/5} = 0.0839$$

To construct a 95% confident interval, we consult the z-table (Table 4) with  $df = n-2 = 7-2 = 5$ ; the multiplier is  $t_{\{4,0.025\}} = 2.571$ . The confidence interval is

$$b_1 \pm t_{\{4,0.025\}}SEb_1 = 0.02925 \pm (2.571)(0.00159)$$

$$\text{or } 0.0252 < \beta_1 < 0.033 \text{ ng/min}$$

**12.19ba:** We are 95% confident that the rate at which leucine is incorporated into protein in the population of all *Xenopus* oocytes is between 0.0252 ng/min and 0.0333 ng/min

**12.21a:**  $SEb_1 = 8.6/\sqrt{20209} = 0.0605$ , so 95% CI for  $b_1$  is

$$-0.7206 \pm (2.101)(0.0605) \text{ or } -0.7206 \pm 0.1271 \text{ or } (-0.848, -0.593)$$

**12.21b:** We are 95% confident that as plant density increases by 1 plant per plot, average cob weight decreases by between 0.848 gm and 0.593 gm of grain per cob.

**12.22a:** From Exercise 12.6,  $s_{Y|X} = \sqrt{SS(\text{resid})/df} = \sqrt{16.7812/(12-2)} = 1.3$

The standard error of the slope is

$$SE_{b_1} = s_{Y|X} / \sqrt{\sum(x - \bar{x})^2} = 1.3/\sqrt{1303} = 0.36$$

**12.22b:**  $H_0$ : Laetiseric acid has no effect on fungus growth ( $\beta_1 = 0$ )

HA: Laetiseric acid inhibits fungus growth ( $\beta_1 < 0$ )

$t_s = -0.7120/0.36 = -19.8$ . With  $df = 10$ , the t-table (Table 4) gives  $t_{.0005} = 4.587$ .

Thus the P-value  $< .0005$ , so we reject  $H_0$ . There is sufficient evidence (P-value  $< .0005$ ) to conclude that laetiseric acid inhibits fungus growth.

**12.27a:**  $r = 82.8977/\sqrt{(28465.7)(.363708)} = .8147$

**12.27b:**  $s_Y = \sqrt{(.363708)/(13-1)} = .1741 \text{ gm}$

$$s_{Y|X} = \sqrt{SS(\text{resid})/df} = \sqrt{.1223/(13-2)} = .1054 \text{ gm}$$

$$.1054/.1741 = .605; \sqrt{1 - .8147^2} = .580$$

**12.27c:**  $b_1 = 82.8977/28465.7 = .002912$ ;

$$b_0 = 2.174 - (.002912)(443.8) = .882$$

The fitted regression line is  $\hat{y} = .882 - .002912X$

**12.28a:**  $r = -14563.1/\sqrt{(20209)(11831.8)} = -0.942$

**12.28b:** from Exercise 12.5 (d),  $s_Y = 25 \text{ gm}$  and  $s_{Y|X} = 8.6 \text{ gm}$ , so  $s_{Y|X} / s_Y = 0.344$

further,  $\sqrt{1 - r^2} = 0.3356$ , which is nearly equal to 0.344, so the approximate relationship is indeed verified.

**12.28c:**  $b_1 = -14563.1/20209.0 = -.7206$ ;

$$b_0 = 224.1 - (-.7206)(128.05) = 316.4$$

The fitted regression line is  $Y = 316.4 - .7206X$ .

**12.30:** Let  $X = \text{age}$  and let  $Y = \text{blood pressure}$ . The Residual Standard Deviation is  $s_{\{Y|X\}} = \sqrt{(1 - r^2)(s_Y)^2 / (n-2)} = \sqrt{(1 - .43^2)(19.5)^2 / (2668/2667)} = 17.6 \text{ mm Hg}$ .

$s_{\{Y|X\}} = \sqrt{(y - \hat{y})^2 / (n-2)}$  is a measure of the variability about the regression line  $\hat{y} = b_1x + b_0$ .

But  $s_Y = \sqrt{(y - \bar{y})^2 / (n-1)}$  is a measure of the variability about the mean  $\bar{y}$ .

**12.41a:** with (iii),

**12.41b:** with (ii), and

**12.41c:** with (i).

**12.45a:** The slope and intercept of the regression line are

$$b_1 = -.342 / .1512 = -2.262$$

$$b_0 = 1.117 - (-2.262)(.12) = 1.39$$

The fitted regression line is  $Y = 1.39 - 2.262X$ .

**12.45c:**  $s_{\{Y|X\}} = \sqrt{SS(\text{resid}) / (n-2)} = \sqrt{.2955 / 10} = .1719 \text{ kg}$ .

**12.46a:** If  $x = .24$ , then predicted  $y = 1.39 - 2.262(.24) = .84512$ . But the variability of this prediction is given by  $s_{\{Y|X\}} = .17$ .

If  $x$  is unknown, then the best prediction is  $\bar{y} = 1.117$ , and the precision of this prediction is  $\pm s_Y = .31175$ . We write  $y = 1.117 \pm .31175 \text{ kg}$ .

However, if  $x = .24$  is known, then the best prediction for  $y$  is given by the regression line  $y = 1.39 - 2.262(.24) = .84512$ , but the precision of this prediction is  $\pm s_{\{Y|X\}} = \pm .17$ . We write  $y = .84512 \pm .17 \text{ kg}$ .

**12.46b:** The condition that  $\sigma_{\{Y|X\}}$  does not depend on  $X$  appears to be doubtful. Rather, the scatterplot shows that there is more variability in  $Y$  when  $X$  is small than when  $X$  is large.

X	SD
.00	.21
.06	.28
.12	.11
.30	.06

**12.47:** The hypotheses are

$H_0$ : sulfur dioxide has no effect on yield ( $\beta_1 = 0$ ) and

$H_A$ : Increasing sulfur dioxide tends to decrease yield ( $\beta_1 < 0$ ).

The sample slope is  $b_1 = -.342 / .1512 = -2.262$

We note that  $b_1 < 0$ , so the data do deviate from  $H_0$  in the direction specified by  $H_A$ .

The residual standard deviation is  $s_{\{Y|X\}} = \sqrt{SS(\text{resid})/(n-2)} = \sqrt{.2955/10} = .1719$  kg.

The standard error of the slope is  $SE_{\{b1\}} = .1719/\sqrt{.1512} = .4421$ .

The test statistic is  $t_s = (b1 - 0)/SE_{\{b1\}} = -2.261/.4421 = -5.12$ .

Consulting Table 4 with  $df = n - 2 = 10$ , we find that  $P\text{-value} < .0005$ , so we reject  $H_0$ .

There is strong evidence ( $P\text{-value} < .0005$ ) to conclude that increasing sulfur dioxide tends to decrease yield.

**12.54:**  $SE_{\{b1\}} = s_{\{Y|X\}}/\sqrt{n-1}s_X = .0374/\sqrt{506667} = .0000525$

So 95% CI is  $.0003186 \pm (2.228)(.0000525)$  ( $df=10$ )  
or  $(.00020, .00044)$  or  $.00020 < \beta_1 < .00044$ .

## 12.59

### Regression Analysis

The regression equation is  
water-consumption = 157 - 23.6 dose

Predictor	Coef	StDev	T	P
Constant	156.95	11.87	13.22	0.000
dose	-23.580	7.358	-3.20	0.009

$S = 26.01$        $R\text{-Sq} = 50.7\%$        $R\text{-Sq}(\text{adj}) = 45.7\%$

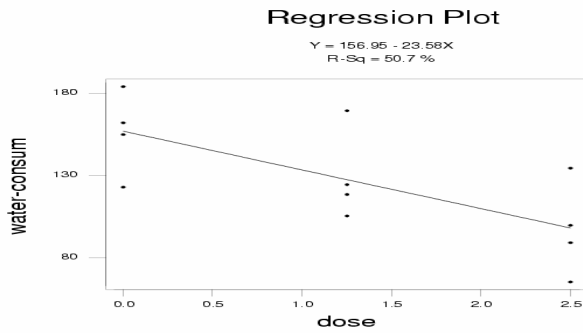
Analysis of Variance

Source	DF	SS	MS	F	P
Regression	1	6950.2	6950.2	10.27	0.009
Residual Error	10	6766.7	676.7		
Total	11	13716.9			

**12.59a:**  $b_0 = 156.95$ ;  $b_1 = -23.580$  (from SOCR, <http://www.socr.ucla.edu> printout)

The fitted regression line is  $\hat{y} = 156.95 - 23.580x$

**12.59b:**



**12.59d:**  $H_0$ : Amphetamine dose has no effect on water consumption ( $\beta_1 = 0$ )  
 $H_A$ : Increasing amphetamine dose tends to reduce water consumption ( $\beta_1 < 0$ )

$t_s = -3.20$  (SOCR, <http://www.socr.ucla.edu> printout) and the P-value =  $0.009/2 = 0.0045$ . Thus we reject  $H_0$ .

There is strong evidence (P-value = 0.0045) to conclude that increasing amphetamine dose tends to reduce water consumption.

**12.59e: One-way Analysis of Variance**

Analysis of Variance for water-co

Source	DF	SS	MS	F	P
dose	2	6972	3486	4.65	0.041
Error	9	6745	749		
Total	11	13717			

Individual 95% CIs For Mean  
Based on Pooled StDev

Level	N	Mean	StDev
0.00	4	156.00	25.32
1.25	4	129.38	27.85
2.50	4	97.05	28.84

Pooled StDev = 27.38

70      105      140      175

H0: The three doses produce the same mean water consumption level ( $\mu_1 = \mu_2 = \mu_3$ )

HA: The mean water consumption levels are not all equal (the  $\mu$ 's are not all equal)

$F_s = 4.65$  (SOCR, <http://www.socr.ucla.edu> printout) and the P-value = 0.041. Note HA cannot be directional because there are three doses). Thus we reject H0.

The conclusion here is similar to that in part (d), in that we reject H0. However, the analysis from (d) gave a smaller P-value, as it made use of the fact that the means are not only different, but they decrease as dose increases.

**12.59f:** The analysis in part (d) requires linearity; that is, the mean water consumption levels must have a linear relationship to dose for the regression model to make sense. The ANOVA in part (e) does not require this condition.

**12.59g:**  $s_{\text{pooled}} = 27.38$  (from ANOVA printout), which is similar to  $s_{\{Y|X\}} = 26.01$  (from regression printout)