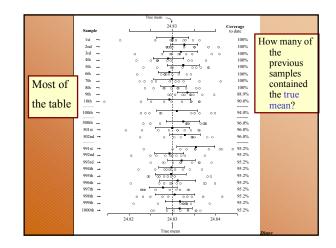
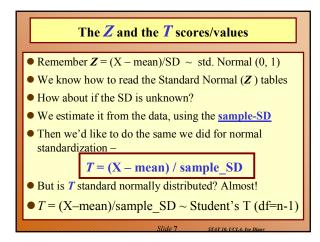
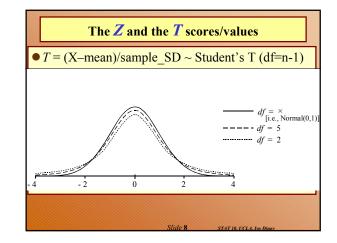


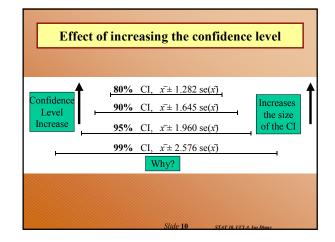
 CI are constructed using the sample read s=SE. But diff. samples yield diff. estimates and → diff. CI's?!? Below is a <u>computer simulation</u> showing how process of taking samples effects the estimates and the CI's. • 1000 samples of size 10 obs's from a Normal(m=24.83, s=0.005) distributions with their 95% CI's. True mean $\overline{x} \pm t SE(\overline{x}) = \overline{x} \pm 2.262SE(\overline{x})$ 24 83 Coverage to date ample 1st 00 100% 0 0 2nd ----ထားစု 100% C 3rd 100% ----0 0 0 999th 95.2% 0 0 1000th Ł 95.29 0. 00 24.82 24.83 24.84 True mean almost always captured in the CI. True mear

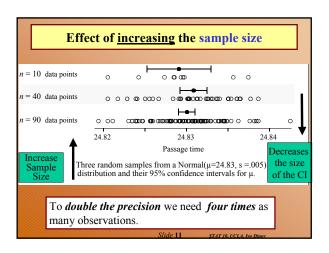


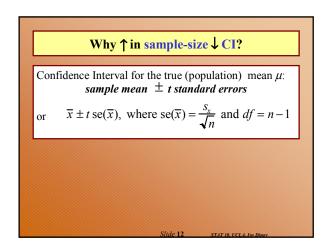


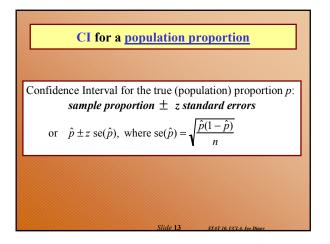


		Sun	nma	ry - (CI fo	r <u>po</u>	pulat	tion	mear	<u>1</u>	
	Confidence Interval for the true (population) mean μ : sample mean $\pm t$ standard errors										
	or	$\overline{x} \pm $	$t \operatorname{se}(\overline{x})$;), w	here	$se(\overline{x})$	$=\frac{S_x}{\sqrt{n}}$, and	l <i>df</i> =	• <i>n</i> − 1	
TAI	BLE 8.1.	1 Value	of the	Multipli	ier, <i>t</i> , fo	or a 95%	6 CI				
df	. 7	8	9	10	11	12	13	14	15	16	17
t	2.365		2.262	2.228	2.201	2.179	2.160	2.145	2.131	2.120	2.110
df	: 18	19	20	25	30	35 2.030	40	45 2.014	50 2.009	60 2.000	∞ 1.960
I	2.101	2.093	2.086	2.060	2.042	2.030	2.021	2.014	2.009	2.000	1.960
						Slide 9			I A Jue Die		

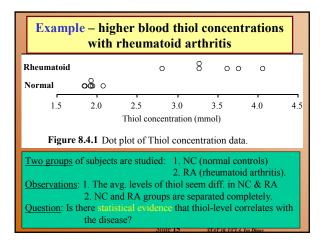


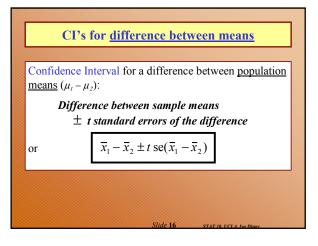


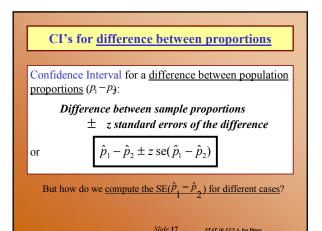


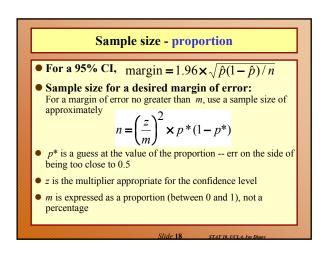


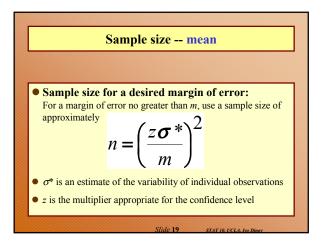
Example – higher blood thiol concentration with rheumatoid arthritis Thiol Concentration (mmol)					
	Normal	Rheumatoid			
Research question:	1.84	2.81			
Is the change in the Thiol status	1.92	4.06			
in the lysate of packed blood	1.94	3.62			
cells substantial to be indicative	1.92	3.27			
of a non trivial relationship	1.85	3.27			
between Thiol-levels and	1.91	3.76			
rheumatoid arthritis?	2.07				
Sample size	7	6			
Samp le mean	1.92143	3.46500			
Sample standard deviation	0.07559	0.44049			
Slie	le 14	STAT 10. UCLA. Ivo Dinov			

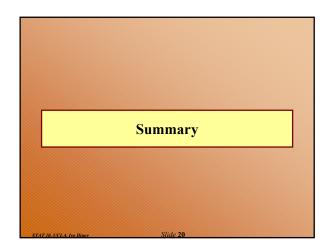


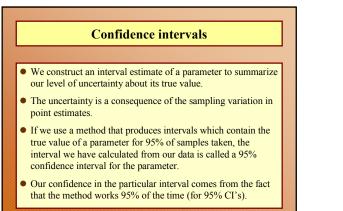












Parameter		Estimate	Standard error of estimate	df
M ean,	μ	\overline{x}	$\frac{s_x}{\sqrt{n}}$	n-1
Proportion,	р	\hat{p}	$\sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$	00
Difference in means,	μ ₁ -μ ₂	$\overline{x}_1 - \overline{x}_2$	$\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$	$Min(n_1-1, n_2-1)$
Difference in proportions,	<i>P</i> ₁ - <i>P</i> ₂	$\hat{p}_1 - \hat{p}_2$	(see Table 8.5.5)	~

Summary cont.

• For a great many situations,

an (approximate) confidence interval is given by

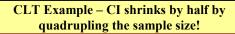
estimate \pm t standard errors

The size of the multiplier, t, depends both on the desired confidence level and the degrees of freedom (df).

- [With proportions, we use the Normal distribution (i.e., $df = \infty$) and it is conventional to use *z* rather than *t* to denote the multiplier.]
- The *margin of error* is the quantity added to and subtracted from the estimate to construct the interval (i.e. *t* standard errors).

Summary cont.

- If we want greater confidence that an interval calculated from our data will contain the true value, we have to use a wider interval.
- To double the precision of a 95% confidence interval (i.e.halve the width of the confidence interval), we need to take 4 times as many observations.



- If I ask 30 of you the question "Is 5 credit hour a reasonable load for Stat10?", and say, 15 (50%) said *no*. Should we change the format of the class?
- Not really the 2SE interval is about [0.32; 0.68]. So, we • Not really – the 2SE interval is about [0.32; 0.68]. So, we have little concrete evidence of the proportion of students who think we need a change in Stat 10 format, $\hat{p} \pm 2 \times SE(\hat{p}) = 0.5 \pm 2 \times \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = 0.5 \pm -0.18$ • If I ask all 300 Stat 10 students and 150 say *no* (still 50%), then 2SE interval around 50% is: [0.44; 0.56].

- So, large sample is much more useful and this is due to CLT effects, without which, we have no clue how useful our estimate actually is ...

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