

# **Stat 100a: Introduction to Probability.**

## Outline for the day

1. Exam 2.
2. Random walks.
3. Reflection principle.
4. Ballot theorem.
5. Avoiding zero.
6. Chip proportions and induction.
7. Doubling up.
8. Examples.

The computer project is due on Sat Dec2 8:00pm.

HW3 is due Tue Dec 5.

Thu Dec 7 is the final exam, here in class, 11am to 12:15pm.

Again any notes and books are fine, and bring a pencil and a calculator. Also bring your student ID to the exam.

## 1. Midterm 2.

The scores are on the course website, <http://www.stat.ucla.edu/~frederic/100a/F17> in 100aexam2scores.txt .

$P(\text{straight flush on the turn}) = (4 \cdot 9 \cdot 46 + 4 \cdot 47) / C(52, 6) = 0.0000906$ ,  
not 0.000906%. I accepted B or E.

It is easy to doublecount 234567 all of hearts. We will count it only in 34567x. Count all straight flushes but not royal flushes, like 23456x, where x cannot be the 7 of hearts. So there are 46 possibilities for x. Similarly in counting 34567x, x cannot be the 8 of hearts but it can be the 2 of hearts. There are  $4 \cdot 9 \cdot 46$  such straight flushes. But with 10JQKAx, now x can truly be any of the remaining 47 cards.

If you play 400,000 hands, you expect 36.2 straight flushes on the turn, not 3.62. Again I accepted C or E.

$P(\text{straight flush on the turn but not royal flush}) = (4 \cdot 9 \cdot 46) / C(52, 6) = 0.0000813$ .

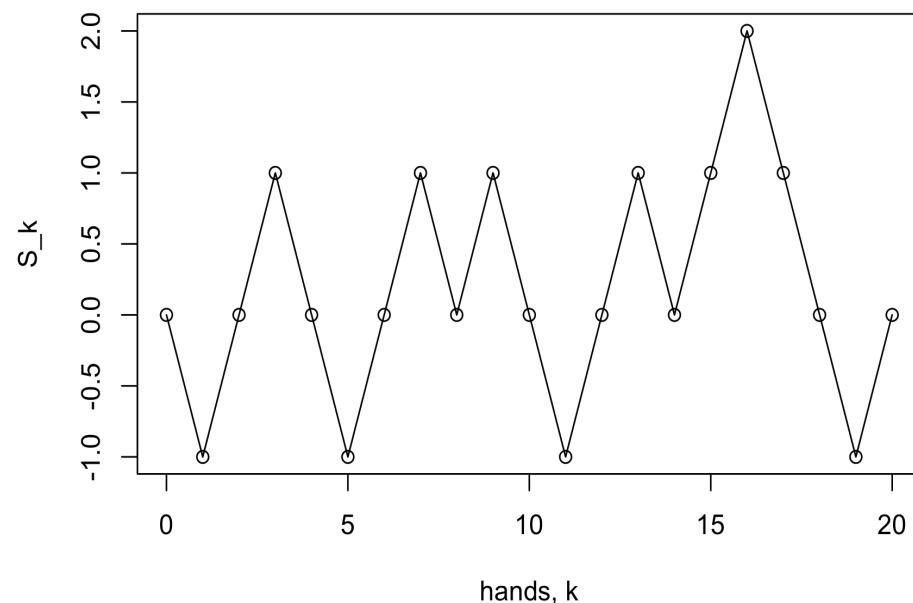
## 2. Random walks, ch. 7.6.

Suppose that  $X_1, X_2, \dots$ , are iid,

and  $S_k = X_0 + X_1 + \dots + X_k$  for  $k = 0, 1, 2, \dots$

The totals  $\{S_0, S_1, S_2, \dots\}$  form a random walk.

The classical (*simple*) case is when each  $X_i$  is 1 or -1 with probability  $\frac{1}{2}$  each.



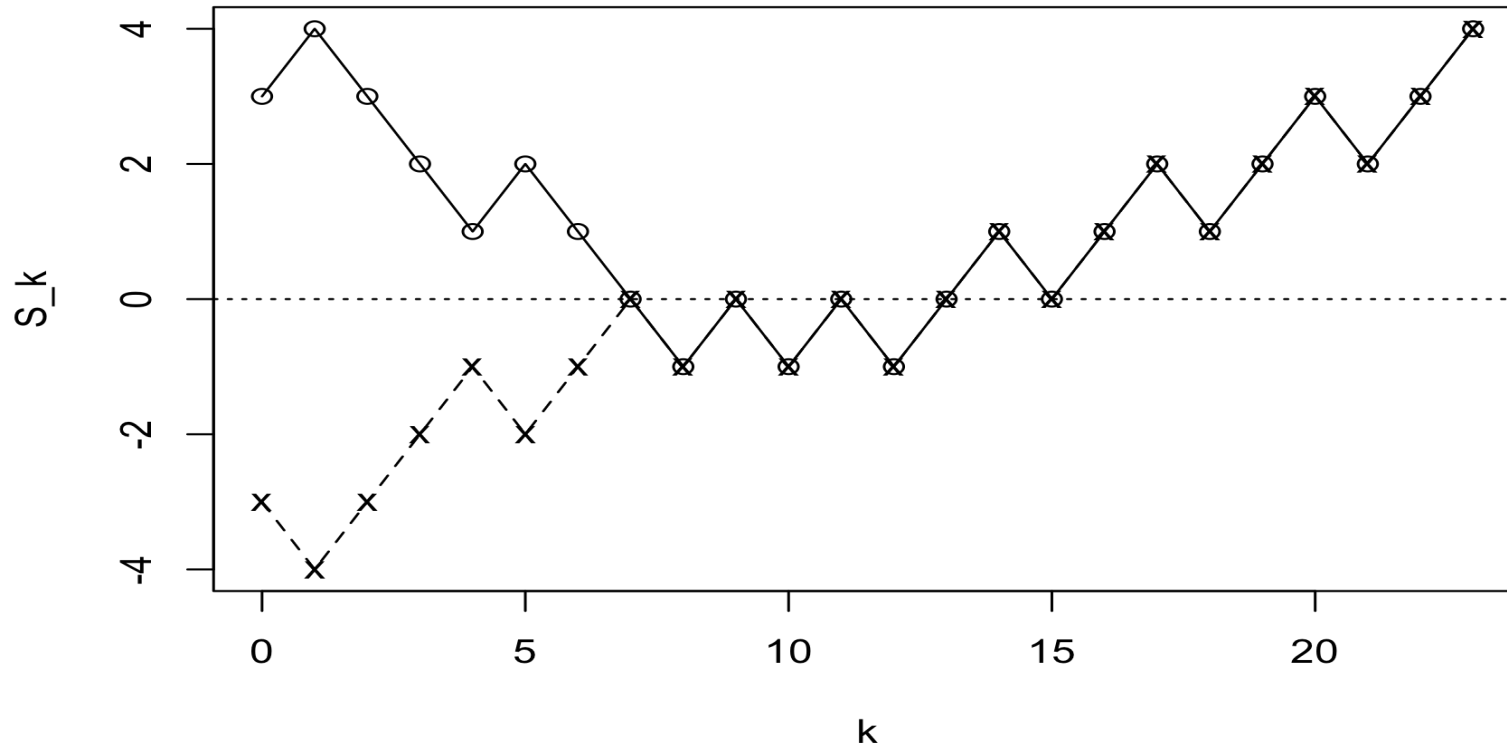
\* Reflection principle: The number of paths from  $(0, X_0)$  to  $(n, y)$  that touch the x-axis = the number of paths from  $(0, -X_0)$  to  $(n, y)$ , for any  $n, y$ , and  $X_0 > 0$ .

\* Ballot theorem: In  $n = a+b$  hands, if player A won  $a$  hands and B won  $b$  hands, where  $a > b$ , and if the hands are aired in random order,  $P(\text{A won more hands than B throughout the telecast}) = (a-b)/n$ .

[In an election, if candidate X gets  $x$  votes, and candidate Y gets  $y$  votes, where  $x > y$ , then the probability that X always leads Y throughout the counting is  $(x-y) / (x+y)$ .]

\* For a simple random walk,  $P(S_1 \neq 0, S_2 \neq 0, \dots, S_n \neq 0) = P(S_n = 0)$ , for any even  $n$ .

**3. Reflection Principle.** The number of paths from  $(0, X_0)$  to  $(n, y)$  that touch the x-axis  
 = the number of paths from  $(0, -X_0)$  to  $(n, y)$ , for any  $n, y$ , and  $X_0 > 0$ .



For each path from  $(0, X_0)$  to  $(n, y)$  that touches the x-axis, you can reflect the first part  
 til it touches the x-axis, to find a path from  $(0, -X_0)$  to  $(n, y)$ , and vice versa.

Total number of paths from  $(0, -X_0)$  to  $(n, y)$  is easy to count: it's just  $C(n, a)$ , where you  
 go up  $a$  times and down  $b$  times

[i.e.  $a - b = y - (-X_0) = y + X_0$ .  $a + b = n$ , so  $b = n - a$ ,  $2a - n = y + X_0$ ,  $a = (n + y + X_0)/2$ ].

**4. Ballot theorem.** In  $n = a+b$  hands, if player A won  $a$  hands and B won  $b$  hands, where  $a > b$ , and if the hands are aired in random order, then  $P(\text{A won more hands than B throughout the telecast}) = (a-b)/n$ .

Proof: We know that, after  $n = a+b$  hands, the total difference in hands won is  $a-b$ .

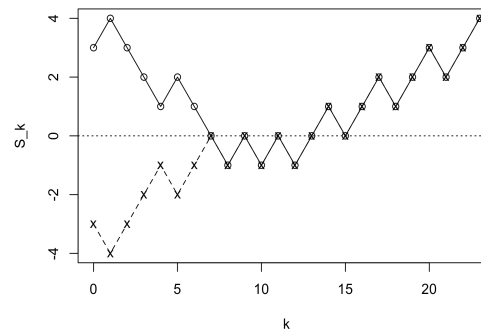
Let  $x = a-b$ .

We want to count the number of paths from  $(1,1)$  to  $(n,x)$  that do not touch the  $x$ -axis.

By the reflection principle, the number of paths from  $(1,1)$  to  $(n,x)$  that **do** touch the  $x$ -axis equals the total number of paths from  $(1,-1)$  to  $(n,x)$ .

So the number of paths from  $(1,1)$  to  $(n,x)$  that **do not** touch the  $x$ -axis equals the number of paths from  $(1,1)$  to  $(n,x)$  minus the number of paths from  $(1,-1)$  to  $(n,x)$

$$\begin{aligned}
 &= C(n-1, a-1) - C(n-1, a) \\
 &= (n-1)! / [(a-1)! (n-a)!] - (n-1)! / [a! (n-a-1)!] \\
 &= \{n! / [a! (n-a)!]\} [(a/n) - (n-a)/n] \\
 &= C(n, a) (a-b)/n.
 \end{aligned}$$



And each path is equally likely, and has probability  $1/C(n,a)$ .

So,  $P(\text{going from } (0,0) \text{ to } (n,x) \text{ without touching the } x\text{-axis}) = (a-b)/n$ .

## 5. Avoiding zero.

For a simple random walk, for any even #  $n$ ,  $P(S_1 \neq 0, S_2 \neq 0, \dots, S_n \neq 0) = P(S_n = 0)$ .

Proof. The number of paths from  $(0,0)$  to  $(n,j)$  that don't touch the x-axis at positive times  
= the number of paths from  $(1,1)$  to  $(n,j)$  that don't touch the x-axis at positive times  
= paths from  $(1,1)$  to  $(n,j)$  - paths from  $(1,-1)$  to  $(n,j)$  by the *reflection principle*  
=  $N_{n-1,j-1} - N_{n-1,j+1}$

Let  $Q_{n,j} = P(S_n = j)$ .

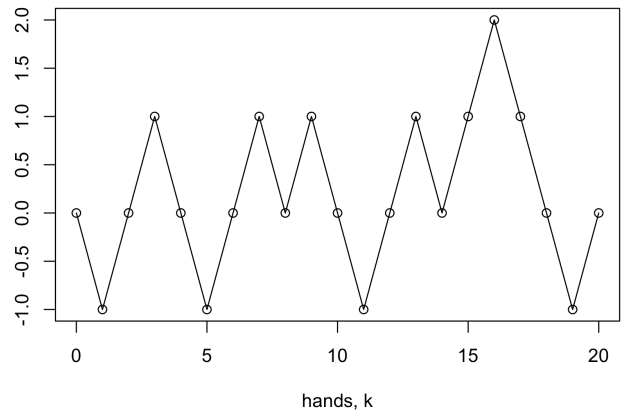
$$P(S_1 > 0, S_2 > 0, \dots, S_{n-1} > 0, S_n = j) = \frac{1}{2}[Q_{n-1,j-1} - Q_{n-1,j+1}].$$

Summing from  $j = 2$  to  $\infty$ ,

$$\begin{aligned} P(S_1 > 0, S_2 > 0, \dots, S_{n-1} > 0, S_n > 0) \\ &= \frac{1}{2}[Q_{n-1,1} - Q_{n-1,3}] + \frac{1}{2}[Q_{n-1,3} - Q_{n-1,5}] + \frac{1}{2}[Q_{n-1,5} - Q_{n-1,7}] + \dots \\ &= (1/2) Q_{n-1,1} \\ &= (1/2) P(S_n = 0), \text{ because to end up at } (n, 0), \text{ you have to be at } (n-1, +/-1), \\ &\text{ so } P(S_n = 0) = (1/2) Q_{n-1,1} + (1/2) Q_{n-1,-1} = Q_{n-1,1}. \end{aligned}$$

By the same argument,  $P(S_1 < 0, S_2 < 0, \dots, S_{n-1} < 0, S_n < 0) = (1/2) P(S_n = 0)$ .

So,  $P(S_1 \neq 0, S_2 \neq 0, \dots, S_n \neq 0) = P(S_n = 0)$ .



## **6. Chip proportions and induction, Theorem 7.6.6.**

$P(\text{win a tournament})$  is proportional to your number of chips.

Simplified scenario. Suppose you either go up or down 1 each hand, with prob.  $1/2$ .

Suppose there are  $n$  chips, and you have  $k$  of them.

Let  $p_k = P(\text{win tournament given } k \text{ chips}) = P(\text{random walk goes } k \rightarrow n \text{ before hitting } 0)$ .

Now, clearly  $p_0 = 0$ . Consider  $p_1$ . From 1, you will either go to 0 or 2.

So,  $p_1 = 1/2 p_0 + 1/2 p_2 = 1/2 p_2$ . That is,  $p_2 = 2 p_1$ .

We have shown that  $p_j = j p_1$ , for  $j = 0, 1$ , and  $2$ .

**(induction:)** Suppose that, for  $j = 0, 1, 2, \dots, m$ ,  $p_j = j p_1$ .

**We will show that  $p_{m+1} = (m+1) p_1$ .**

**Therefore,  $p_j = j p_1$  for all  $j$ .**

That is,  $P(\text{win the tournament})$  is prop. to your number of chips.

$p_m = 1/2 p_{m-1} + 1/2 p_{m+1}$ . If  $p_j = j p_1$  for  $j \leq m$ , then we have

$$m p_1 = 1/2 (m-1) p_1 + 1/2 p_{m+1},$$

$$\text{so } p_{m+1} = 2m p_1 - (m-1) p_1 = (m+1) p_1.$$

**7. Doubling up.** Again,  $P(\text{winning}) = \text{your proportion of chips}$ .

Theorem 7.6.7, p152, describes another simplified scenario.

Suppose you either double each hand you play, or go to zero, each with probability  $1/2$ .

Again,  $P(\text{win a tournament})$  is prop. to your number of chips.

Again,  $p_0 = 0$ , and  $p_1 = 1/2$   $p_2 = 1/2$   $p_2$ , so again,  $p_2 = 2 p_1$ .

We have shown that, for  $j = 0, 1$ , and  $2$ ,  $p_j = j p_1$ .

**(induction:)** Suppose that, for  $j \leq m$ ,  $p_j = j p_1$ .

**We will show that  $p_{2m} = (2m) p_1$ .**

**Therefore,  $p_j = j p_1$  for all  $j = 2^k$ .** That is,  $P(\text{win the tournament})$  is prop. to # of chips.

This time,  $p_m = 1/2 p_0 + 1/2 p_{2m}$ . If  $p_j = j p_1$  for  $j \leq m$ , then we have

$mp_1 = 0 + 1/2 p_{2m}$ , so  $p_{2m} = 2mp_1$ . Done.

In Theorem 7.6.8, p152, you have  $k$  of the  $n$  chips in play. Each hand, you gain 1 with prob.  $p$ , or lose 1 with prob.  $q=1-p$ .

Suppose  $0 < p < 1$  and  $p \neq 0.5$ . Let  $r = q/p$ . Then  $P(\text{you win the tournament}) = (1-r^k)/(1-r^n)$ .

The proof is again by induction, and is similar to the proof we did of Theorem 7.6.6.



## 8. Examples.

(Chen and Ankenman, 2006). Suppose that a \$100 winner-take-all tournament has  $1024 = 2^{10}$  players. So, you need to double up 10 times to win. Winner gets \$102,400.

Suppose you have probability  $p = 0.54$  to double up, instead of 0.5.

What is your expected profit in the tournament? (Assume only doubling up.)

Answer.  $P(\text{winning}) = 0.54^{10}$ , so exp. return =  $0.54^{10} (\$102,400) = \$215.89$ . So exp. profit = \$115.89.

What if each player starts with 10 chips, and you gain a chip with  $p = 54\%$  and lose a chip with  $p = 46\%$ ? What is your expected profit?

Answer.  $r = q/p = .46/.54 = .852$ .  $P(\text{you win}) = (1-r^{10})/(1-r^{10240}) = 79.9\%$ .  
So exp. profit =  $.799(\$102400) - \$100 \sim \$81700$ .

## Random Walk example.

Suppose you start with 1 chip at time 0 and that your tournament is like a simple random walk, but if you hit 0 you are done.  $P(\text{you have not hit zero by time } 47)?$

We know that starting at 0,  $P(Y_1 \neq 0, Y_2 \neq 0, \dots, Y_{2n} \neq 0) = P(Y_{2n} = 0)$ .

So, for a random walk starting at (0,0),

by symmetry  $P(Y_1 > 0, Y_2 > 0, \dots, Y_{48} > 0) = \frac{1}{2} P(Y_1 \neq 0, Y_2 \neq 0, \dots, Y_{2n} \neq 0)$   
 $= \frac{1}{2} P(Y_{48} = 0) = \frac{1}{2} \text{Choose}(48,24)(\frac{1}{2})^{48}.$

Also  $P(Y_1 > 0, Y_2 > 0, \dots, Y_{48} > 0) = P(Y_1 = 1, Y_2 > 0, \dots, Y_{48} > 0)$   
 $= P(\text{start at 0 and win your first hand, and then stay above 0 for at least 47 more hands})$   
 $= P(\text{start at 0 and win your first hand}) \times P(\text{from (1,1), stay above 0 for } \geq 47 \text{ more hands})$   
 $= \frac{1}{2} P(\text{starting with 1 chip, stay above 0 for at least 47 more hands}).$

So, multiplying both sides by 2,

$P(\text{starting with 1 chip, stay above 0 for at least 47 hands}) = \text{Choose}(48,24)(\frac{1}{2})^{48}$   
 $= 11.46\%.$