Stat 13, Intro. to Statistical Methods for the Life and Health Sciences.

- 1. t versus normal and assumptions.
- 2. Confounding and lefties example.
- 3. Comparing two proportions using numerical and visual summaries, good or bad year example.
- 4. Comparing 2 proportions with CIs + testing using simulation, dolphin example.
- 5. Comparing 2 props. with theory-based testing, smoking and gender example.

Read ch5.

http://www.stat.ucla.edu/~frederic/13/F17.

Bring a PENCIL and CALCULATOR and any books or notes you want to the exams.

The midterm is Tue Nov 7.

There is no lecture Thu Nov 2 because of faculty retreat and also no lecture Thu Nov 9.

The first exam will be on everything through today. We will do review next class.

Why do we sometimes use the t distribution and sometimes the normal distribution in testing and confidence intervals?

The central limit theorem states that, for any iid random variables X_1 , ..., X_n with mean μ and SD σ , $(\bar{x} - \mu) \div (\sigma/vn)$ -> standard normal, as $n \to \infty$.

iid means independent and identically distributed, like draws from the same large population. standard means mean 0 and SD 1.

CLT: $(\bar{x} - \mu) \div (\sigma/\sqrt{n})$ -> standard normal. If Z is std. normal, then P(|Z| < 1.96) = 95%.

So, if n is large, then

$$P(|(\bar{x} - \mu) \div (\sigma/vn)| < 1.96) \sim 95\%.$$

Mult. by (σ/vn) and get

$$P(|\bar{x} - \mu| < 1.96 \sigma/vn) \sim 95\%$$
.

 $P(\mu - \bar{x} \text{ is in the range } 0 +/- 1.96 \sigma/vn) \sim 95\%.$

P(μ is in the range \bar{x} +/- 1.96 σ / ν n) ~ 95%.

This all assumes n is large. What if n is small?

CLT: $(\bar{x} - \mu) \div (\sigma/\sqrt{n})$ -> standard normal.

What about if n is small?

A property of the normal distribution is that the sum of independent normals is also normal, and from this it follows that if $X_1, ..., X_n$ are iid and normal, then $(\bar{x} - \mu) \div (\sigma/vn)$ is standard normal.

So again P(μ is in the range \bar{x} +/- 1.96 σ / ν n) = 95%. This assumes you know σ . What if σ is unknown?

Suppose $X_1, ..., X_n$ are iid with mean μ and SD σ .

CLT: $(\bar{x} - \mu) \div (\sigma/\sqrt{n}) \sim \text{std. normal.}$

If $X_1, ..., X_n$ are normal, then $(\bar{x} - \mu) \div (\sigma/v)$ is std. normal.

 σ is the SD of the population from which $X_1, ..., X_n$ are drawn. s is the SD of the sample, $X_1, ..., X_n$.

Gosset/student (1908) showed that replacing σ with s, if $X_1, ..., X_n$ are normal, then $(\bar{x} - \mu) \div (s/vn)$ is t distributed. So we need the multiplier from the t distribution.

To sum up,

if the observations are iid and n is large, then

P(μ is in the range \bar{x} +/- 1.96 σ / ν n) ~ 95%.

If the observations are iid and normal, then

P(μ is in the range \bar{x} +/- 1.96 σ / ν n) ~ 95%.

If the obs. are iid and normal and σ is unknown, then

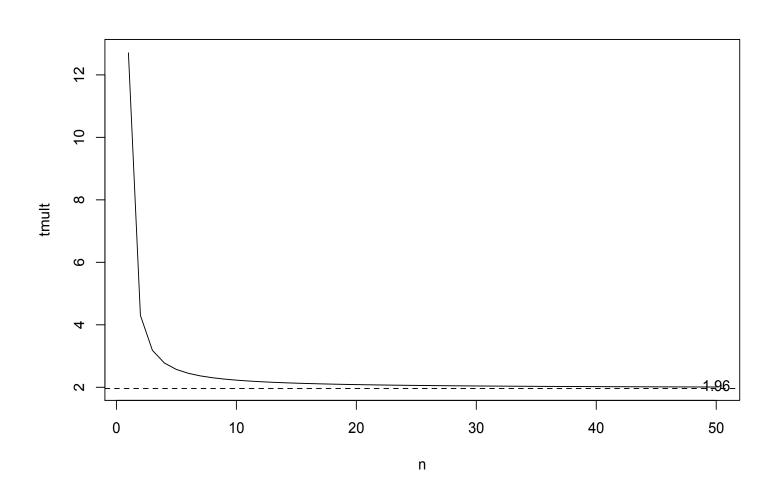
P(μ is in the range \bar{x} +/- t_{mult} s/vn) ~ 95%.

where t_{mult} is the multiplier from the t distribution.

This multiplier depends on n.

If n is large (like 30 or more) then t_{mult} is close to 1.96 anyway.

If obs. are iid, n is not large, and the obs. are not normal? Then you need to do simulations.



2. More about confounding factors and lefties example.

- A confounding factor must be plausibly linked to both the explanatory and response variables. So for instance saying "perhaps a higher proportion of the smokers are men" would not be a very convincing confounding factor, unless you have some reason to think gender is strongly linked to liver cancer.
- Another example: left-handedness and age at death.
 Psychologists Diane Halpern and Stanley Coren looked at 1,000 death records of those who died in Southern California in the late 1980s and early 1990s and contacted relatives to see if the deceased were righthanded or lefthanded. They found that the average ages at death of the lefthanded was 66, and for the righthanded it was 75. Their results were published in prestigious scientific journals, Nature and the New England Journal of Medicine.

All sorts of causal conclusions were made about how this shows that the stress of being lefthanded in our righthanded world leads to premature death.

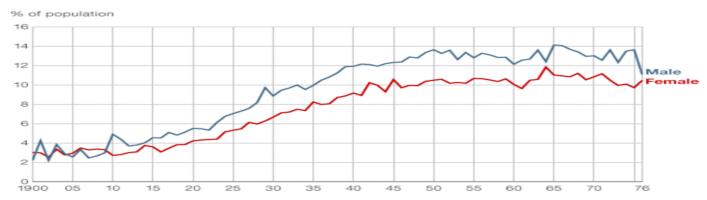


Is this an observational study or an experiment?

- Is this an observational study or an experiment? It is an observational study.
- Are there plausible confounding factors you can think of?

 A confounding factor is the age of the two populations in general. Lefties in the 1980s were on average younger than righties. Many old lefties were converted to righties at infancy, in the early 20th century, but this practice has subsided. Thus in the 1980s and 1990s, there were relatively few old lefties but many young lefties in the overall population. This alone explains the discrepancy.

Left handedness 1900-1976



Source: Chris McManus Right Hand, Left Hand

Unit 2: Comparing Two Groups

- In Unit 1, we learned the basic process of statistical inference using tests and confidence intervals. We did all this by focusing on a single proportion.
- In Unit 2, we will take these ideas and extend them to comparing two groups. We will compare two proportions, two independent means, and paired data.

Chapter 5: Comparing Two Proportions

- 5.1: Comparing Two Groups: Categorical Response
- 5.2: Comparing Two Proportions: Simulation-Based Approach
- 5.3: Comparing Two Proportions: Theory-Based Approach

3. Comparing two proportions using numerical and visual summaries, and the good or bad year example.

Section 5.1

Example 5.1: Positive and Negative Perceptions

- Consider these two questions:
 - Are you having a good year?
 - Are you having a bad year?

• Do people answer each question in such a way that would indicate the same answer? (e.g. Yes for the first one and No for the second.)

Positive and Negative Perceptions

- Researchers questioned 30 students (randomly giving them one of the two questions).
- They then recorded if a positive or negative response was given.
- They wanted to see if the wording of the question influenced the answers.

Positive and negative perceptions

- Observational units
 - The 30 students
- Variables
 - Question wording (good year or bad year)
 - Perception of their year (positive or negative)
- Which is the explanatory variable and which is the response variable?
- Is this an observational study or experiment?

Raw Data in a Spreadsheet

Individual	Type of Question	Response
1	Good Year	Positive
2	Good Year	Negative
3	Bad Year	Positive
4	Good Year	Positive
5	Good Year	Negative
6	Bad Year	Positive
7	Good Year	Positive
8	Good Year	Positive
9	Good Year	Positive
10	Bad Year	Negative
11	Good Year	Negative
12	Bad Year	Negative
13	Good Year	Positive
14	Bad Year	Negative
15	Good Year	Positive

Individual	Type of Question	Response
16	Good Year	Positive
17	Bad Year	Positive
18	Good Year	Positive
19	Good Year	Positive
20	Good Year	Positive
21	Bad Year	Negative
22	Good Year	Positive
23	Bad Year	Negative
24	Good Year	Positive
25	Bad Year	Negative
26	Good Year	Positive
27	Bad Year	Negative
28	Good Year	Positive
29	Bad Year	Positive
30	Bad Year	Negative

Two-Way Tables

- A two-way table organizes data
 - Summarizes two categorical variables
 - Also called contingency table
- Are students more likely to give a positive response if they were given the good year question?

	Good Year	Bad Year	Total
Positive response	15	4	19
Negative response	3	8	11
Total	18	12	30

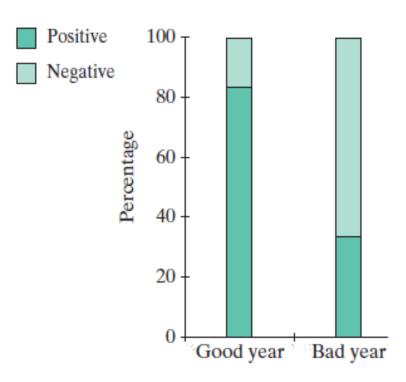
Two-Way Tables

- Conditional proportions will help us better determine if there is an association between the question asked and the type of response.
- We can see that the subjects with the positive question were *more likely* to respond positively.

	Good Year	Bad Year	Total
Positive response	15/18 ≈ 0.83	4/12 ≈ 0.33	19
Negative response	3	8	11
Total	18	12	30

Segmented Bar Graphs

 We can also use segmented bar graphs to see this association between the "good year" question and a positive response.



Statistic

 The statistic we will mainly use to summarize this table is the difference in proportions of positive responses is 0.83 – 0.33 = 0.50.

	Good Year	Bad Year	Total
Positive response	15 (83%)	4 (33%)	19
Negative response	3	8	11
Total	18	12	30

Another Statistic

- Another statistic that is often used, called relative risk, is the ratio of the proportions: 0.83/0.33 = 2.5.
- We can say that those who were given the good year question were 2.5 times as likely to give a positive response.

	Good Year	Bad Year	Total
Positive response	15 (83%)	4 (33%)	19
Negative response	3	8	11
Total	18	12	30

4. Comparing 2 proportions with CIs and testing using simulation, dolphin example.

Section 5.2

Example 5.2

Is swimming with dolphins therapeutic for patients suffering from clinical depression?

- Researchers Antonioli and Reveley (2005), in British Medical Journal, recruited 30 subjects aged 18-65 with a clinical diagnosis of mild to moderate depression
- Discontinued antidepressants and psychotherapy 4 weeks prior to and throughout the experiment
- 30 subjects went to an island near Honduras where they were randomly assigned to two treatment groups

- Both groups engaged in one hour of swimming and snorkeling each day
- One group swam in the presence of dolphins and the other group did not
- Participants in both groups had identical conditions except for the dolphins
- After two weeks, each subjects' level of depression was evaluated, as it had been at the beginning of the study
- The response variable is whether or not the subject achieved substantial reduction in depression

Null hypothesis: Dolphins do not help.

 Swimming with dolphins is not associated with substantial improvement in depression

Alternative hypothesis: Dolphins help.

 Swimming with dolphins increases the probability of substantial improvement in depression symptoms

- The parameter is the (long-run) difference between the probability of improving when receiving dolphin therapy and the prob. of improving with the control ($\pi_{\text{dolphins}} \pi_{\text{control}}$)
- So we can write our hypotheses as:

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\begin{aligned} &\mathbf{H_0:} \ \pi_{\text{dolphins}} - \pi_{\text{control}} = 0. \\ &\mathbf{H_a:} \ \pi_{\text{dolphins}} - \pi_{\text{control}} > 0. \\ &\mathbf{or} \\ &\mathbf{H_0:} \ \pi_{\text{dolphins}} = \pi_{\text{control}} \\ &\mathbf{H_a:} \ \pi_{\text{dolphins}} > \pi_{\text{control}} \end{aligned}
```

(Note: we are not saying our parameters equal any certain number.)

Results:

	Dolphin group	Control group	Total
Improved	10 (66.7%)	3 (20%)	13
Did Not Improve	5	12	17
Total	15	15	30

The difference in proportions of improvers is:

$$\hat{p}_d - \hat{p}_c = 0.667 - 0.20 = 0.467.$$

- There are two possible explanations for an observed difference of 0.467.
 - A tendency to be more likely to improve with dolphins (alternative hypothesis)
 - The 13 subjects were going to show improvement with or without dolphins and random chance assigned more improvers to the dolphins (null hypothesis)

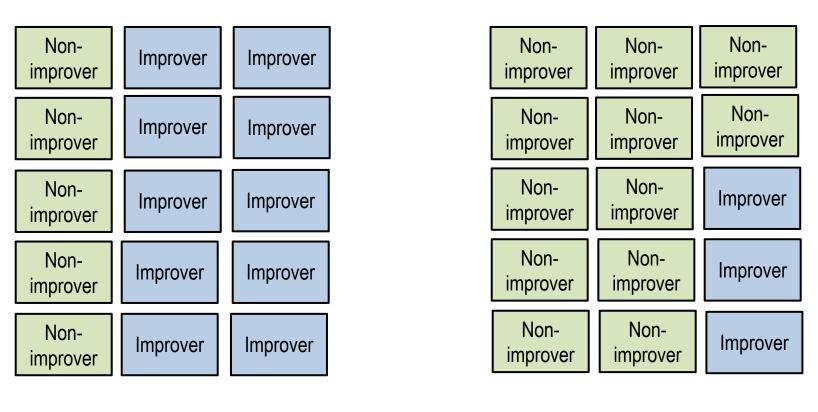
- If the null hypothesis is true (no association between dolphin therapy and improvement) we would have 13 improvers and 17 non-improvers regardless of the group to which they were assigned.
- Hence the assignment doesn't matter and we can just randomly assign the subjects' results to the two groups to see what would happen under a true null hypothesis.

- We can simulate this with cards
 - 13 cards to represent the improvers
 - 17 cards represent the non-improvers
- Shuffle the cards
 - put 15 in one pile (dolphin therapy)
 - put 15 in another (control group)

- Compute the proportion of improvers in the Dolphin Therapy group
- Compute the proportion of improvers in the Control group
- The difference in these two proportions is what could just as well have happened under the assumption there is no association between swimming with dolphins and substantial improvement in depression.

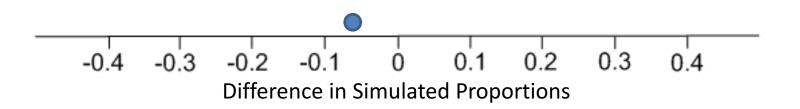
Dolphin Therapy

Control



60.0% Improvers

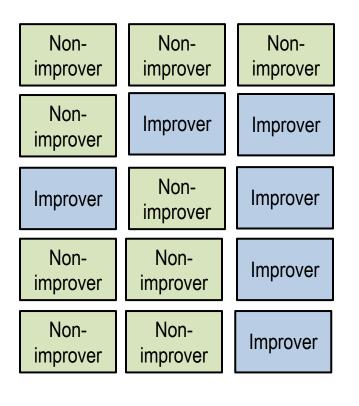
40.0% Improvers

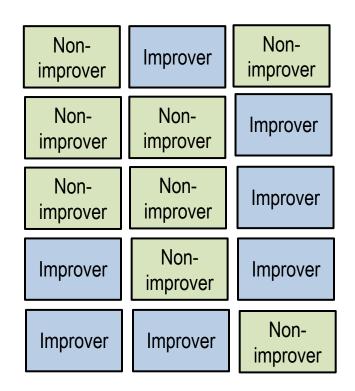


0.400 - 0.467 = -0.067

Dolphin Therapy

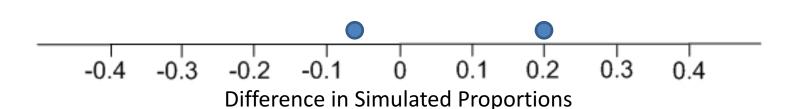
Control





ቆፀ.ፀ% Improvers

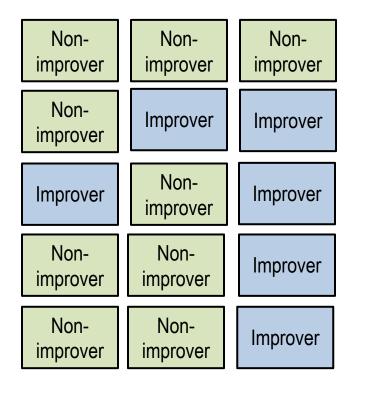
ቆβ.3% Improvers

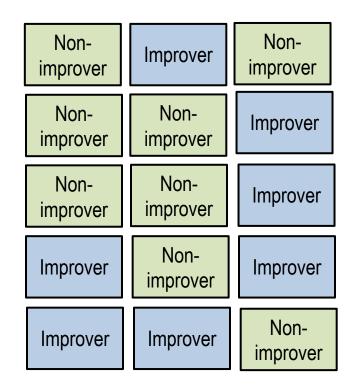


0.533 - 0.333 = 0.200

Dolphin Therapy

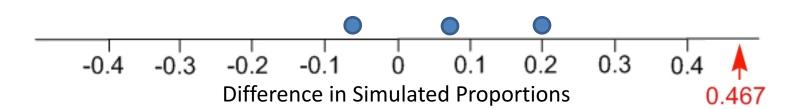
Control





\$B.3% Improvers

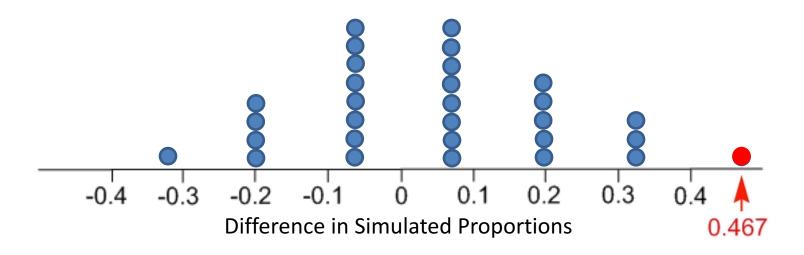
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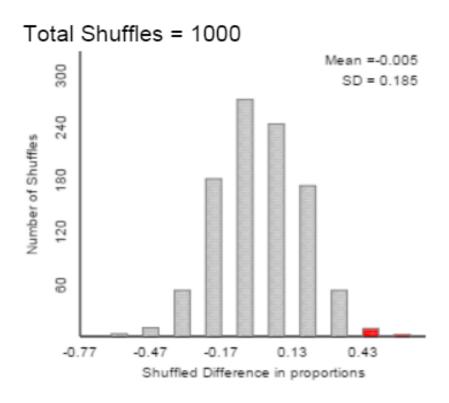
0.467 - 0.400 = 0.067

More Simulations

Only one simulated statistics out of 30 was as large or larger than our observed difference in proportions of 0.467, hence our p-value for this null distribution is $1/30 \approx 0.037$



We did 1000 repetitions to develop a null distribution.



- 13 out of 1000 results had a difference of 0.467 or higher (p-value = 0.013).
- 0.467 is $\frac{0.467-0}{0.185} \approx 2.52$ SD above zero.
 - Using either the p-value or standardized statistic, we have strong evidence against the null and can conclude that the improvement due to swimming with dolphins was statistically significant.

- A 95% confidence interval for the difference in the probability using the standard deviation from the null distribution is $0.467 \pm 2(0.185) = 0.467 \pm 0.370$ or (0.097 to 0.837)
- We are 95% confident that when allowed to swim with dolphins, the probability of improving is between 0.097 and 0.837 higher than when no dolphins are present.
- How does this interval back up our conclusion from the test of significance?
- See if 0 is in the interval.

- Can we say that the presence of dolphins caused this improvement?
 - Since this was a randomized experiment, and assuming everything was identical between the groups, we have strong evidence that dolphins were the cause
- Can we generalize to a larger population?
 - Maybe mild to moderately depressed 18-65 year old patients willing to volunteer for this study
 - We have no evidence that random selection was used to find the 30 subjects. "Outpatients, recruited through announcements on the internet, radio, newspapers, and hospitals."

5. Comparing Two Proportions: Theory-Based Approach, and smoking and gender example.

Section 5.3

Introduction

- Just as with a single proportion, we can often predict results of a simulation using a theory-based approach.
- The theory-based approach also gives a simpler way to generate confidence intervals.
- The main new mathematical fact to use is the SE for the difference between two proportions is

$$\sqrt{\left(\frac{\hat{p}(1-\hat{p})}{n_1} + \frac{\hat{p}(1-\hat{p})}{n_2}\right)}$$
 for testing

or
$$\sqrt{\left(\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}\right)}$$
 for Cls.

A clarification on the new formula

The margin of error for the difference in proportions is

Multiplier × SE, where SE =
$$\sqrt{\left(\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}\right)}$$
.

In testing, the null hypothesis is no difference between the two groups, so we use the SE

$$\sqrt{\left(\frac{\hat{p}(1-\hat{p})}{n_1} + \frac{\hat{p}(1-\hat{p})}{n_2}\right)}$$

where \hat{p} is the proportion in both groups combined. But in

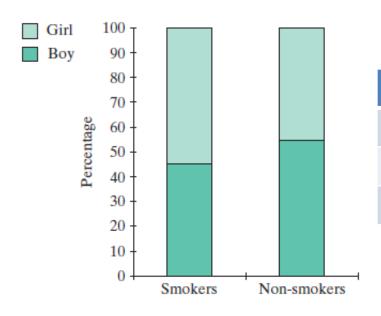
Cls, we use the formula $\sqrt{\left(\frac{\hat{p}_1(1-\hat{p}_1)}{n_1}+\frac{\hat{p}_2(1-\hat{p}_2)}{n_2}\right)}$ because we are not assuming $\hat{p}_1=\hat{p}_2$ with CIs.

Parents' Smoking Status and their Babies' Gender

Example 5.3

- How does parents' behavior affect the gender of their children?
- Fukuda et al. (2002) found the following in Japan.
 - Out of 565 births where both parents smoked more than a pack a day, 255 were boys. This is 45.1% boys.
 - Out of 3602 births where both parents did not smoke,
 1975 were boys. This 54.8% boys.
 - In total, out of 4170 births, 2230 were boys, which is 53.5%.
- Other studies have shown a reduced male to female birth ratio where high concentrations of other environmental chemicals are present (e.g. industrial pollution, pesticides)

- A segmented bar graph and 2-way table
- Let's compare the proportions to see if the difference is statistically significantly.



	Both Smoked	Neither Smoked
Boy	255 (45.1%)	1,975 (54.8%)
Girl	310	1,627
Total	565	3,602

Null Hypothesis:

- There is no association between smoking status of parents and sex of child.
- The probability of having a boy is the same for parents who smoke and don't smoke.
- π_{smoking} $\pi_{\text{nonsmoking}}$ = 0

Alternative Hypothesis:

- There is an association between smoking status of parents and sex of child.
- The probability of having a boy is not the same for parents who smoke and don't smoke
- π_{smoking} $\pi_{\text{nonsmoking}} \neq 0$

- What are the observational units in the study?
- What are the variables in this study?
- Which variable should be considered the explanatory variable and which the response variable?
- What is the parameter of interest?
- Can you draw cause-and-effect conclusions for this study?

Using the 3S Strategy to asses the strength

1. Statistic:

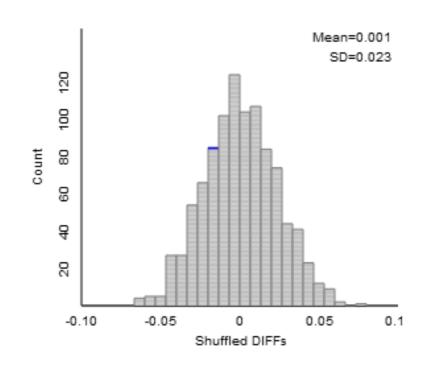
• The proportion of boys born to nonsmokers minus the proportion of boys born to smokers is 0.548 - 0.451 = 0.097.

2. Simulate:

- Use the Two Proportions applet to simulate
- Many repetitions of shuffling the 2230 boys and 1937 girls to the 565 smoking and 3602 nonsmoking parents
- Calculate the difference in proportions of boys between the groups for each repetition.
- Shuffling simulates the null hypothesis of no association

3. Strength of evidence:

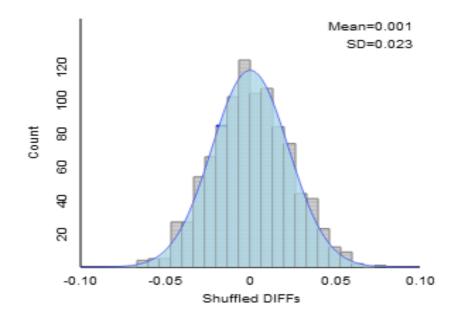
- Nothing as extreme as our observed statistic (≥ 0.097 or ≤ -0.097) occurred in 5000 repetitions,
- How many SDs is 0.097
 above the mean?
 Z = 0.097/0.023 = 4.22
 using simulations. What about using the theory-based approach?



Count Samples Beyond .097

Count = 0/1000 (0.0000)

- Notice the null distribution is centered at zero and is bell-shaped.
- This can be approximated by the normal distribution.



Formulas

• The theory-based approach yields z = 4.30.

$$z = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

• Here
$$z = \frac{.548 - .451}{\sqrt{.535 (1 - .535) \left(\frac{1}{3602} + \frac{1}{565}\right)}} = 4.30.$$

• p-value is 2*(1-pnorm(4.30)) = 0.00171%.

- Fukuda et al. (2002) found the following in Japan.
 - Out of 3602 births where both parents did not smoke,
 1975 were boys. This 54.8% boys.
 - Out of 565 births where both parents smoked more than a pack a day, 255 were boys. This is 45.1% boys.
 - In total, out of 4170 births, 2230 were boys, which is 53.5% boys.

Formulas

 How do we find the margin of error for the difference in proportions?

Multiplier
$$\times \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

- Here n is large and obs. are iid, so we use the multipliers of
 - 1.645 for 90% confidence
 - 1.96 for 95% confidence
 - 2.576 for 99% confidence
- We can write the confidence interval in the form:
 - statistic ± margin of error.

- Our statistic is the observed sample difference in proportions, 0.097.
 - Plugging in 1.96 × $\sqrt{\left(\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}\right)} = 0.044$, we get 0.097 ± 0.044 as our 95% CI.
 - We could also write this interval as (0.053, 0.141).
- We are 95% confident that the probability of a boy baby where neither family smokes minus the probability of a boy baby where both parents smoke is between 0.053 and 0.141.