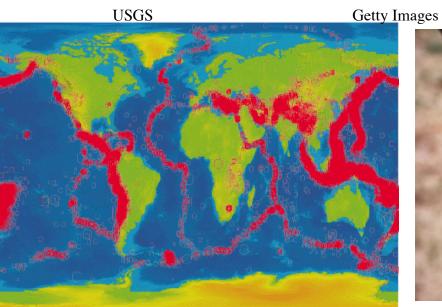
# **Statistics 222, Spatial Statistics.**

## Outline for the day:

- 0. Examples of nonparametric estimation of Hawkes models.
- 1. Deviance residuals.
- 2. Voronoi deviance residuals.
- 3. Superthinning.
- 4. Exercises.
- No class or OH Thu May6.
- Next class we will discuss fithawkes.r, code for fitting space-time Hawkes models.

# Applications to earthquakes and US plague.



## Application to Loma Prieta earthquake data.

Loma Prieta earthquake was Mw 6.9 on Oct 17, 1989.

As an illustration, we will estimate g on its 5566 aftershocks  $M \ge 3$  within 15 months.



(Google images)

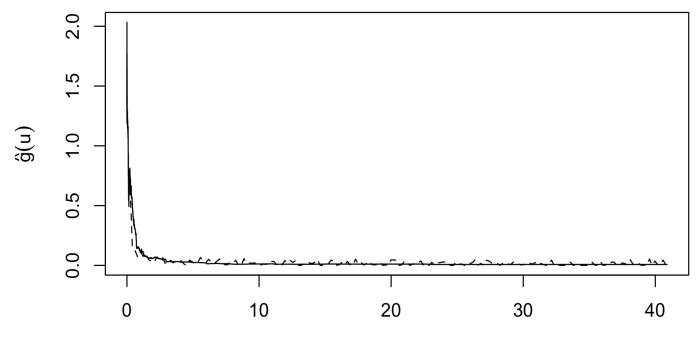
## Application to Loma Prieta earthquake data.

(SCEC.ORG)

Estimated triggering function for 5567 Loma Prieta M  $\geq$  3 events, 10/16/1989 to 1/17/1990.

Solid curve is the analytic method and dashed curve is Marsan and Lengliné (2008).

- Dotted curves are estimates based on analytic method +/- 1 or 2 SEs, respectively, for light grey and dark grey.
- SEs were computed using the SD of analytic estimates in 100 simulations of Hawkes processes with triggering functions sampled from the solid curve.



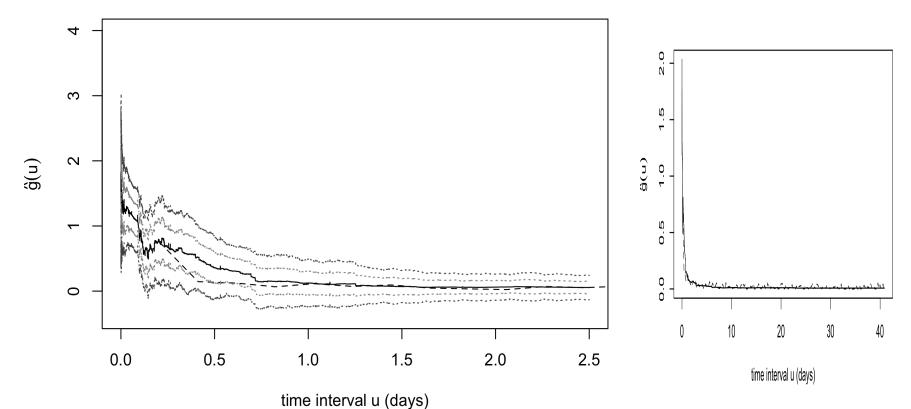
time interval u (days)

## Application to Loma Prieta earthquake data.

Estimated triggering function for Loma Prieta seismicity  $M \ge 3$ , 10/16/1989 to 1/17/1990.

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#### Application to US plague data.

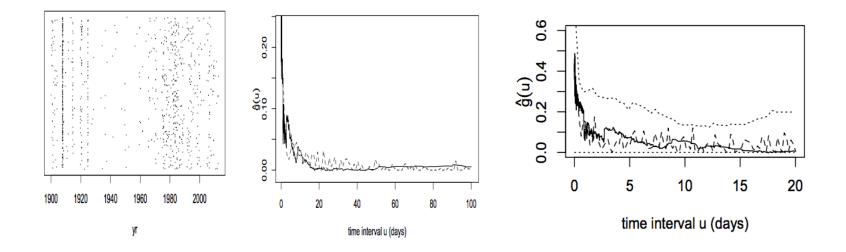


Figure 4: (a) Onset dates of reported and confirmed occurrences of plague in the United States from 1900-2012, according to data from the CDC. The y-coordinates are scattered uniformly at random on the y-axis for ease of visualization. (b) Estimated triggering function,  $\hat{g}$ , for the reported onset times of U.S. plague cases. (c) Estimated triggering function  $\hat{g}$ , for U.S. plague data, for intervals up to 20 days. In (b) and (c), the solid curves correspond to equation (9), the dashed curves result from the method of Marsan and Lengliné (2008), and the dotted curves are the middle 95% range for  $\hat{g}$  from equation (9) resulting from simulating Hawkes models where the true triggering function is that estimated from the data using equation (9).

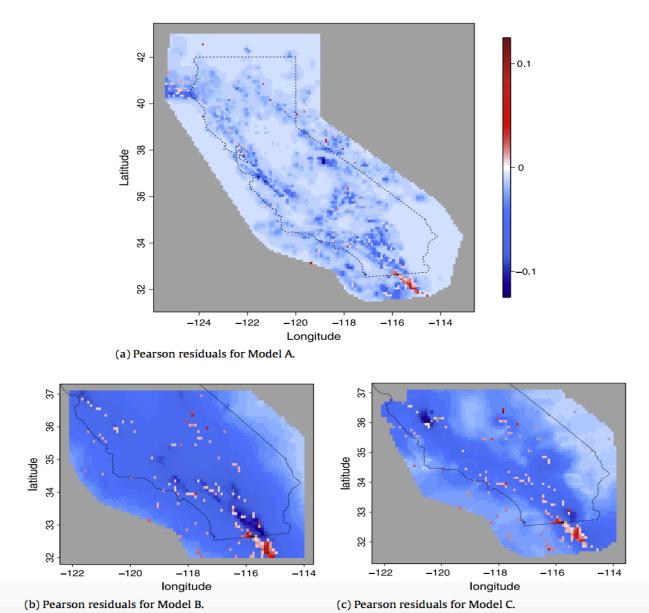
## Recent methods for point process models for occurrences.

## 1. Deviance residuals

- 2. Voronoi residuals
- 3. Superthinned residuals.

Given two competing models, can consider the difference between residuals, number of observed points – number expected, over each pixel. Divide by the estimated SE to get *Pearson residuals* (Baddeley et al. 2005).
Problem: Hard to interpret. If difference = 3, is this because model A overestimated by 3? Or because model B underestimated by 3? Or because model A overestimated by 1 and model B underestimated by 2?
Also, the results are rarely visually appealing or useful.

Pearson residuals tend to look just like a map of the points, unless pixels are very large.



With two competing models, it is better to consider the difference between *log-likelihoods*, in each pixel. The result may be called *deviance residuals* (Clements et al. 2011), ~ resids from gen. linear models.

$$R_{\mathrm{D}}(B_i) = \sum_{i:(t_i, x_i, y_i) \in B_i} \log(\hat{\lambda}_1(t_i, x_i, y_i)) - \int_{B_i} \hat{\lambda}_1(t, x, y) \, \mathrm{d}t \, \mathrm{d}x \, \mathrm{d}y$$
$$- \left(\sum_{i:(t_i, x_i, y_i) \in B_i} \log(\hat{\lambda}_2(t_i, x_i, y_i)) - \int_{B_i} \hat{\lambda}_2(t, x, y) \, \mathrm{d}t \, \mathrm{d}x \, \mathrm{d}y\right)$$

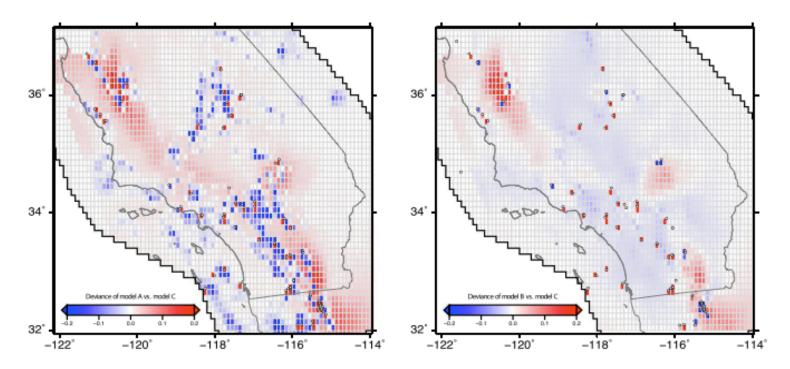


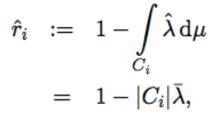
FIG. 4. Left panel (a): deviance residuals for model A versus C. Sum of deviance residuals is 86.427. Right panel (b): deviance residuals for model B versus C. Sum of deviance residuals is -7.468.

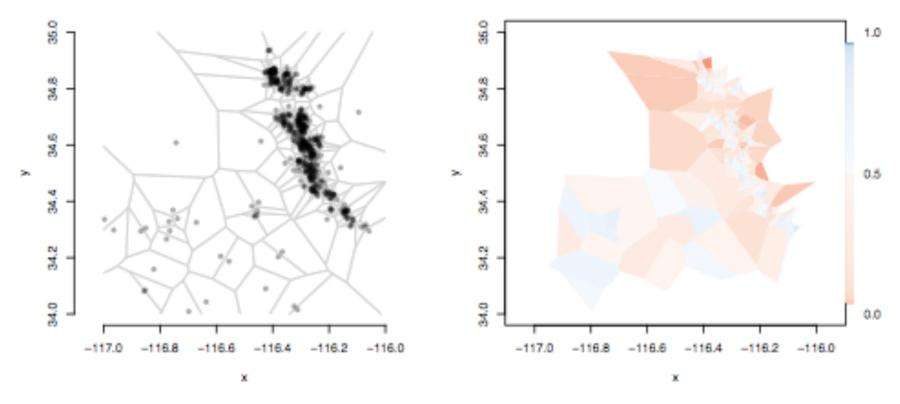
## **<u>2. Voronoi residuals</u>** (Bray et al. 2013)

A Voronoi tessellation divides a space into cells  $C_i$ , where  $C_i$  contains all locations closer to event i than any other observed event.

Within each cell, calculate residuals

 $r \sim 1 - X$ ;  $X \sim \Gamma(3.569, 3.569)$  (Tanemura 2003)





spatially adaptive and nonparametric.

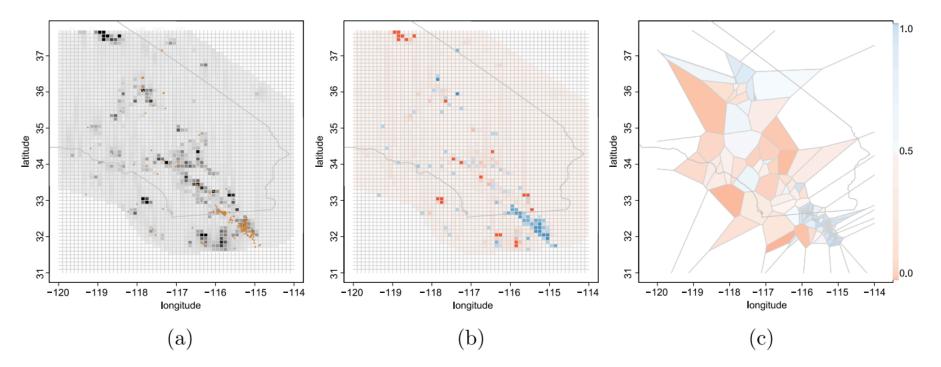
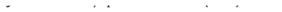
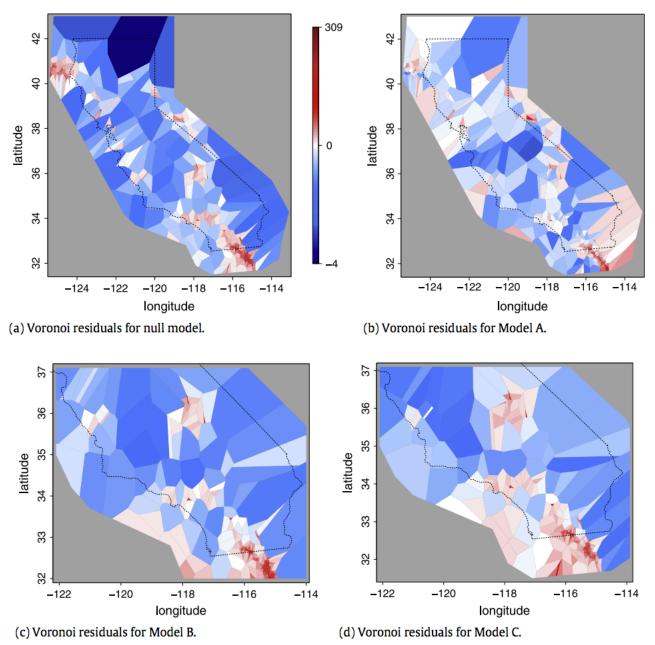


FIG. 2. (a) Estimated rates under the Helmstetter, Kagan and Jackson (2007) model, with epicentral locations of observed earthquakes with  $M \ge 4.0$  in Southern California between January 1, 2006 and January 1, 2011 overlaid. (b) Raw pixel residuals for Helmstetter, Kagan and Jackson (2007) with pixels colored according to their corresponding p-values. (c) Voronol residuals for Helmstetter, Kagan and Jackson (2007) with pixels colored according to their corresponding p-values.





### With 2 models, can compare loglikelihoods across pixels or Voronoi cells.

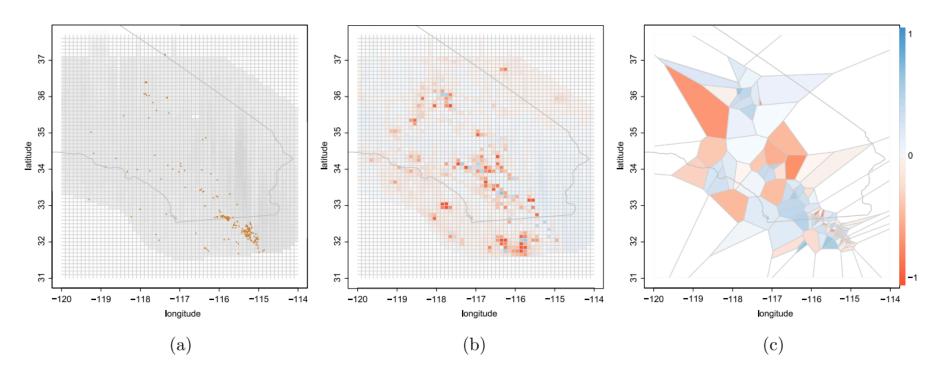


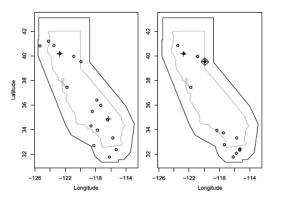
FIG. 3. (a) Estimated rates under the Shen, Jackson and Kagan (2007) model, with epicentral locations of observed earthquakes with  $M \ge 4.0$  in Southern California between January 1, 2006 and January 1, 2011 overlaid. (b) Pixel deviance plot with blue favoring model A, Helmstetter, Kagan and Jackson (2007), versus model B, Shen, Jackson and Kagan (2007). Coloration is on a linear scale. (c) Voronoi deviance plot with blue favoring model A, Helmstetter, Kagan and Jackson and Kagan (2007), versus model B, Shen, Jackson and Kagan (2007). Coloration is on a linear scale.

**<u>3. Superthinning.</u>** (Clements et al., 2012)

Choose some number  $c \sim mean(\hat{\lambda})$ .

Superpose: where  $\hat{\lambda}(t,x,y) < c$ , add in points of a simulated Poisson process of rate  $c - \hat{\lambda}(t,x,y)$ .

Thin: where  $\hat{\lambda}(t_i, x_i, y_i) > c$ , keep each point  $(t_i, x_i, y_i)$  with prob.  $c / \hat{\lambda}(t_i, x_i, y_i)$ .



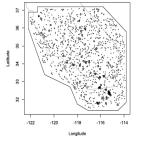


FIG. 9. Superposed residuals for model C. Simulated points make up 90.7% of all points.

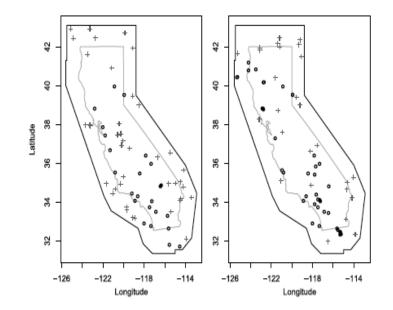


FIG. 11. One realization of super-thinned residuals for the five models considered (circles = observed earthquakes; plus signs = simulated points). Top-left panel (a): model A (k = 2.76). Top-center panel (b): model B (k = 2.95). Top-right panel (c): model C (k = 2.73). Bottom-left panel (d): ETAS (k = 1.35). Bottom-right panel (e): STEP (k = 0.75).

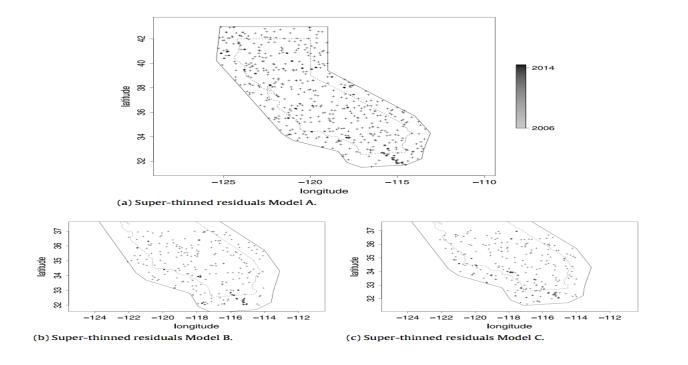
## 3. Superthinning. (Clements et al., 2013)

Choose some number c. mean(  $\hat{\lambda}$  ) at the points is a suggested default.

Superpose: where  $\hat{\lambda}(t,x,y) < c$ , add in points of a simulated Poisson process of rate  $c - \hat{\lambda}(t, x, y)$ .

Thin: where  $\hat{\lambda}(t_i, x_i, y_i) > c$ , keep each point  $(t_i, x_i, y_i)$  with prob.  $c / \hat{\lambda}(t_i, x_i, y_i)$ .

Result is Poisson with rate c, if the model for  $\lambda$  is correct.



Exercises. Superposition.

Suppose  $N_1$  is a Poisson process with rate 3, and  $N_2$  is a Poisson process with rate 2 + x + 4t, independent of  $N_1$ , and both are on  $[0,10] \times [0,1] \times [0,1]$ . t x y.

Let  $M = N_1 + N_2$ . Is M a Poisson process? What is its intensity?

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For any disjoint measurable sets  $B_1, B_2, ..., M(B_i) = N_1(B_i) + N_2(B_i)$  is independent of  $\{N_1(B_j), j \neq i\}$  and  $\{N_2(B_j), j \neq i\}$  and thus is independent of  $\{N_1(B_j) + N_2(B_j), j \neq i\}$ .

So yes, M is a Poisson process and since  $EM(B) = EN_1(B) + EN_2(B)$ , M has rate 5 + x + 4t.

Exercises.

Suppose N is homogeneous Poisson process with rate 1, and M is a clustered Hawkes process.

Both M and N have 40 points on  $B = [0,10] \times [0,1] \times [0,1]$ 

Let v1 = the average size of a Voronoi cell in a Voronoi tessellation of N, and v2 = the average size of a Voronoi cell in a Voronoi tessellation of M. Which is bigger, v1 or v2, or will they be the same?

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The same, since  $v1 = v2 = \frac{1}{4}$ . Each cell has one point, and the 40 cells occupy an area of size 10.

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