

# Analysis of Experiments With Functional Responses

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- Joint work with Dr. Qing Shen (Edmunds.com Inc.)

# Outline

- Introduction
- Linear models with functional responses
- Diagnostics
- An Example
- Summary

## Introduction

- Many designed experiments have functional responses.
  - Robust parameter designs with a signal factor.
- Functional linear models are useful for such experiments
  - functional response, scalar predictors
- Ramsay and Silverman (1997, 2005):  
*Introduction to Functional Data Analysis*
- Ramsay and Silverman (2002):  
*Applied Functional Data Analysis: Methods and Case Studies*
- Nair, Taam and Ye (2002, JQT) analyzed three real robust design experiments with functional responses.

## Example: Alternator Experiment

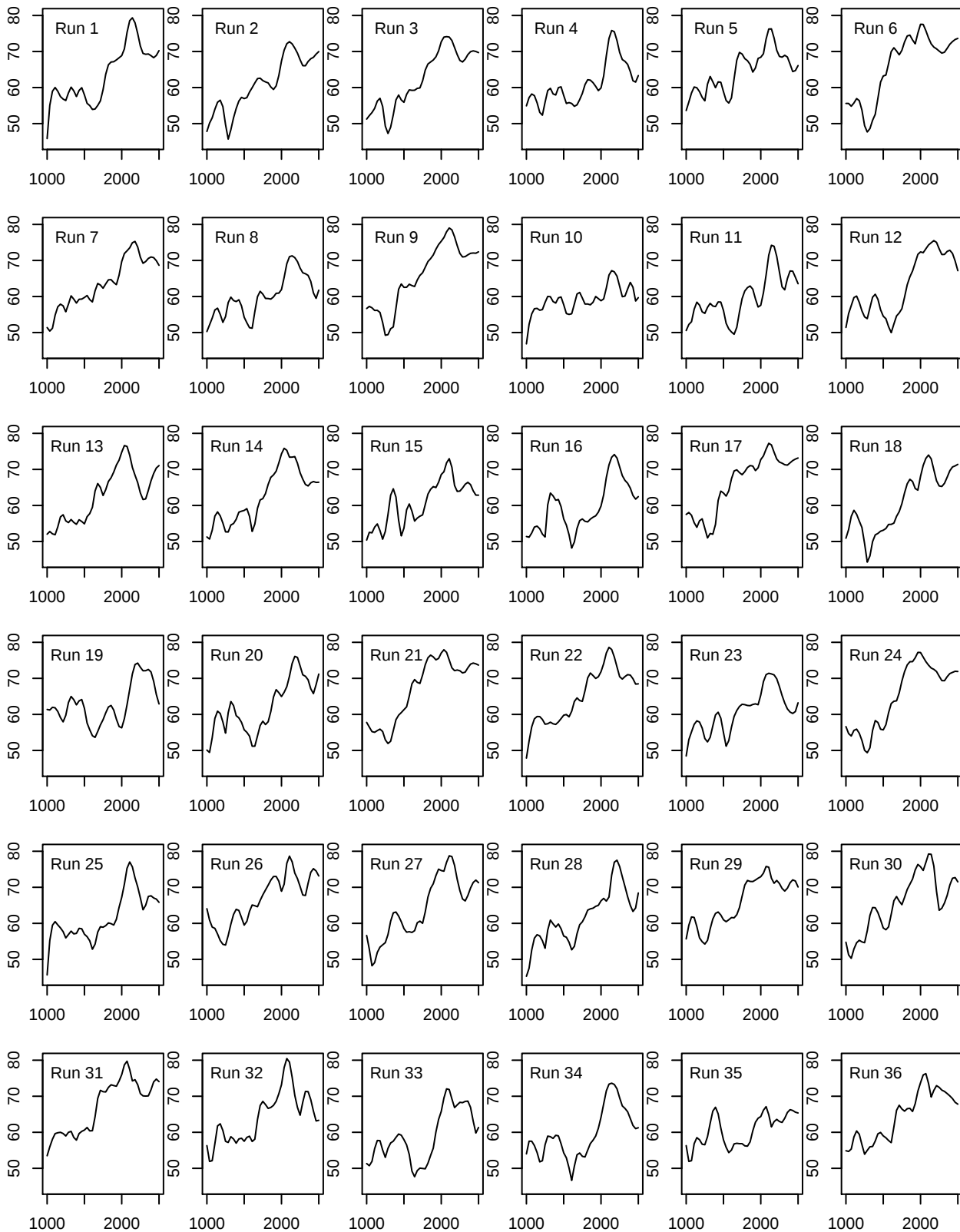
- A robust design experiment from Nair, Taam and Ye (2002)
- Study the effects of 7 process assembly parameters (factors  $A-G$ ) on the audible noise levels of alternators.
- Noise factor:  $D$ .
- A  $2_{IV}^{7-2}$  design plus 4 additional runs at the high levels
- signal factor: rotating speed
- recorded 43 measurements of sound pressure levels (responses) at rotating speeds ranging from 1000 to 2500 revolution per minute.

(Fig. Observed response curves)

| Factor   | Description             |
|----------|-------------------------|
| <i>A</i> | Through Bolt Torque     |
| <i>B</i> | Rotor Balance           |
| <i>C</i> | Stator Varnish          |
| <i>D</i> | Air Gap Variation       |
| <i>E</i> | Stator Orientation      |
| <i>F</i> | Housing Stator Slip Fit |
| <i>G</i> | Shaft Radial Alignment  |

| Run | <i>A</i> | <i>B</i> | <i>C</i> | <i>D</i> | <i>E</i> | <i>F</i> | <i>G</i> |
|-----|----------|----------|----------|----------|----------|----------|----------|
| 1   | -1       | -1       | -1       | -1       | -1       | 1        | 1        |
| 2   | -1       | -1       | -1       | -1       | 1        | 1        | -1       |
| 3   | -1       | -1       | -1       | 1        | -1       | -1       | -1       |
| 4   | -1       | -1       | -1       | 1        | 1        | -1       | 1        |
| 5   | -1       | -1       | 1        | -1       | -1       | -1       | 1        |
| 6   | -1       | -1       | 1        | -1       | 1        | -1       | -1       |
| 7   | -1       | -1       | 1        | 1        | -1       | 1        | -1       |
| 8   | -1       | -1       | 1        | 1        | 1        | 1        | 1        |
| 9   | -1       | 1        | -1       | -1       | -1       | -1       | -1       |
| 10  | -1       | 1        | -1       | -1       | 1        | -1       | 1        |
| 11  | -1       | 1        | -1       | 1        | -1       | 1        | 1        |
| 12  | -1       | 1        | -1       | 1        | 1        | 1        | -1       |
| 13  | -1       | 1        | 1        | -1       | -1       | 1        | -1       |
| 14  | -1       | 1        | 1        | -1       | 1        | 1        | 1        |
| 15  | -1       | 1        | 1        | 1        | -1       | -1       | 1        |
| 16  | -1       | 1        | 1        | 1        | 1        | -1       | -1       |
| 17  | 1        | -1       | -1       | -1       | -1       | -1       | -1       |
| 18  | 1        | -1       | -1       | -1       | 1        | -1       | 1        |
| 19  | 1        | -1       | -1       | 1        | -1       | 1        | 1        |
| 20  | 1        | -1       | -1       | 1        | 1        | 1        | -1       |
| 21  | 1        | -1       | 1        | -1       | -1       | 1        | -1       |
| 22  | 1        | -1       | 1        | -1       | 1        | 1        | 1        |
| 23  | 1        | -1       | 1        | 1        | -1       | -1       | 1        |
| 24  | 1        | -1       | 1        | 1        | 1        | -1       | -1       |
| 25  | 1        | 1        | -1       | -1       | -1       | 1        | 1        |
| 26  | 1        | 1        | -1       | -1       | 1        | 1        | -1       |
| 27  | 1        | 1        | -1       | 1        | -1       | -1       | -1       |
| 28  | 1        | 1        | -1       | 1        | 1        | -1       | 1        |
| 29  | 1        | 1        | 1        | -1       | -1       | -1       | 1        |
| 30  | 1        | 1        | 1        | -1       | 1        | -1       | -1       |
| 31  | 1        | 1        | 1        | 1        | -1       | 1        | -1       |
| 32  | 1        | 1        | 1        | 1        | 1        | 1        | 1        |
| 33  | 1        | 1        | 1        | 1        | 1        | 1        | 1        |
| 34  | 1        | 1        | 1        | 1        | 1        | 1        | 1        |
| 35  | 1        | 1        | 1        | 1        | 1        | 1        | 1        |
| 36  | 1        | 1        | 1        | 1        | 1        | 1        | 1        |

# Observed Response Curves



# Linear Models With Functional Responses

Model:

$$\mathbf{Y}(t) = \mathbf{X}\boldsymbol{\beta}(t) + \boldsymbol{\epsilon}(t), \quad t \in [a, b]$$

- $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_n)^T$  is an  $n \times p$  matrix
- $\boldsymbol{\beta}(t) = (\beta_1(t), \dots, \beta_p(t))^T$
- $\boldsymbol{\epsilon}(t) = (\epsilon_1(t), \dots, \epsilon_n(t))^T$
- each  $\epsilon_i(t)$  is an independent realization of a Gaussian process with mean 0 and covariance function  $\gamma(s, t)$ .
- eigen-decomposition

$$\gamma(s, t) = \sum_{i=1}^{\infty} \lambda_i \phi_i(s) \phi_i(t),$$

where  $\lambda_i$  are eigenvalues and  $\phi_i$  are orthonormal eigenfunctions.

## Estimation

- Each curve is a point in a functional space.
- The least squares estimates minimize the residual SS in  $L_2$ .

$$\hat{\boldsymbol{\beta}}(t) = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}(t)$$

- This is point-wise l.s. estimation.
- Fitted responses:  $\hat{y}_i(t) = \mathbf{x}_i^T \hat{\boldsymbol{\beta}}(t)$
- Residuals:  $\hat{\epsilon}_i(t) = y_i(t) - \hat{y}_i(t)$ .
- Residual SS:  $r_{SS} = \sum_{i=1}^n \|\hat{\epsilon}_i\|^2 = \sum_{i=1}^n \int \hat{\epsilon}_i(t)^2 dt$ .

## Testing Nested Linear Hypothesis

$$H_0(\omega) : \quad \mathbf{Y}(t) = \mathbf{X}_1\boldsymbol{\beta}_1(t) + \boldsymbol{\epsilon}(t)$$

$$H_1(\Omega) : \quad \mathbf{Y}(t) = \mathbf{X}_1\boldsymbol{\beta}_1(t) + \mathbf{X}_2\boldsymbol{\beta}_2(t) + \boldsymbol{\epsilon}(t)$$

- Traditional multivariate tests are not appropriate.
- Ramsay and Silverman (1997) used pointwise  $t$  statistics
  - problem of multiple comparison and correlated responses
- Faraway (1997) proposed a bootstrap based test
  - heavy computation
- Fan and Lin (1998) proposed adaptive transform-based tests
- Shen and Faraway (2004) proposed a functional F test
  - similar to an ordinary F test
  - easy to use

## Functional F Test

$$F = \frac{(rss_{\omega} - rss_{\Omega})/(p - q)}{rss_{\Omega}/(n - p)}$$

- Under the null hypothesis,

$$F \sim \frac{\sum_{i=1}^{\infty} \lambda_i \chi_{p-q}^2 / (p - q)}{\sum_{i=1}^{\infty} \lambda_i \chi_{n-p}^2 / (n - p)},$$

where  $\lambda_i$  are eigenvalues of  $\gamma(s, t)$  and the  $\chi^2$  random variables are independent.

- approximate it by an ordinary F distribution,

$$F \approx F(\lambda(p - q), \lambda(n - p))$$

where

$$\lambda = (\sum \lambda_i)^2 / \sum \lambda_i^2$$

is the *degrees-of-freedom-adjustment-factor*.

## Testing Individual Predictors

$$H_0 : \beta_j(t) = 0 \text{ vs. } H_1 : \beta_j(t) \neq 0$$

- Functional  $F$  test statistic

$$F_j = \frac{rss_j - rss}{rss/(n - p)}$$

where  $rss_j$  is the residual SS under  $H_0$ .

$$F_j = \frac{\|\hat{\beta}_j\|^2}{(rss/(n - p))(\mathbf{X}^T \mathbf{X})_{jj}^{-1}} = \frac{\|\hat{\beta}_j\|^2}{(se)^2}.$$

- approximated by an ordinary F distribution

$$F_j \approx F(\lambda, \lambda(n - p))$$

## In Practice

- only observe  $y_i(t)$  for  $t_1, \dots, t_m$ ,

$$y_i(t_j) = \mathbf{x}_i^T \boldsymbol{\beta}(t_j) + \epsilon(t_j), i = 1, \dots, n, j = 1, \dots, m,$$

- do point-wise estimation and regression.
- compute  $\|\hat{\epsilon}_i\|^2 = \sum_{k=1}^m \hat{\epsilon}_i(t_k)^2 / m$
- compute  $\|\hat{\beta}_j\|^2 = \sum_{k=1}^m \hat{\beta}_j(t_k)^2 / m$ .
- estimate the covariance function  $\gamma(s, t)$  by the empirical covariance matrix  $\hat{\Sigma}$
- estimate the degrees-of-freedom-adjustment-factor by

$$\hat{\lambda} = \text{trace}(\hat{\Sigma})^2 / \text{trace}(\hat{\Sigma}^2).$$

- A large sample size ( $n \geq 30$ ) is desired.

## Diagnostics

- Outliers exist not only in observational studies but also in designed experiments.
- Outliers and influential cases often lead to misleading conclusions.
- Point-wise diagnostics are useful, but
  - residuals are correlated over time.
  - inconsistency often emerges from different time points.
- Truly functional diagnostics is desired

## Studentized Residuals

- Hat matrix:  $\mathbf{H} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$
- Leverage  $h_{ii}$  is the  $i$ th diagonal element of  $\mathbf{H}$ .

Studentized residuals (for scalar regression):

$$s_i = \frac{\hat{\epsilon}_i}{\sqrt{1 - h_{ii}} \hat{\sigma}}$$

Studentized residuals (for functional regression):

$$S_i = \frac{\|\hat{\epsilon}_i\|}{\sqrt{1 - h_{ii}} \hat{\Delta}}$$

where  $\hat{\Delta}^2 = rss/(n - p)$  is an unbiased estimate of total variance  
 $\Delta^2 = \int \gamma(t, t) dt = \sum \lambda_k$ .

- $S_i$  and  $\|\hat{y}_i\|$  are uncorrelated.

## Studentized Residuals

- If the error process is Gaussian,

$$\frac{\|\hat{\epsilon}_i\|^2}{1 - h_{ii}} \sim \sum_{k=1}^{\infty} \lambda_k \chi_1^2,$$

where  $\lambda_k$  are eigenvalues of  $\gamma(s, t)$  and the  $\chi^2$  random variables are independent.

$$\frac{\|\hat{\epsilon}_i\|^2}{1 - h_{ii}} \approx c \chi_{\lambda}^2,$$

- $\lambda = (\sum \lambda_k)^2 / (\sum \lambda_k^2)$  is the adjustment factor.

Residual plots:

- Residual vs. Fitted Plot:  $S_i$  against  $\|\hat{y}_i\|$ .
- Chi-square Q-Q plot:  $S_i^2$  against quantiles of a  $\chi^2$  distribution with  $\text{df} = \hat{\lambda}$ .

## Outliers and Jackknife Residuals

Scalar:

$$j_i = \frac{y_i - \tilde{y}_i}{\hat{\sigma}_{(i)} \sqrt{1 + \mathbf{x}_i^T (\mathbf{X}_{(i)}^T \mathbf{X}_{(i)})^{-1} \mathbf{x}_i}}$$

Functional:

$$J_i = \frac{\|y_i - \tilde{y}_i\|}{\sqrt{1 + \mathbf{x}_i^T (\mathbf{X}_{(i)}^T \mathbf{X}_{(i)})^{-1} \mathbf{x}_i} \hat{\Delta}_{(i)}}$$

where  $\hat{\Delta}_{(i)}^2$  is an unbiased estimate of total variance  $\Delta^2$  without the  $i$ th case.

## Outliers and Jackknife Residuals

To test whether the  $i$ th case is an outlier, consider the *mean shift outlier model*:

$$y_j(t) = \mathbf{x}_j^T \boldsymbol{\beta}(t) + \epsilon_j(t), \quad j \neq i,$$

$$y_i(t) = \mathbf{x}_i^T \boldsymbol{\beta}(t) + \delta(t) + \epsilon_i(t).$$

- The  $i$ th case is not an outlier is equivalent to  $\delta(t) = 0$ .
- A functional F test can be used for the latter.
- $F_i = J_i^2$ .

## Outliers and Jackknife Residuals

- If the  $i$ th case is not an outlier,

$$J_i^2 \sim \frac{\sum_{k=1}^{\infty} \lambda_k \chi_1^2}{\sum_{k=1}^{\infty} \lambda_k \chi_{n-p-1}^2 / (n-p-1)},$$

where  $\lambda_k$  are eigenvalues of  $\gamma(s, t)$  and the  $\chi^2$  random variables are independent.

- Use functional F test to formally detect outliers.

$$J_i^2 \approx F(\lambda, \lambda(n-p-1))$$

- Formula for computation:

$$J_i = S_i \sqrt{\frac{n-p-1}{n-p-S_i^2}}$$

## Influential Cases and Cook's Distance

Scalar:

$$d_i = \frac{(\hat{\beta}_{(i)} - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\hat{\beta}_{(i)} - \hat{\beta})}{p \hat{\sigma}^2}$$

Functional:

$$D_i = \frac{\int (\hat{\beta}_{(i)}(t) - \hat{\beta}(t))^T (\mathbf{X}^T \mathbf{X}) (\hat{\beta}_{(i)}(t) - \hat{\beta}(t)) dt}{p \hat{\Delta}^2}.$$

Formula for computation:

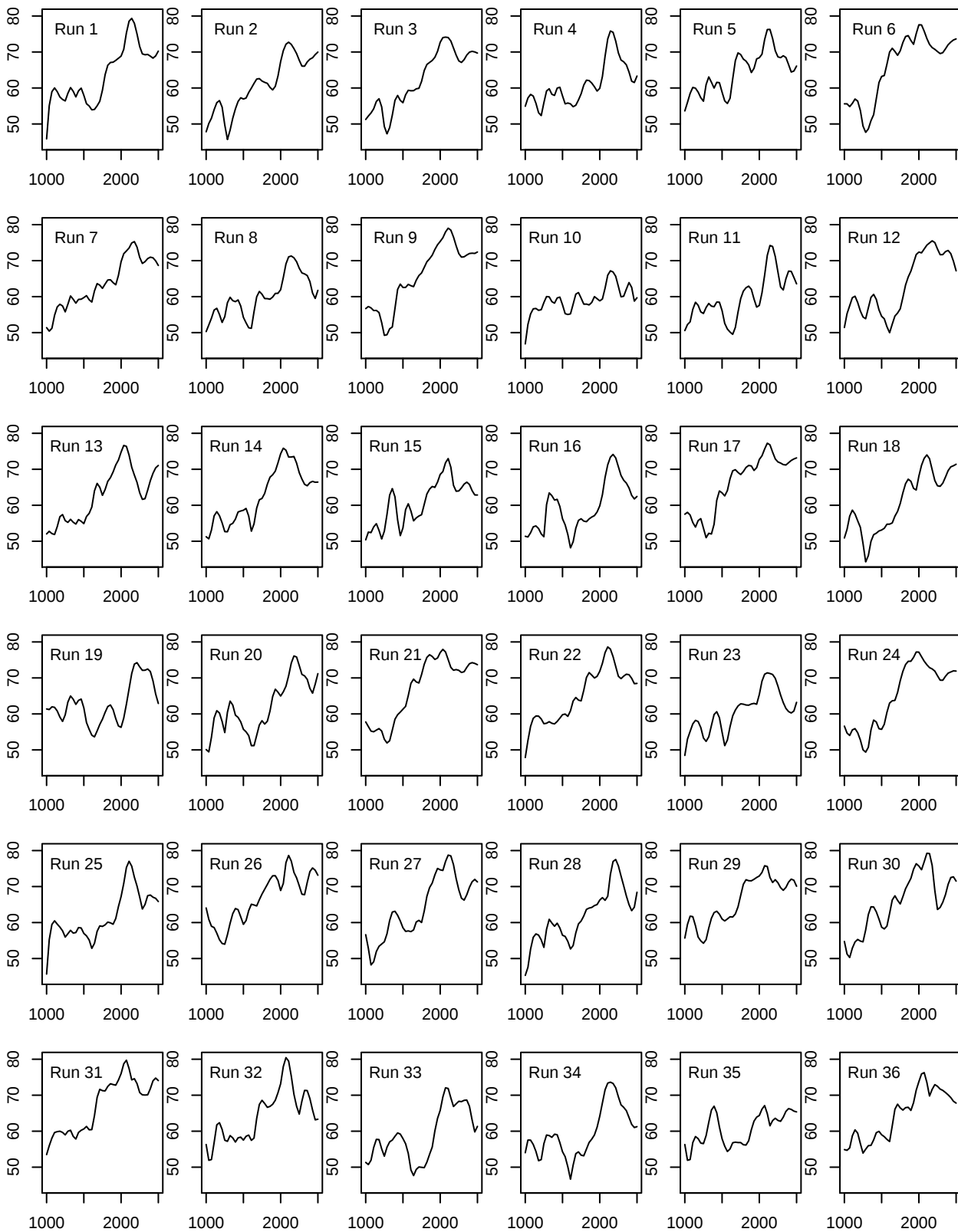
$$D_i = \frac{1}{p} \frac{h_{ii}}{1 - h_{ii}} S_i^2.$$

## Example: Alternator Experiment

- A  $2_{IV}^{7-2}$  design with defining relation  
 $I = C E F G = A B C D F = A B D E G$ .
- It allows the clear estimation of all 7 main effects and 15 two-factor interactions.
- The 6 remaining 2-factor interactions are aliased in three pairs (i.e.,  $C E = F G$ ,  $C F = E G$ ,  $C G = E F$ ).
- For simplicity, ignore the last four runs in the analysis.
- Start with a main effects model.
- Indeed, none of the 2-factor interactions is significant.

(Fig. Observed response curves)

# Observed Response Curves



## Main Effects Model

|           | Numerator | Denominator | F Value | P-Value |
|-----------|-----------|-------------|---------|---------|
| Intercept | 4035.1    | 0.3556      | 11348   | 0.0000  |
| A         | 1.1633    | 0.3556      | 3.2716  | 0.0093  |
| B         | 0.4084    | 0.3556      | 1.1485  | 0.3388  |
| C         | 1.2838    | 0.3556      | 3.6103  | 0.0051  |
| D         | 1.9419    | 0.3556      | 5.4611  | 0.0002  |
| E         | 0.3225    | 0.3556      | 0.9071  | 0.4762  |
| F         | 0.3602    | 0.3556      | 1.0130  | 0.4117  |
| G         | 2.7768    | 0.3556      | 7.8091  | 0.0000  |

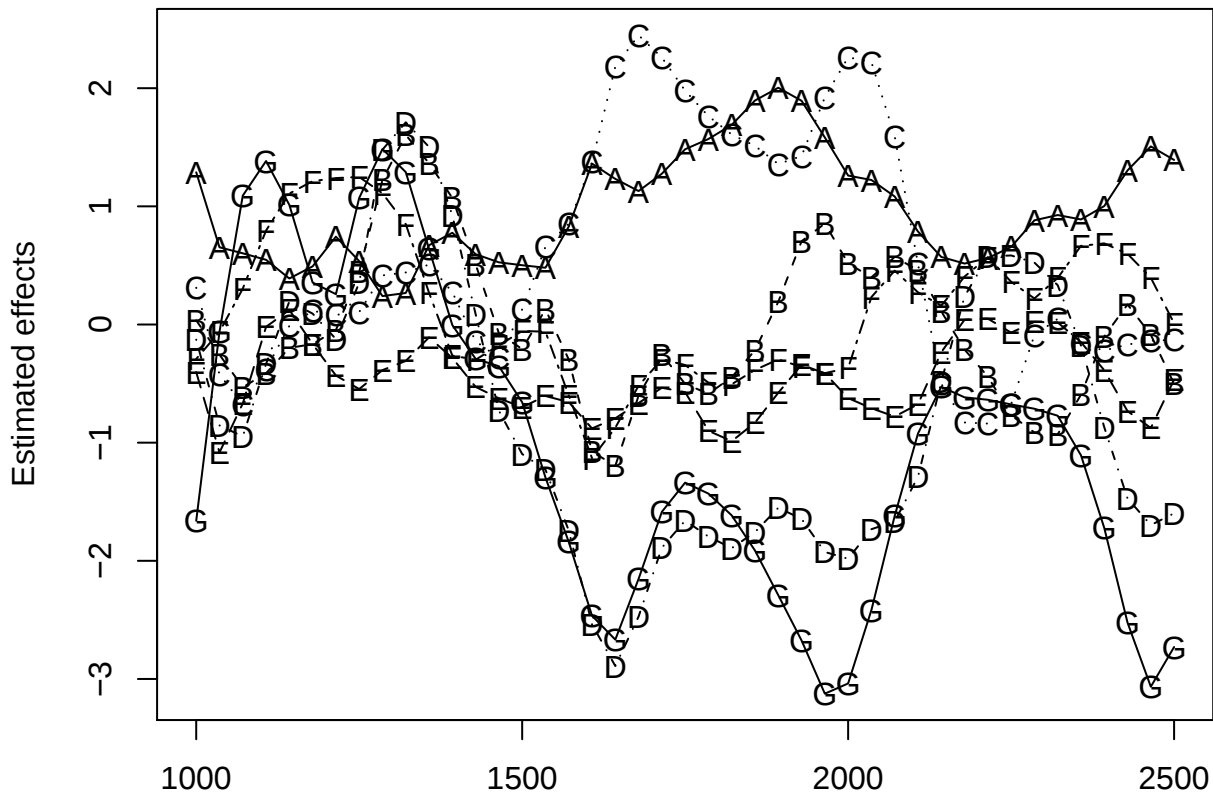
Average residual SS: 273.0895 on 24 degrees of freedom

Adjustment factor: 4.808803 (for 43 grids)

- $D$  and  $G$  are very significant.
- $A$  and  $C$  are significant at 1% margin.

(Fig. Estimated effect curves from the main effects model)

# Estimated Effect Curves (Main Effects Model)



## Reduced Model

|           | Numerator | Denominator | F Value | P-Value |
|-----------|-----------|-------------|---------|---------|
| Intercept | 4035.1    | 0.3565      | 11319   | 0.0000  |
| A         | 1.1633    | 0.3565      | 3.2633  | 0.0076  |
| C         | 1.2838    | 0.3565      | 3.6012  | 0.0040  |
| D         | 1.9419    | 0.3565      | 5.4473  | 0.0001  |
| G         | 2.7768    | 0.3565      | 7.7893  | 0.0000  |

Average residual SS: 308.0066 on 27 degrees of freedom

Adjustment factor: 5.135548 (for 43 grids)

## ANOVA Table

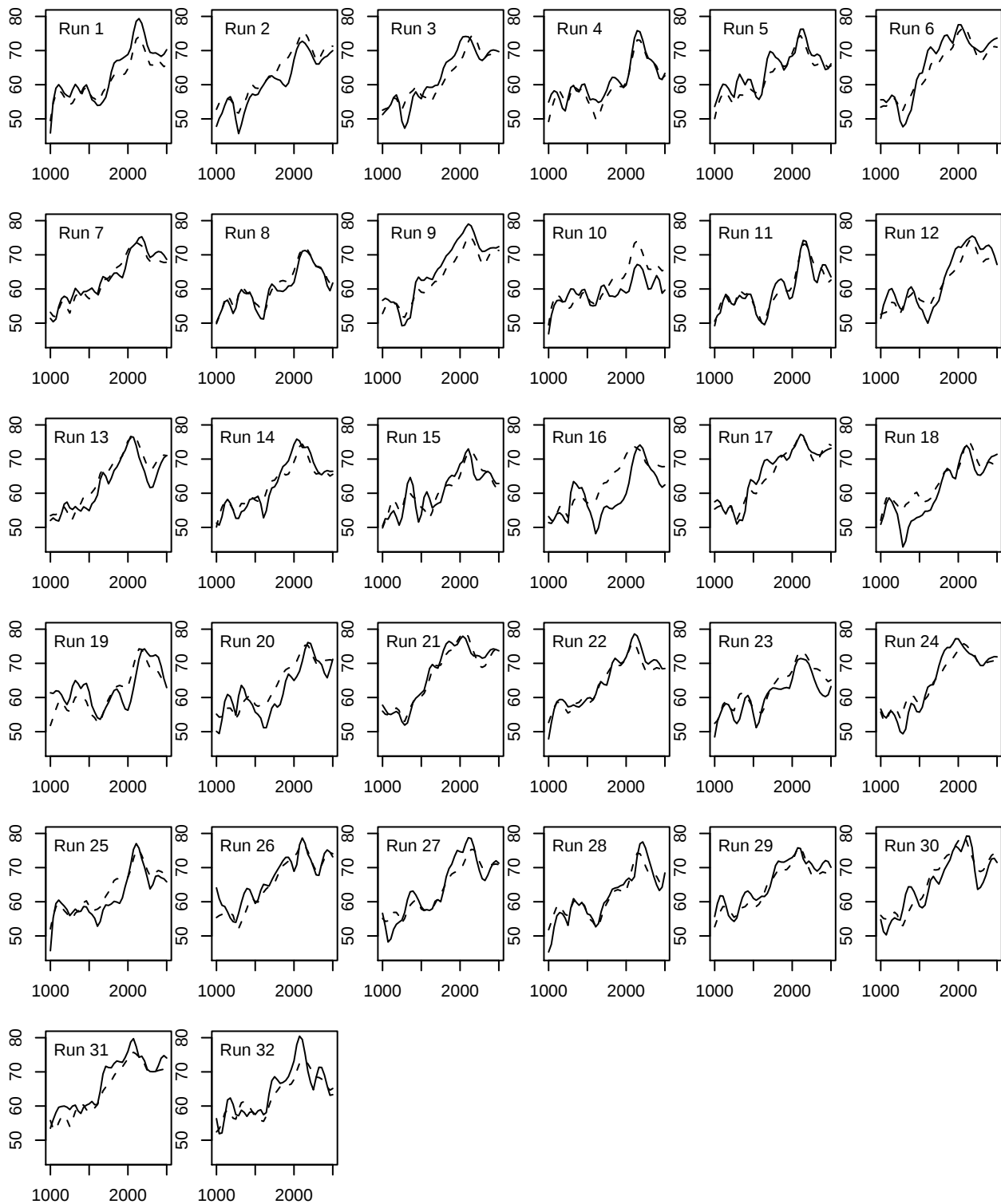
|              | DF | Resid SS | Mean SS | F    |
|--------------|----|----------|---------|------|
| Reduced      | 27 | 308.0066 | 11.4077 | 1.02 |
| Main-Effects | 24 | 273.0895 | 11.3787 |      |

Adjusted DF: 14.4 and 115.4, p-value: 0.4365

Adjustment factor: 4.808803 (for 43 grids)

(Fig. Observed and fitted curves for the reduced model)

# Observed and Fitted Curves (Reduced Model)



Observed (solid), fitted (dashed)

## Diagnosics

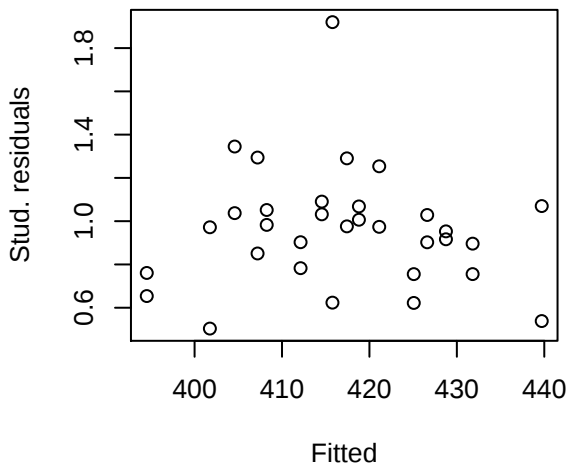
(Fig. Diagnostics plots for the reduced model)

- The formal outlier test declares that case 16 is an outlier.
- The horizontal line on the jackknife plot shows the (Bonferoni adjusted) critical value at the 5% level.

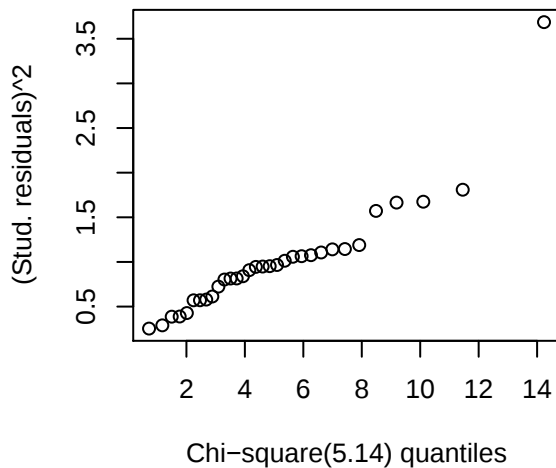
(Fig. Observed and fitted curves for the reduced model)

# Diagnostics Plots (Reduced Model)

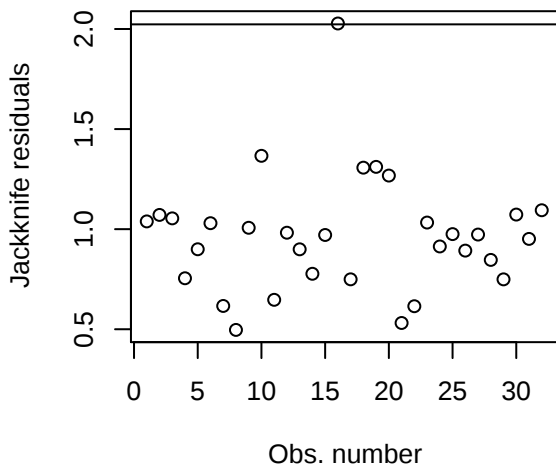
## Residuals vs Fitted



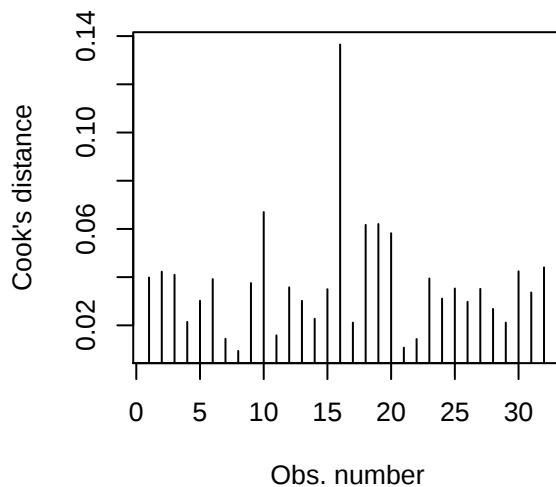
## Chi-square Q-Q plot



## Jackknife residuals plot



## Cook's distance plot



## Reduced Model Without Case 16

|           | Numerator | Denominator | F Value | P-Value |
|-----------|-----------|-------------|---------|---------|
| Intercept | 4052.1    | 0.3315      | 12224   | 0.0000  |
| A         | 0.8192    | 0.3315      | 2.4712  | 0.0256  |
| C         | 1.7529    | 0.3315      | 5.2879  | 0.0001  |
| D         | 1.4225    | 0.3315      | 4.2911  | 0.0005  |
| G         | 3.4725    | 0.3315      | 10.4753 | 0.0000  |

Average residual SS: 265.9543 on 26 degrees of freedom

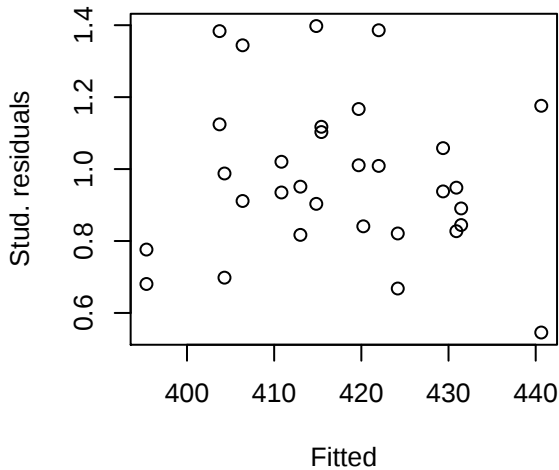
Adjustment factor: 6.055215 (for 43 grids)

- $G$  is still most significant,
- but  $C$  is more significant than  $D$ ,
- $A$  becomes less significant  
(p-value increases from .0076 to .0256).

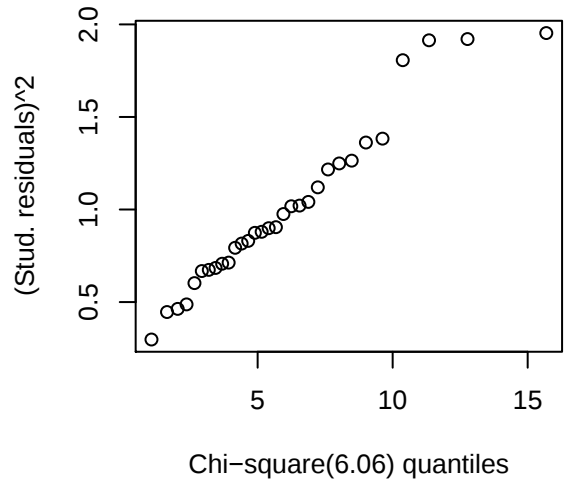
(Fig. Diagnostics plots for the reduced model w/o case 16)

# Diagnostics Plots (reduced model w/o case 16)

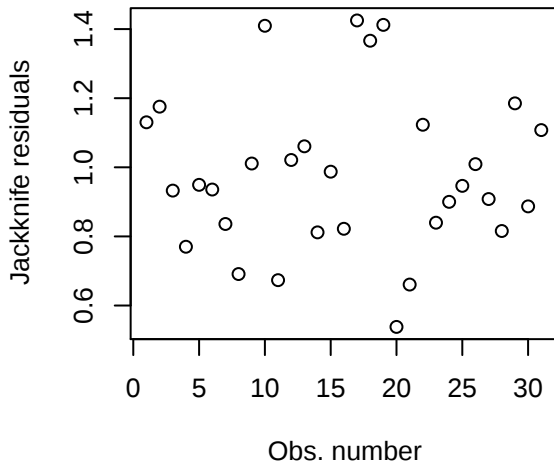
## Residuals vs Fitted



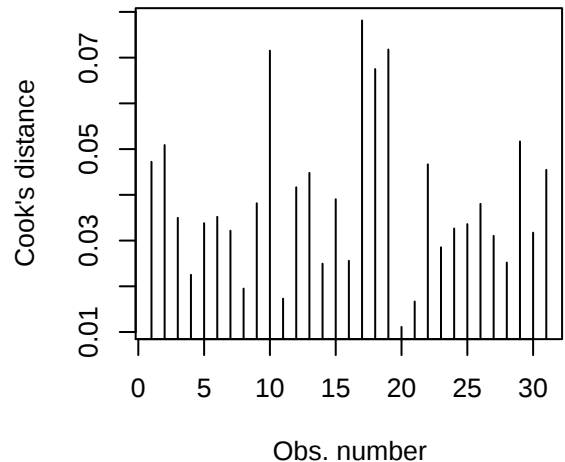
## Chi-square Q-Q plot



## Jackknife residuals plot



## Cook's distance plot



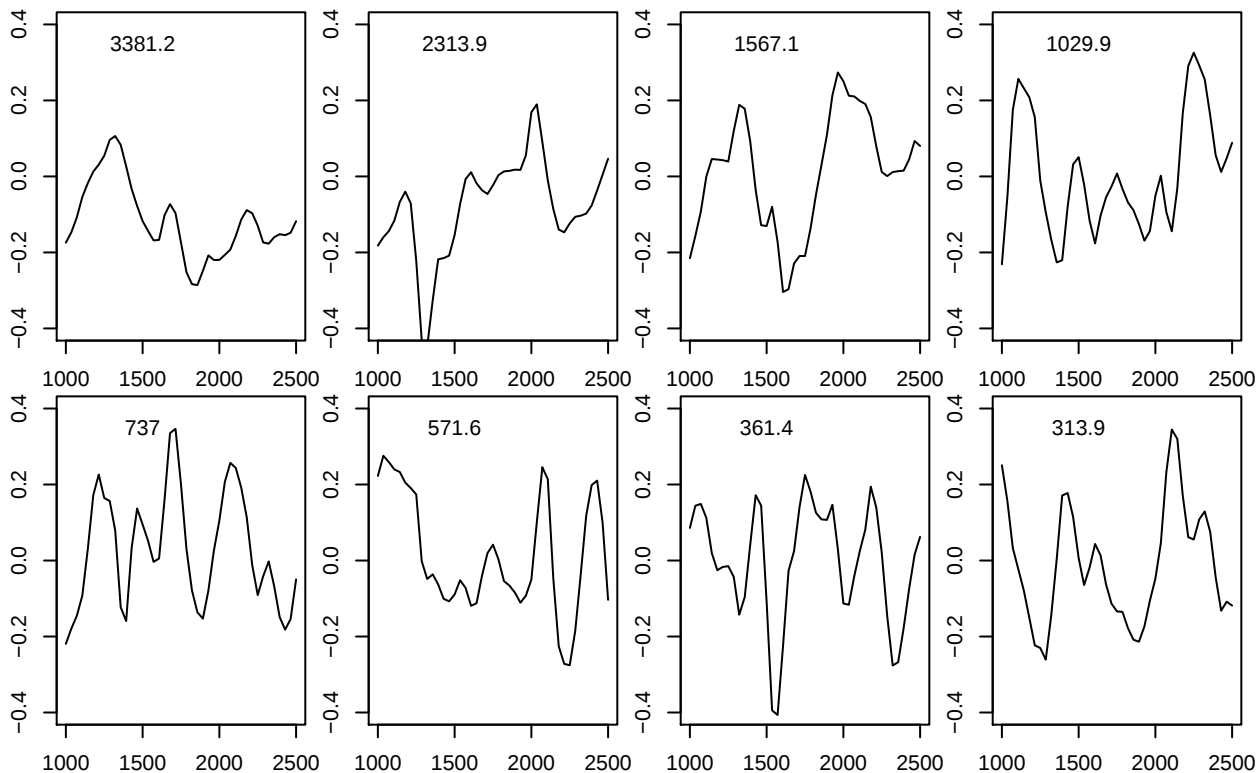
## Discussion

- Nair, Taam and Ye (2002)
  - used point-wise CI to declare significant effects.
  - did not perform diagnostics.

(Fig. Estimated eigenfunctions and their associated eigenvalues)

- The error process is rather complicated.
- The first 8 eigenvalues explain 90% variation of the residual functions.
- Difficult to decide the dimension of the error process.
- The functional F test does not suffer from this difficulty.

# Estimated Eigenfunctions and Associated Eigenvalues (reduced model w/o case 16)



## Summary

Functional linear regression models

- functional responses and scale predictors
- functional F test
  - simple and flexible
  - results are not sensitive to small eigenvalues
- diagnostics
  - studentized residuals (chi-square Q-Q plot)
  - jackknife residuals (formal test for outliers)
  - Cook's distance
  - easy computation formulas

## References

1. Nair, V. N., Taam, W., and Ye, K. Q. (2002), “Analysis of Functional Responses From Robust Design Studies,” *Journal of Quality Technology*, 34, 355–370.
2. Shen, Q. and Faraway, J. (2004). An  $F$  Test for Linear Models With Functional Responses. *Statistica Sinica* **14** 1239–1257.
3. Shen, Q. and Xu, H. (2005). Diagnostics for Linear Models With Functional Responses. UCLA Department of Statistics, preprint 439.