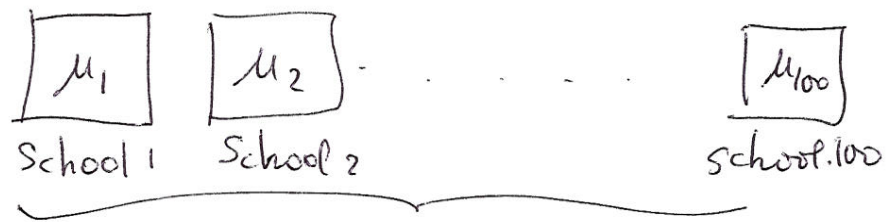


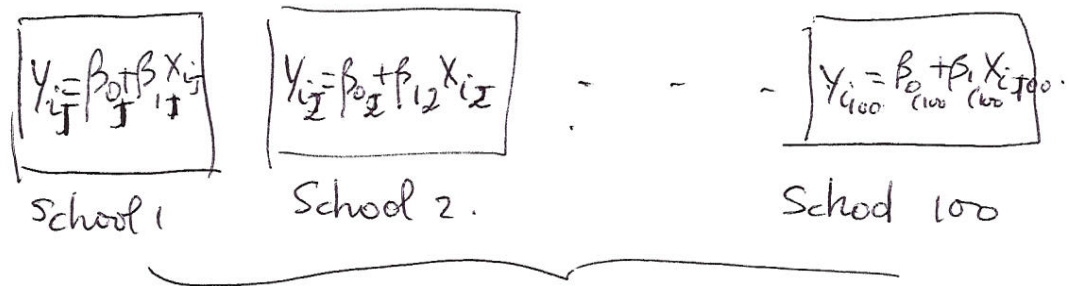
The Normal Hierarchical Regression Model

$m=100$ schools. Y_{ij} = score of student i in school $J=1, \dots, m$.
Chapter 8:



Is there a difference in average score?
 (between schools?)

Chapter 11:



Are the regressions (relation SES and score) different between schools?

General model for any # of indep. variables. (1)

Y_J = vector of scores of n_J children in school $J=1, \dots, m$.
 X_J = Matrix ($n_J \times p$) with p variates school J $m=100$
 β_J = Regression coefficient vector for school J .

Model within group (within school)

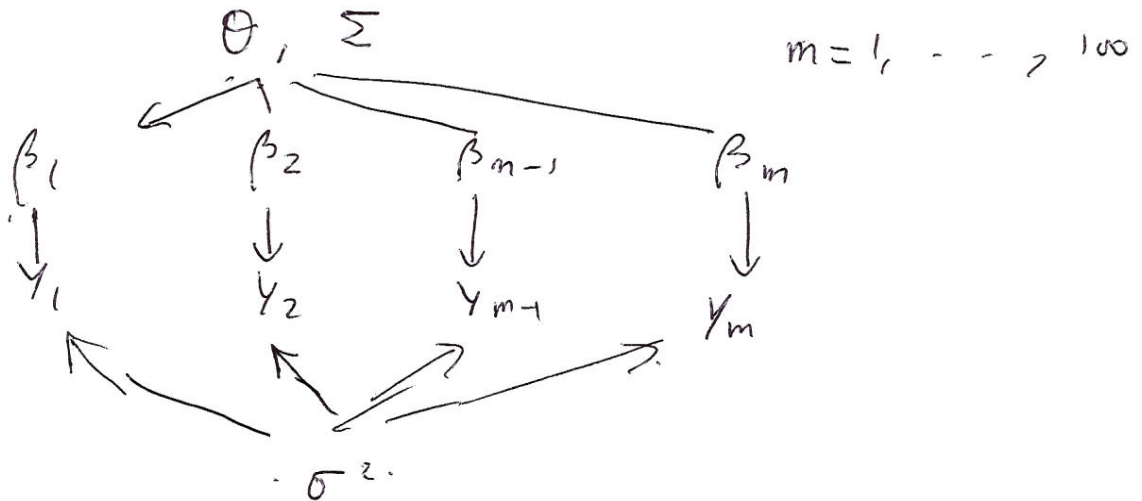
$$Y_J \sim \text{MVN}(X_J \beta_J, \sigma^2 I) \quad \sigma^2 = \text{error variance}$$

Notice: μ_J is now $X_J \beta_J$. $J=1, \dots, 100$

Model ~~within~~ between schools

$$\beta_J \sim \text{MVN}(\theta, \Sigma) \quad \begin{aligned} \theta &= \text{Population mean} \\ \Sigma &= \text{Population variance} \end{aligned}$$

Unknowns: $\theta, \Sigma, \beta_j, \sigma^2$



Alternative parametrization: Linear mixed (2) effect model.

$$\beta_j = \theta + \gamma_j. \quad Y_{ij} = \theta^T X_{ij} + \gamma_j^T X_{ij} + \epsilon_{ij}.$$

$$Y_1, \dots, Y_m \sim \text{iid } \text{MVN}(0, \Sigma).$$

θ = fixed effects $\gamma_1, \dots, \gamma_m$ = random effects.

Regressors could be different for fixed and random effects.

$$Y_{ij} = \theta^T X_{ij} + \gamma_j^T Z_{ij} + \epsilon_{ij}.$$

Regardless of parametrization

Given Prior for θ, Σ, σ^2 ,

Having observed $\vec{Y}_1, \vec{Y}_2, \dots, \vec{Y}_m \quad m=1, \dots, 100$

Posterior distribution

Many parameters: $P(\beta_1, \dots, \beta_m, \theta, \Sigma, \sigma^2 | X_1, \dots, X_m, Y_1, \dots, Y_m)$

Full Conditional Distributions

Information across groups is shared via $\theta, \Sigma,$ and σ^2 .

Conditional on θ, Σ, σ^2 , the β_J elements $\beta_1, \beta_2, \dots, \beta_m$ are independent. $m=100$.

① $(\beta_J | Y_J, X_J, \theta, \Sigma, \sigma^2) \sim \text{MVN}$ with

$$\text{Var}(\beta_J | Y_J, X_J, \sigma^2, \theta, \Sigma) = (\Sigma^{-1} + X_J^T X_J / \sigma^2)^{-1}$$

$$E(\beta_J | Y_J, X_J, \sigma^2, \theta, \Sigma) = (\Sigma^{-1} + X_J^T X_J / \sigma^2)^{-1} (\Sigma^{-1} \theta + X_J^T Y_J / \sigma^2)$$

② $(\theta | \beta_1, \dots, \beta_m, \Sigma) \sim \text{MVN}$ with. (May want to look at chapter 7)

mean $\mu_m = \Lambda_m^{-1} (\Lambda_0^{-1} \mu_0 + m \bar{\beta})$

Notice $\bar{\beta} = \text{average of } \beta.$
 $= \frac{1}{m} \sum \beta_J$

Variance $\Lambda_m = (\Lambda_0^{-1} + m \Sigma^{-1})^{-1}$

③ $(\Sigma | \theta, \beta_1, \dots, \beta_m) \sim \text{Inverse-Wishart}(n_0 + m, (S_0 + S_\theta)^{-1})$

$$S_\theta = \sum_{J=1}^m (\beta_J - \theta)(\beta_J - \theta)^T$$

④ $\sigma^2 \sim \text{IG}([v_0 + \sum n_J] / 2, [v_0 \sigma_0^2 + \text{SSR}] / 2)$

$$\text{SSR} = \sum_{J=1}^m \sum_{i=1}^{n_J} (y_{i,J} - \beta_J^T x_{i,J})^2$$

YOU INPUT

Note: all items $\lambda_0, \mu_0, \tau_0, S_0, \nu_0, \sigma_0, m$ must be provided (not computed by Gibbs or Metropolis sampling).

→ $\mu_0 =$ prior expectation of θ .

(i) = Average of ordinary least squared regression estimates
or

(ii) If you have prior info from literature, put your own numbers

In (i) ..

$$\mu_0 = \begin{pmatrix} \overline{\hat{\beta}}_0 \\ \vdots \\ \overline{\hat{\beta}}_p \end{pmatrix} = \begin{pmatrix} \frac{\sum_{j=1}^m \hat{\beta}_{0(j)}}{m} \\ \vdots \\ \frac{\sum_{j=1}^m \hat{\beta}_{p(j)}}{m} \end{pmatrix}$$

Say, for school 1 you get

$$\hat{\beta}_0 = 0.2 \quad \hat{\beta}_1 = 3 \quad \hat{\beta}_2 = 2$$

for school 2 you get

$$\hat{\beta}_0 = 0.1 \quad \hat{\beta}_1 = 1.5 \quad \hat{\beta}_3 = 1.$$

$$\text{Then } \overline{\hat{\beta}}_0 = \frac{0.2+0.1}{2} \quad \overline{\hat{\beta}}_1 = \frac{3+1.5}{2} \quad \overline{\hat{\beta}}_2 = \frac{2+1}{2}$$

→ $\Lambda_0 =$ prior variance of θ

⇒ Think that we have $m=100$ sets of $\hat{\beta}$, i.e. 100 $\hat{\beta}_1$, 4 for each school, 100 $\hat{\beta}_2$, etc.

$$\text{Var}(\hat{\beta}_1) = \frac{\sum_{j=1}^m (\hat{\beta}_{1(j)} - \bar{\beta}_1)^2}{m-1}$$

$$\text{Var}(\hat{\beta}_2) = \frac{\sum_{j=1}^m (\hat{\beta}_{2(j)} - \bar{\beta}_2)^2}{m-1}$$

$$\text{Cov}(\hat{\beta}_1, \hat{\beta}_2) = \frac{\sum (\hat{\beta}_{1(j)} - \bar{\beta}_1)(\hat{\beta}_{2(j)} - \bar{\beta}_2)}{m-1}$$

So you calculate all this,

$$\begin{pmatrix} \text{Var}(\hat{\beta}_1) & \text{Cov}(\hat{\beta}_1, \hat{\beta}_2) & \dots & \text{Cov}(\hat{\beta}_1, \hat{\beta}_p) \\ \text{Cov}(\hat{\beta}_2, \hat{\beta}_1) & \text{Var}(\hat{\beta}_2) & \dots & \text{Cov}(\hat{\beta}_2, \hat{\beta}_p) \\ \text{Cov}(\hat{\beta}_3, \hat{\beta}_1) & \text{Cov}(\hat{\beta}_3, \hat{\beta}_2) & \text{Var}(\hat{\beta}_3) & \dots \\ \vdots & \vdots & \vdots & \ddots \\ \text{Cov}(\hat{\beta}_p, \hat{\beta}_1) & \text{Cov}(\hat{\beta}_p, \hat{\beta}_2) & \dots & \text{Var}(\hat{\beta}_p) \end{pmatrix}$$

This and the m_0 given represent information of a person with unbiased but weak priors

→ $\eta_0 = \#$ of parameters in the regression + 2
 $= p + 2$.

This makes the $E(\Sigma) = S_0$ but the prior distribution is diffuse.

→ $S_0 =$ also equal to $\Delta = 0$.

→ $\sigma_0^2 =$ average of within group sample variance

i.e. Compute $S_1^2, S_2^2, \dots, S_m^2$.

$\sum_{j=1}^m \frac{S_j^2}{m}$ is your σ_0^2 .

recall $S_j^2 = \frac{\sum (Y_{ij} - \bar{Y}_j)^2}{n_j - 1}$.

→ $\nu_0 = 1$

Please, read section 11.1 - 11.3.

What if the Normal Model
for Y is not good?

Can we do regression still?

Yes, Generalized linear model. Assume
 $Y \sim$ some distribution (Poisson, Binomial)
and use a link to regression

e.g. $Y_{iJ} \sim \text{Poisson}(\lambda_{iJ})$.

$$\log(\lambda_{iJ}) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \varepsilon_p$$

Some θ, Σ have similar
full conditional distributions but
not the β 's.

Need Metropolis Hastings.

What is Metropolis - Hastings?

Chapter 10.

and 11.4 - end. of chapter 10.