

University of California, Los Angeles  
Department of Statistics

Statistics 100B

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Practice 1 - solutions

**EXERCISE 1**

Find the distribution of the random variable  $X$  for each of the following moment-generating functions:

- a.  $M_X(t) = \left[\frac{1}{3}e^t + \frac{2}{3}\right]^5$ .  $X$  follows binomial with  $n = 5$ ,  $p = \frac{1}{3}$ .
- b.  $M_X(t) = \frac{e^t}{2-e^t} = \frac{\frac{1}{2}e^t}{1-\frac{1}{2}e^t}$ .  $X$  follows geometric with  $p = \frac{1}{2}$ .
- c.  $M_X(t) = e^{2(e^t-1)}$ .  $X$  follows Poisson with  $\lambda = 2$ .

**EXERCISE 2**

Let  $M_X(t) = \frac{1}{6}e^t + \frac{2}{6}e^{2t} + \frac{3}{6}e^{3t}$  be the moment-generating function of a random variable  $X$ .

- a. To find  $E(X)$  we need the first derivative of  $M_X(t)$  w.r.t.  $t$ :  
 $M'_X(t) = \frac{1}{6}e^t + \frac{2}{6}2e^{2t} + \frac{3}{6}3e^{3t}$ . The first moment is the previous derivative evaluated at  $t = 0$ . The result is  $E(X) = \frac{1}{6} + \frac{4}{6} + \frac{9}{6} = \frac{7}{3}$ .
- b. To find  $\text{var}(X)$  we need the second moment. We need the second derivative of  $M_X(t)$  w.r.t.  $t$ .  
 $M''_X(t) = \frac{1}{6}e^t + \frac{8}{6}2e^{2t} + \frac{27}{6}3e^{3t}$ . And at  $t = 0$  we obtain the second moment. It is equal to  $EX^2 = \frac{1}{6} + \frac{8}{6} + \frac{27}{6} = 6$ . The variance is equal to:  $\text{var}(X) = EX^2 - (E(X))^2 = 6 - \left(\frac{7}{3}\right)^2 = \frac{5}{9}$ .
- c. From the definition of moment-generating functions  $M_X(t) = Ee^{tX}$  we see that  $X$  is discrete with possible values 1, 2, and 3, and corresponding probabilities  $\frac{1}{6}$ ,  $\frac{2}{6}$ , and  $\frac{3}{6}$ .

**EXERCISE 3**

Let  $X$  follow the Poisson probability distribution with parameter  $\lambda$ . Its moment-generating function is  $M_X(t) = e^{\lambda(e^t-1)}$ .

- a. The moment-generating function of  $Z = \frac{X-\lambda}{\sqrt{\lambda}}$  is:

$$M_Z(t) = M_{\frac{X-\lambda}{\sqrt{\lambda}}}(t) = e^{-\frac{\lambda}{\sqrt{\lambda}}t} M_X\left(\frac{t}{\sqrt{\lambda}}\right) \Rightarrow M_Z(t) = e^{-\sqrt{\lambda}t} e^{\lambda(e^{\frac{t}{\sqrt{\lambda}}}-1)}$$

- b. Using the series expansion of

$$e^{\frac{t}{\sqrt{\lambda}}} = 1 + \frac{\frac{t}{\sqrt{\lambda}}}{1!} + \frac{\left(\frac{t}{\sqrt{\lambda}}\right)^2}{2!} + \frac{\left(\frac{t}{\sqrt{\lambda}}\right)^3}{3!} + \dots$$

we get:

$$M_Z(t) = e^{-\sqrt{\lambda}t - \lambda + \lambda\left[1 + \frac{t}{\sqrt{\lambda}} + \frac{t^2}{2\lambda} + \frac{t^3}{3!(\sqrt{\lambda})^3} + \dots\right]} = e^{-\sqrt{\lambda}t - \lambda + \lambda + \sqrt{\lambda}t + \frac{t^2}{2} + \frac{\lambda t^3}{3!(\sqrt{\lambda})^3} + \dots}$$

Therefore,

$$\lim_{\lambda \rightarrow \infty} M_Z(t) = e^{\frac{1}{2}t^2}.$$

In other words, as  $\lambda \rightarrow \infty$ , the ratio  $Z = \frac{X-\lambda}{\sqrt{\lambda}}$  converges to the standard normal distribution.

**EXERCISE 4**

Here we use the normal approximation to Poisson. It is given that  $X$  follows the Poisson distribution with  $\lambda = 100$ . We know that for large  $\lambda$  the ratio  $Z = \frac{X-\lambda}{\sqrt{\lambda}}$  follows the standard normal distribution. Therefore:

$$P(X \leq 110) = P\left(Z < \frac{110.5 - 100}{\sqrt{100}}\right) = P(Z < 1.05) = 0.8531.$$

The exact probability is

$$P(X \leq 110) = \sum_{x=0}^{110} \frac{e^{-100} 100^x}{x!} = \dots = 0.8529.$$

The approximation is not bad!

**EXERCISE 5**

We know that the moment generating function of a normal random variable is  $e^{\mu t + \frac{1}{2}\sigma^2 t^2}$ . Because the sample is i.i.d the moment generating function of  $T = \sum_{i=1}^n X_i$  is:

$M_T(t) = Ee^{Tt} = Ee^{(X_1 + \dots + X_n)t} = Ee^{X_1 t} \dots Ee^{X_n t} \Rightarrow M_T(t) = e^{n\mu t + \frac{1}{2}n\sigma^2 t^2}$ . This is the moment generating function of a normal random variable with mean  $n\mu$  and variance  $n\sigma^2$ . Therefore  $T \sim N(n\mu, \sigma\sqrt{n})$ .

**EXERCISE 6**

The two sample are independent with  $X \sim N(\mu_1, \sigma_1)$  and  $Y \sim N(\mu_2, \sigma_2)$ .

a.  $E(\bar{X} - \bar{Y}) = E(\bar{X}) - E(\bar{Y}) = \mu_1 - \mu_2$ .

b.  $Var(\bar{X} - \bar{Y}) = Var(\bar{X}) + Var(\bar{Y}) - 2Cov(\bar{X}, \bar{Y}) = \frac{\sigma_1^2}{n} + \frac{\sigma_2^2}{n}$ . The covariance is zero because the two samples are independent.

c.  $M_{\bar{X}-\bar{Y}} = Ee^{(\bar{X}-\bar{Y})t} = Ee^{\bar{X}t}Ee^{-\bar{Y}t} = M_{X_1}(\frac{t}{m}) \cdots M_{X_n}(\frac{t}{m})M_{Y_1}(-\frac{t}{n}) \cdots M_{Y_n}(-\frac{t}{n}) = e^{m(\mu_1 \frac{t}{m} + \frac{1}{2}\sigma_1^2 \frac{t^2}{m^2})} e^{-n(\mu_2 \frac{t}{n} - \frac{1}{2}\sigma_2^2 \frac{t^2}{n^2})} = e^{t(\mu_1 - \mu_2) + \frac{1}{2}t^2(\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n})}$ . This is the moment generating function of a normal random variable with mean  $\mu_1 - \mu_2$  and variance  $\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}$ .

d. It is given that  $n = m$  and  $\sigma_1^2 = 2, \sigma_2^2 = 2.5$ . We want  $P(-1 < \bar{X} - \bar{Y} - (\mu_1 - \mu_2) < 1) = 0.95$ .

Since  $\bar{X} - \bar{Y} \sim N(\mu_1 - \mu_2, \sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}})$  we get:

$$P\left(\frac{-1}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} < Z < \frac{1}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}}\right) = 0.95 \Rightarrow P\left(\frac{-1}{\sqrt{\frac{2}{n} + \frac{2.5}{n}}} < Z < \frac{1}{\sqrt{\frac{2}{n} + \frac{2.5}{n}}}\right) = 0.95$$

Therefore,  $1.96 = \frac{1}{\sqrt{\frac{2}{n} + \frac{2.5}{n}}} \Rightarrow 1.96 = \frac{1}{\sqrt{\frac{4.5}{n}}} \Rightarrow n = 17.3 \approx 18$ .

**EXERCISE 7**

We are given that  $X$  follows the normal distribution with  $\mu = 0$  and  $\sigma^2 = 1$ . Therefore its probability density function (p.d.f) is  $f(x) = \frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}x^2}$ ,  $-\infty < x < \infty$ . The random variable  $Y$  is said to follow the lognormal distribution if  $Y = e^X$  because  $\log(y)$  follows the normal distribution. To find its p.d.f we start with its cumulative distribution function which by definition is:

$$F(y) = P(Y \leq y) = P(e^X \leq y) = P(X \leq \log(y)) = F_x(\log(y)) \Rightarrow f(y) = \frac{1}{y} f_x(\log(y)) \Rightarrow \frac{1}{y} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\log(y))^2}$$

**EXERCISE 8**

The radius of a circle,  $X$ , is an exponential random variable with parameter  $\lambda$ . Therefore its p.d.f. is  $f(x) = \lambda e^{-\lambda x}$ . Let  $Y$  be the area of the circle. Then  $Y = \pi X^2$ . We want to find the p.d.f of  $Y$ . We start with the c.d.f. function:  $F(y) = P(Y \leq y) = P(\pi X^2 \leq y) = P(X^2 \leq \frac{y}{\pi}) = P(-\sqrt{\frac{y}{\pi}} \leq X \leq +\sqrt{\frac{y}{\pi}}) = F_X(\sqrt{\frac{y}{\pi}}) - F_X(-\sqrt{\frac{y}{\pi}}) \Rightarrow f(y) = F(y)' = \frac{1}{2\sqrt{\pi y}} f_X(\sqrt{\frac{y}{\pi}}) - 0 \Rightarrow f(y) = \frac{1}{2\sqrt{\pi y}} \lambda e^{-\lambda \sqrt{\frac{y}{\pi}}}$ .