

University of California, Los Angeles  
Department of Statistics

Statistics 100B

Instructor: Nicolas Christou

Practice 2 - solutions

**EXERCISE 1**

This population has mean  $\mu = 3$ , and standard deviation  $\sigma = 0.28$ .

- a. According to the central limit theorem  $\bar{X}$  is distributed as  $\bar{X} \sim N(3, \frac{0.28}{\sqrt{36}})$  or  $\bar{X} \sim N(3, 0.047)$ . We expect that the range  $3 \pm 3(0.047)$  or  $3 \pm 0.141$  or  $(2.86, 3.141)$  will cover almost all the area. We conclude that the histogram should have been much narrower.
- b. According to the central limit theorem  $T \sim N(36(3), 0.28\sqrt{36})$  or  $T \sim N(108, 1.68)$ . The histogram should be centered around 108 with spread  $(103, 113)$ .

**EXERCISE 2**

It is given  $X \sim N(2700, 400)$ . The total supply for  $n = 12$  weeks is  $4000 + 12(2500) = 34000$ . We want the supply to be below 2000 pounds or the total sugar use in these 12 weeks to be more than 32000 pounds:

$$P(T > 32000) = P\left(Z > \frac{32000 - 12(2700)}{400\sqrt{12}}\right) = P(Z > -0.29) = 0.6141.$$

**EXERCISE 3**

We know the the moment generating function of  $N(\mu, \sigma)$  is  $M_X(t) = e^{\mu t + \frac{1}{2}t^2\sigma^2}$ .

- a. Moment generating function of  $X + Y$ :

$$M_{X+Y}(t) = M_X(t)M_Y(t) = e^{2\mu t + t^2\sigma^2}$$

Moment generating function of  $X - Y$ :

$$M_{X-Y}(s) = M_X(s)M_{-Y}(s) = e^{s^2\sigma^2}$$

- b. Joint moment generating function of  $X + Y, X - Y$ :

$$\begin{aligned} M_{X+Y, X-Y}(t, s) &= Ee^{(X+Y)t + (X-Y)s} \\ &= Ee^{X(t+s) + Y(t-s)} \\ &= M_X(t+s)M_Y(t-s) \\ &= e^{\mu(t+s) + \frac{1}{2}(t+s)^2\sigma^2} e^{\mu(t-s) + \frac{1}{2}(t-s)^2\sigma^2} \\ &= e^{2\mu t + t^2\sigma^2} e^{t^2\sigma^2} = M_{X+Y}(t)M_{X-Y}(s). \end{aligned}$$

- c. Since the joint moment generating function of  $X + Y$  and  $X - Y$  can be expressed as the product of the moment generating functions of  $X + Y$  and  $X - Y$  we conclude that  $X + Y$  and  $X - Y$  are independent.

**EXERCISE 4**

- a. We can write  $X_1 - 2X_2 + X_3$  as  $\mathbf{a}'\mathbf{X}$  where  $\mathbf{a}' = (1, -2, 1)$ . Therefore

$$var(\mathbf{aX}) = \mathbf{a}'\Sigma\mathbf{a} = \begin{pmatrix} 1 & -2 & 1 \end{pmatrix} \begin{pmatrix} 5 & 2 & 3 \\ 2 & 3 & 0 \\ 3 & 0 & 3 \end{pmatrix} \begin{pmatrix} 1 \\ -2 \\ 1 \end{pmatrix} = 18.$$

- b. Let  $\mathbf{A} = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 1 \end{pmatrix}$ . Then  $\mathbf{Y} = \begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} = \mathbf{AX}$ . Therefore,

$$var(\mathbf{Y}) = \mathbf{A}\Sigma\mathbf{A}' = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 1 \end{pmatrix} \begin{pmatrix} 5 & 2 & 3 \\ 2 & 3 & 0 \\ 3 & 0 & 3 \end{pmatrix} \begin{pmatrix} 1 & 1 \\ 1 & 1 \\ 0 & 1 \end{pmatrix} = \begin{pmatrix} 12 & 15 \\ 15 & 21 \end{pmatrix}.$$

### EXERCISE 5

- a. Let  $k$  be the minimum number of plays the casino must win. Then  $1 \times k - (10000 - k) \times 35 > 400$ . Solve for  $k$  to get  $k = 9734$ . Let  $Y$  be the number of games the casino must win, so  $Y \sim b(10000, \frac{37}{38})$ . The casino must win at least 9734 plays and this probability is  $P(Y \geq 9734) = \sum_{y=9734}^{10000} \binom{10000}{y} \left(\frac{37}{38}\right)^y \left(\frac{1}{38}\right)^{10000-y}$ . Using normal approximation to binomial:  $P(Y \geq 9734) = P(Z > \frac{9733.5 - 10000 \frac{37}{38}}{10000 \sqrt{\frac{37}{38} \frac{1}{38}}}) = P(Z > -0.21) = 0.5832$ .
- b. If we view the 10000 outcomes as a random sample from the following distribution then we can use the central limit theorem: Let  $T = X_1 + \dots + X_{10000}$  be the sum of the 10000 outcomes.

$X$	$P(X)$
1	$\frac{37}{38}$
-35	$\frac{1}{38}$

This distribution has  $\mu = 0.05263$  and  $\sigma = 5.76$ .

$$P(T > 400) = P(Z > \frac{400 - 10000(0.05263)}{5.76\sqrt{10000}}) = P(Z > -0.22) = 0.5871.$$

### EXERCISE 6

We have  $\mathbf{Z} = \begin{pmatrix} X \\ Y \end{pmatrix} \sim N_2(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  and therefore the joint moment generating function of  $(X_i, Y_i)$  is  $e^{\mathbf{t}'\boldsymbol{\mu} + \frac{1}{2}\mathbf{t}'\boldsymbol{\Sigma}\mathbf{t}}$ . The joint moment generating function of  $(\bar{X}, \bar{Y})$  is:

$$\begin{aligned} Ee^{t_1\bar{X} + t_2\bar{Y}} &= Ee^{t_1 \frac{X_1 + \dots + X_n}{n} + t_2 \frac{Y_1 + \dots + Y_n}{n}} \\ &= Ee^{t_1 \frac{X_1}{n} + t_2 \frac{Y_1}{n}} \times \dots \times Ee^{t_1 \frac{X_n}{n} + t_2 \frac{Y_n}{n}}, \text{ because the pairs } (X_i, Y_i) \text{ are independent.} \end{aligned}$$

Each one of these expectations is the joint moment generating function of  $\mathbf{AZ}$  with  $\mathbf{A} = \begin{pmatrix} \frac{1}{n} & 0 \\ 0 & \frac{1}{n} \end{pmatrix}$  and  $\mathbf{Z} = \begin{pmatrix} X \\ Y \end{pmatrix}$ . Since  $E(\mathbf{AZ}) = \frac{\boldsymbol{\mu}}{n}$  and  $\text{var}(\mathbf{AZ}) = \frac{\boldsymbol{\Sigma}}{n^2}$ , it follows that the joint moment generating function of  $\mathbf{AZ}$  is  $e^{\mathbf{t}'\frac{\boldsymbol{\mu}}{n} + \frac{1}{2}\mathbf{t}'\frac{\boldsymbol{\Sigma}}{n^2}\mathbf{t}}$ . But we have  $n$  independent pairs, therefore the joint moment generating function of  $(\bar{X}, \bar{Y})$  is  $\left(e^{\mathbf{t}'\frac{\boldsymbol{\mu}}{n} + \frac{1}{2}\mathbf{t}'\frac{\boldsymbol{\Sigma}}{n^2}\mathbf{t}}\right)^n = e^{\mathbf{t}'\boldsymbol{\mu} + \frac{1}{2}\mathbf{t}'\frac{\boldsymbol{\Sigma}}{n}\mathbf{t}}$ . This shows that the joint distribution of  $(\bar{X}, \bar{Y})$  is bivariate normal  $N_2(\boldsymbol{\mu}, \frac{\boldsymbol{\Sigma}}{n})$ .

### EXERCISE 7

We write  $\mathbf{Y} = \begin{pmatrix} Y_1 \\ Y_2 \\ Y_3 \end{pmatrix} = \mathbf{AX}$ , where  $\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} \sim N_3(\mathbf{0}, \mathbf{I})$  and  $\mathbf{A} = \begin{pmatrix} \frac{1}{\sqrt{3}} & \frac{1}{\sqrt{3}} & \frac{1}{\sqrt{3}} \\ \frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} & 0 \\ \frac{1}{\sqrt{6}} & \frac{1}{\sqrt{6}} & -\frac{2}{\sqrt{6}} \end{pmatrix}$ . Therefore,  $\text{var}(\mathbf{Y}) = \text{var}(\mathbf{AX}) = \mathbf{AA}' = \mathbf{I}_3$ , which means  $Y_1, Y_2, Y_3$  are i.i.d.  $N(0, 1)$ .

### EXERCISE 8

Let  $\mathbf{X} = (X_1, X_2, X_3)$  has joint moment generating function

$$M_{\mathbf{X}}(t_1, t_2, t_3) = (1 - t_1 + 2t_2)^{-4}(1 - t_1 + 3t_3)^{-3}(1 - t_1)^{-2}.$$

Answer the following questions:

- Find the moment generating function of  $(X_1, X_3)$ .  
 $M_{X_1, X_3}(t_1, t_3) = M_{\mathbf{X}}(t_1, 0, t_3) = (1 - t_1)^{-6}(1 - t_1 + 3t_3)^{-3}$ .
- Find the moment generating function of  $X_1$ .  
 $M_{X_1}(t_1) = M_{\mathbf{X}}(t_1, 0, 0) = (1 - t_1)^{-9}$ .
- Find the moment generating function of  $X_3$ .  
 $M_{X_3}(t_3) = M_{\mathbf{X}}(0, 0, t_3) = (1 + 3t_3)^{-3}$ .
- Are  $X_1, X_3$  independent?  
 No, because  $M_{X_1, X_3}(t_1, t_3) \neq M_{X_1}(t_1) \times M_{X_3}(t_3)$ .
- Find the moment generating function of  $(X_2, X_3)$ .  
 $M_{X_2, X_3}(t_2, t_3) = M_{\mathbf{X}}(0, t_2, t_3) = (1 + 2t_2)^{-4}(1 + 3t_3)^{-3}$ .
- Are  $X_2, X_3$  independent?  
 $M_{X_2}(t_2) = M_{\mathbf{X}}(0, t_2, 0) = (1 + 2t_2)^{-4}$ . Yes,  $X_2, X_3$  are independent because  $M_{X_2, X_3}(t_2, t_3) = M_{X_2}(t_2) \times M_{X_3}(t_3)$ .