

Homework 3

Exercise 1

THE JOINT PDF OF Y_1 AND Y_2 IS:

$$f(y_1, y_2) = f(y_1) \cdot f(y_2) = e^{-y_1} \cdot e^{-y_2} = e^{-(y_1 + y_2)}$$

(INDEPENDENT)

LET $U = \frac{Y_1}{Y_1 + Y_2}$ AND $V = Y_2$

IT FOLLOWS THAT $Y_2 = V$ AND $Y_1 = \frac{UV}{1-U}$

JACOBIAN:

$$J = \begin{vmatrix} \frac{\partial U}{\partial Y_1} & \frac{\partial U}{\partial Y_2} \\ \frac{\partial V}{\partial Y_1} & \frac{\partial V}{\partial Y_2} \end{vmatrix} = \begin{vmatrix} \frac{Y_2}{(Y_1 + Y_2)^2} & \frac{-Y_1}{(Y_1 + Y_2)^2} \\ 0 & 1 \end{vmatrix} = \frac{Y_2}{(Y_1 + Y_2)^2} = \frac{(1-U)^2}{V}$$

THEREFORE THE JOINT PDF OF U AND V IS:

$$f(u, v) = e^{-\left[\frac{uv}{1-u} + v\right]} \frac{v}{(1-u)^2} = e^{-\frac{v}{1-u}} \frac{v}{(1-u)^2}$$

TO FIND PDF OF U INTEGRATE THE JOINT PDF W.R.T. V :

$$f(u) = \int_0^{\infty} e^{-\frac{v}{1-u}} \frac{v}{(1-u)^2} dv = 1.$$

EXERCISE 2 :

$$(a) X \sim N(0,1), \quad Y \sim N(0,1)$$

X, Y ARE
INDEPENDENT

JOINT PDF OF X, Y:

$$f(x,y) = f(x) \cdot f(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2}$$

$$= \frac{1}{2\pi} e^{-\frac{1}{2}x^2 - \frac{1}{2}y^2}$$

$$\left. \begin{array}{l} U = X+Y \\ V = X-Y \end{array} \right\} \rightarrow x = \frac{U+V}{2} \text{ AND } y = \frac{U-V}{2}$$

JACOBIAN :

$$J = \begin{vmatrix} \frac{\partial U}{\partial X} & \frac{\partial U}{\partial Y} \\ \frac{\partial V}{\partial X} & \frac{\partial V}{\partial Y} \end{vmatrix} = \begin{vmatrix} 1 & 1 \\ 1 & -1 \end{vmatrix} = -2$$

JOINT PDF OF U, V:

$$f(u,v) = \frac{1}{2\pi} e^{-\frac{1}{2}\left(\frac{u+v}{2}\right)^2} \cdot e^{-\frac{1}{2}\left(\frac{u-v}{2}\right)^2} \cdot |-2|^{-1}$$

THIS IS SIMPLIFIED TO:

$$f(u,v) = \frac{1}{\sqrt{2}\sqrt{2\pi}} e^{-\frac{1}{2}\frac{u^2}{2}} \cdot \frac{1}{\sqrt{2}\sqrt{2\pi}} e^{-\frac{1}{2}\frac{v^2}{2}} \rightarrow$$

EXERCISE 3 :

$$X \sim \text{beta}(\alpha, \beta), \quad Y \sim \text{beta}(\alpha + \beta, \gamma)$$

$$f(x, y) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} \cdot \frac{\Gamma(\alpha + \beta + \gamma)}{\Gamma(\alpha + \beta)\Gamma(\gamma)} x^{\alpha + \beta - 1} (1-x)^{\gamma-1}$$

$$\left. \begin{array}{l} u = xy \\ v = x \end{array} \right\} \begin{array}{l} y = \frac{u}{v} \\ x = v \end{array}$$

$$J = \begin{vmatrix} \frac{\partial u}{\partial x} & \frac{\partial u}{\partial y} \\ \frac{\partial v}{\partial x} & \frac{\partial v}{\partial y} \end{vmatrix} = \begin{vmatrix} y & x \\ 1 & 0 \end{vmatrix} = -x$$

$$f(u, v) = \frac{\Gamma(\alpha + \beta + \gamma)}{\Gamma(\alpha)\Gamma(\beta)\Gamma(\gamma)} v^{\alpha-1} (1-v)^{\beta-1} \left(\frac{u}{v}\right)^{\alpha + \beta - 1} \left(1 - \frac{u}{v}\right)^{\gamma-1} \frac{1}{v}$$

Exercise 4

(a) PDF of R:

$$P(R \leq r) = 1 - P(R > r) = 1 - P\left(X = 0 \text{ PARTICLES IN SPHERE WITH VOLUME } \frac{4}{3}\pi r^3\right)$$

$$\text{HERE } \lambda = \lambda \frac{4}{3}\pi r^3$$

$$= 1 - \frac{\lambda^0 e^{-\lambda}}{0!} = 1 - \frac{e^{-\lambda \frac{4}{3}\pi r^3}}{1}$$

$$f(r) = 4\lambda\pi r^2 e^{-\lambda \frac{4}{3}\pi r^3}$$



(b) Let $U = R^3$

$$F_U(u) = P(U \leq u) = P(R^3 \leq u)$$

$$= P(R \leq u^{1/3}) = F_R(u^{1/3})$$

$$f_U(u) = \frac{1}{3} u^{-2/3} 4\lambda\pi u^{2/3} e^{-\lambda \frac{4}{3}\pi u}$$

$$f_U(u) = \frac{4}{3}\lambda\pi e^{-\frac{4}{3}\lambda\pi u}$$

$$\therefore U \sim \text{EXP}\left(\frac{4}{3}\lambda\pi\right)$$

$$\text{Ans } E U = \frac{3}{4\lambda\pi}$$

Exercise 5

$$X_1 \sim \chi_{r_1}^2, \quad X_2 \sim \chi_{r_2}^2$$

$$f(x_1, x_2) = f(x_1) f(x_2) = \frac{x_1^{\frac{r_1}{2}-1} e^{-x_1/2}}{\Gamma(\frac{r_1}{2}) 2^{r_1/2}} \cdot \frac{x_2^{\frac{r_2}{2}-1} e^{-x_2/2}}{\Gamma(\frac{r_2}{2}) 2^{r_2/2}}$$

$$J = \begin{vmatrix} \frac{\partial y_1}{\partial x_1} & \frac{\partial y_1}{\partial x_2} \\ \frac{\partial y_2}{\partial x_1} & \frac{\partial y_2}{\partial x_2} \end{vmatrix} = \begin{vmatrix} \frac{1}{x_2} & -\frac{x_1}{x_2^2} \\ 1 & 1 \end{vmatrix} = \frac{1}{x_2} + \frac{x_1}{x_2^2} = \frac{x_2 + x_1}{x_2^2} = \frac{y_2}{x_2^2}$$

$$\left. \begin{aligned} y_1 &= \frac{x_1}{x_2} \\ y_2 &= x_1 + x_2 \end{aligned} \right\} \begin{aligned} x_1 &= \frac{y_1 y_2}{1 + y_1} \\ x_2 &= \frac{y_2}{1 + y_1} \end{aligned}$$

$$f(y_1, y_2) = \frac{\left(\frac{y_1 y_2}{1 + y_1}\right)^{\frac{r_1}{2}-1} e^{-\frac{y_1 y_2}{2(1+y_1)}} \left(\frac{y_2}{1 + y_1}\right)^{\frac{r_2}{2}-1} e^{-\frac{y_2}{2(1+y_1)}}}{\Gamma(\frac{r_1}{2}) 2^{r_1/2} \Gamma(\frac{r_2}{2}) 2^{r_2/2}} \cdot \frac{y_2^2}{(1+y_1)^2 y_2} \cdot \frac{\Gamma(\frac{r_1+r_2}{2})}{\Gamma(\frac{r_1}{2}) \Gamma(\frac{r_2}{2})}$$

$$= \frac{y_1^{\frac{r_1}{2}-1} (1+y_1)^{-\frac{r_1+r_2}{2}} \Gamma(\frac{r_1+r_2}{2})}{\Gamma(\frac{r_1}{2}) \Gamma(\frac{r_2}{2})} \cdot \frac{y_2^{\frac{r_1+r_2}{2}-1} e^{-y_2/2}}{\Gamma(\frac{r_1+r_2}{2}) 2^{r_1+r_2}}$$

$$= f(y_1) \cdot f(y_2)$$

$\therefore y_1, y_2$ ARE INDEPENDENT

AND $y_2 \sim \chi_{r_1+r_2}^2$

SINCE $\frac{x_1}{x_2}$ IND OF x_3 (GIVEN) } $\frac{x_1/r_1}{x_2/r_2}$ IND OF $\frac{x_3/r_3}{x_1+r_2/(r_1+r_2)}$

AND $\frac{x_1}{x_2}$ IND OF x_1+x_2

Exercise 6

$$\begin{aligned}
 \text{(a), } P(X_1, \dots, X_{r-1} | X_r) &= \frac{P(X_1, \dots, X_r)}{P(X_r)} \quad \text{NOTE: } X_r \sim \text{bin}(n, p_r) \\
 &= \frac{\frac{n!}{x_1! x_2! \dots x_r!} p_1^{x_1} p_2^{x_2} \dots p_r^{x_r}}{\frac{n!}{(n-x_r)! x_r!} p_r^{x_r} (1-p_r)^{n-x_r}} \\
 &= \frac{(n-x_r)!}{x_1! x_2! \dots x_{r-1}!} \left(\frac{p_1}{1-p_r}\right)^{x_1} \dots \left(\frac{p_{r-1}}{1-p_r}\right)^{x_{r-1}}
 \end{aligned}$$

$\therefore \sim$ MULTINOMIAL with
 $n - x_r$ # of trials

AND PROBABILITIES $\frac{p_1}{1-p_r}, \frac{p_2}{1-p_r}, \dots, \frac{p_{r-1}}{1-p_r}$

Exercise 7

$$X_i | p_i \sim \text{Bernoulli}(p_i), \quad p_i \sim \text{beta}(\alpha, \beta)$$

$$\text{(a) } 1. E X_i = E[E(X_i | p_i)] = E p_i = \frac{\alpha}{\alpha + \beta}$$

$$E \sum X_i = n \frac{\alpha}{\alpha + \beta}$$

$$2. \text{VAR}(X_i) = E[\text{VAR}(X_i | p_i)] + \text{VAR}(E(X_i | p_i))$$

$$= E[p_i(1-p_i)] + \text{VAR}(p_i)$$

$$\begin{aligned}
&= E p_i - E p_i^2 + \text{VAR}(p_i) \\
&= \frac{\alpha}{\alpha+\beta} - \left(\frac{\alpha^2}{(\alpha+\beta)^2} + \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)} \right) + \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)} \\
&= \frac{\alpha(\alpha+\beta)(\alpha+\beta+1) - \alpha^2(\alpha+\beta+1) - \alpha\beta + \alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)} = \frac{\alpha\beta}{(\alpha+\beta)^2} \\
\therefore \text{VAR}\left(\sum X_i\right) &= n \frac{\alpha\beta}{(\alpha+\beta)^2}
\end{aligned}$$

(b) $X_i | p_i \sim b(n_i, p_i), \quad p_i \sim \text{beta}(\alpha, \beta)$

$$E Y = \frac{\alpha}{\alpha+\beta} \sum n_i$$

$$\text{VAR}(Y) = \sum \text{VAR}(X_i)$$

$$\text{VAR}(X_i) = n_i \frac{\alpha\beta(\alpha+\beta+n_i)}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

Exercise 8

CDF of $Q = \prod X_i$:

$$F_Q(q) = P(Q \leq q)$$

$$= P(\prod X_i \leq q)$$

$$= P(-\ln \prod X_i > \ln q)$$

$$= P(-\ln X_1, \dots, -\ln X_n > \ln q)$$

Now each $-\ln X_i \sim \text{Exp}(1)$ ($= F_T(-\ln q)$)

Then $T = -\sum \ln X_i \sim \Gamma(n, 1)$

Therefore, $F_Q(q) = 1 - F_T(-\ln q)$

$$f(q) = -f_T(-\ln q) = -\frac{1}{q} \frac{(-\ln q)^{n-1} e^{-\ln q}}{\Gamma(n)} = \frac{(-\ln q)^{n-1}}{\Gamma(n)}$$
