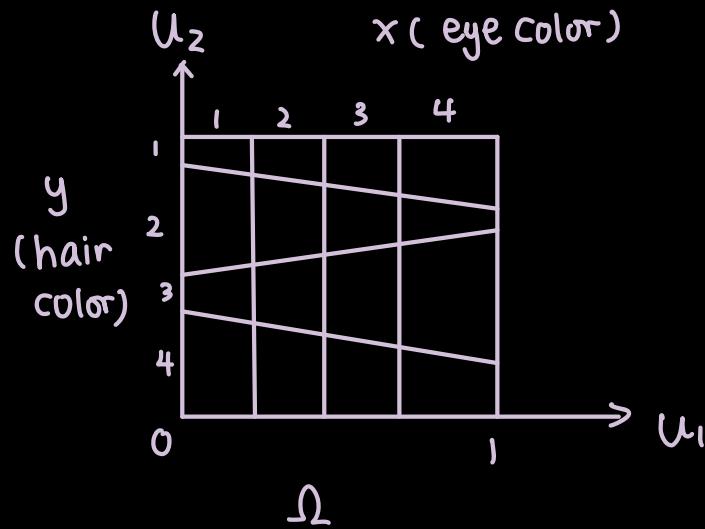


# Gibbs Sampler

(Iterative Conditional Sampler)

Background in Probability (conditioning)

Discrete Case



Random Sample (or throw into) a pair in  $\Omega$

$$P(x,y) = \text{area of cell } (x,y)$$

$$\left. \begin{aligned} P_x(x) &= \text{area of } (X) = \sum_y P(x,y) \\ P_y(y) &= \sum_x P(x,y) \end{aligned} \right\} \begin{array}{l} \text{Operation ① :} \\ \text{marginalization} \end{array}$$

Operation ② : Conditioning

$$P_{Y|X}(y|x) = \frac{P(x,y)}{P_x(x)} \quad \begin{array}{l} \text{Joint} \\ \text{marginal} \end{array}$$

↓  
randomly throw into  $(x)$   
change sample space to  $(x)$

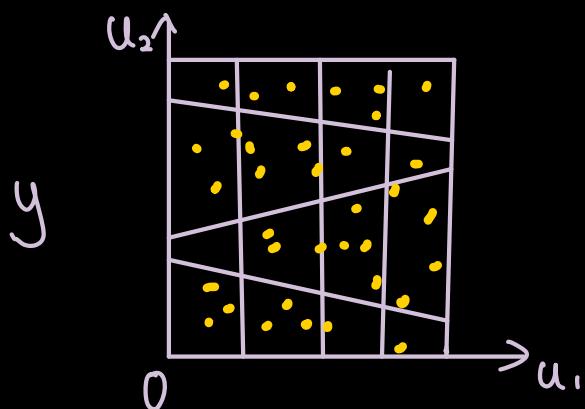
### Operation ③: Factorization

$$P(x,y) = P(x) P(y|x) = \frac{\text{area of } (x,y)}{\text{area of } (x)}$$

↓              ↓              ↓  
 area of  $(x,y)$     area of  $(x)$      $\frac{\text{area of } (x,y)}{\text{area of } (x)}$

A lot of machine learning algorithms are based on these three operations.

### New discrete case



Repeat a large # of trees

$P(x,y) =$  how often fall into  $(x,y)$

$P(x) = \dots - - - - (x)$   
 ↓  
 frequency

$P(y|x) =$  when fall into  $x$ , how often also fall into  $y$   
 ↳ relative frequency

$$P(\text{alarm}|\text{fire}) = 1 \quad P(\text{fire}|\text{alarm}) \approx 0$$

## Back to Gibbs Sampler

Target :  $P(x, y)$

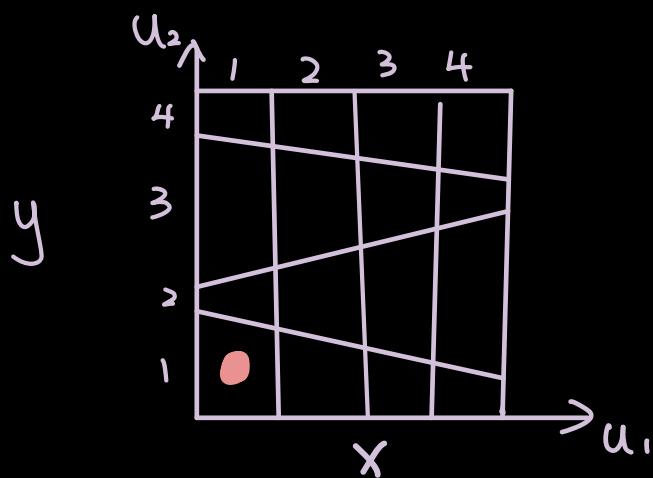
Algorithm : Start from  $(x_0, y_0)$

Iterative,  $t = 1, 2, \dots, T$

Step 1 : Sample  $x_t \sim P(x|y_{t-1})$

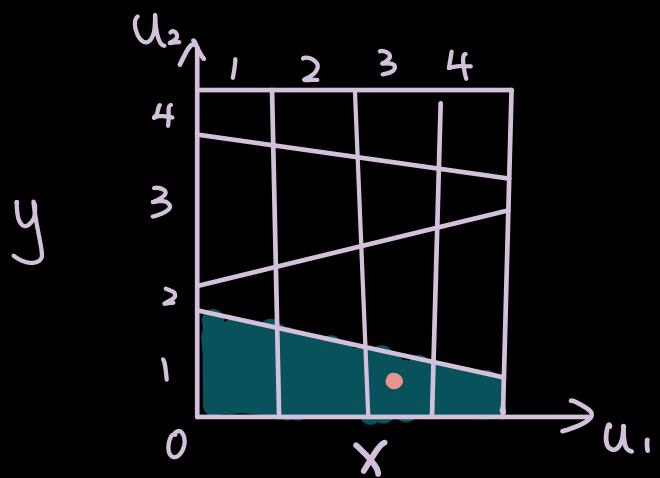
Step 2 : Sample  $y_t \sim (y|x_t)$

## Example :



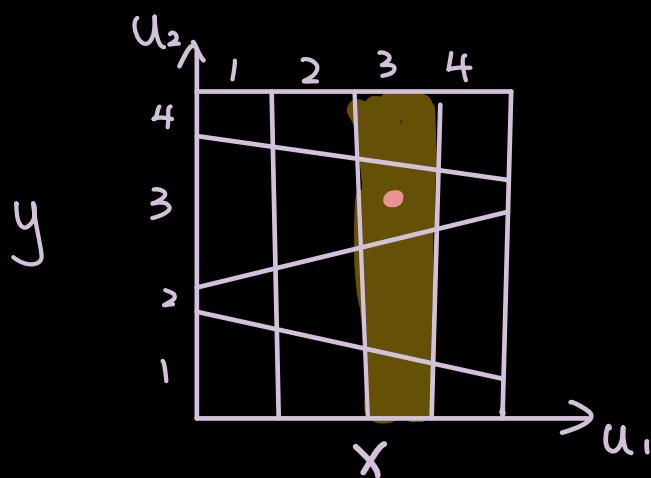
$$(x_0 = 1, y_0 = 1)$$

$X_1 \sim P(X|y=1) \longrightarrow$  randomly relocate  
in row  $(y=1) \dots$



now  $X_1 = 3$  :

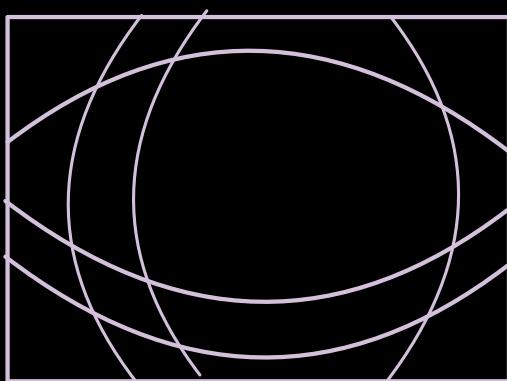
$y_1 \sim P(y_1 | x=3) \rightarrow$  randomly relocate in column ( $x=3$ ) . . .



now  $y_1 = 3$

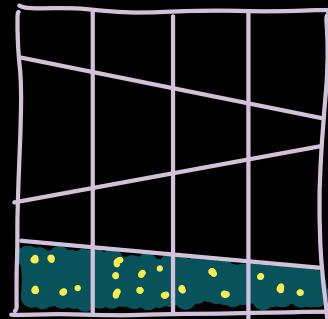
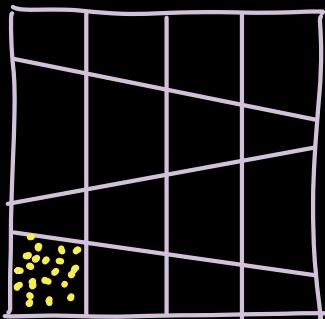
$\xrightarrow[t]{\infty}$  your location becomes  
a random point in  $\Omega$

Example:

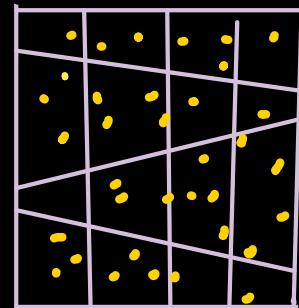
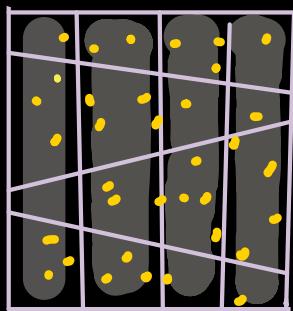


Replace sampling by  
integration  
 $\rightarrow$  local maximum of  
 $P(x, y)$   
"coordinate ascent"

# Population migration



1 million



Uniform over  $\Omega$

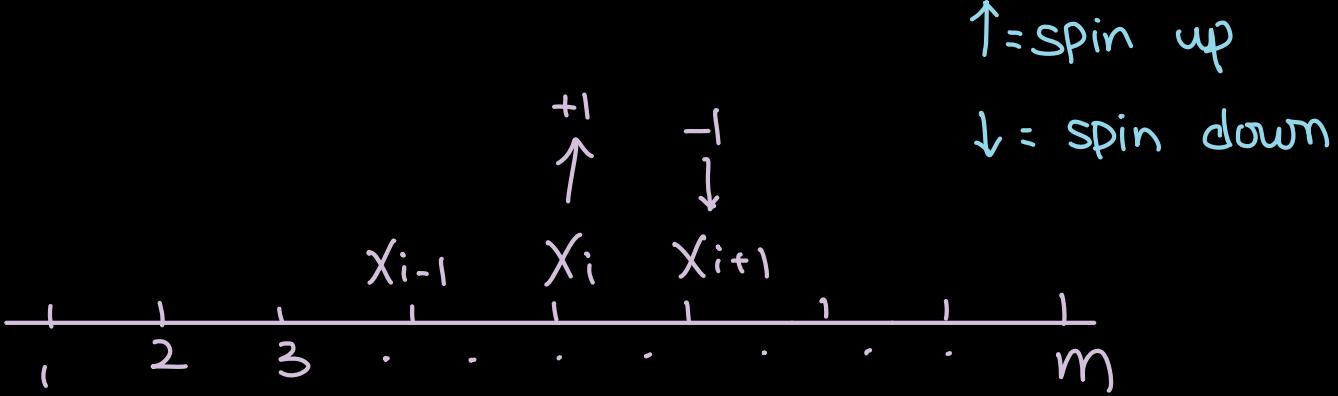
$$p(x, y)$$



Stationary  
equilibrium

Ising model

1 D space



$$X = (x_1, x_2, \dots, x_i, \dots, x_m)$$

energy function

$$\mathcal{E}(x) = - (x_1 x_2 + x_2 x_3 + \dots + x_{i-1} x_i + x_i x_{i+1} +$$

$\downarrow$  gibbs dist       $\dots + x_{m-1} x_m)$

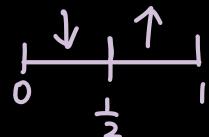
$$\pi(x) = \frac{1}{Z} e^{-\frac{\mathcal{E}(x)}{\tau}} = \frac{1}{Z} e^{\beta \sum_{i=1}^{m-1} x_i x_{i+1}}$$

"Ferromagnetism" , where  $\beta = 1/\tau$

Pseudo code:

Initialize  $X = (x_1, \dots, x_i, \dots, x_m)$

$x_i \stackrel{iid}{\sim}$  Bernoulli ( $\frac{1}{2}$ ) fair coin



Iterate: for (i in 1:m)

Sample  $x_i \sim \pi(x_i | X_{[-i]})$

current values of  $x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_m$

# of possible  $x = 2^m$

local / small changes, feasible

$$\pi(X_i | X_{-i}) \xrightarrow{\text{Conditioning}} \frac{\pi(X_i, X_{-i})}{\pi(X_{-i})}$$

↓ marginalization

$$\pi(X_i = +1, X_{-i}) + \pi(X_i = -1, X_{-i})$$

$$= \frac{1}{Z} e^{\beta \sum_{i=1}^{m-1} X_i X_{i+1}}$$



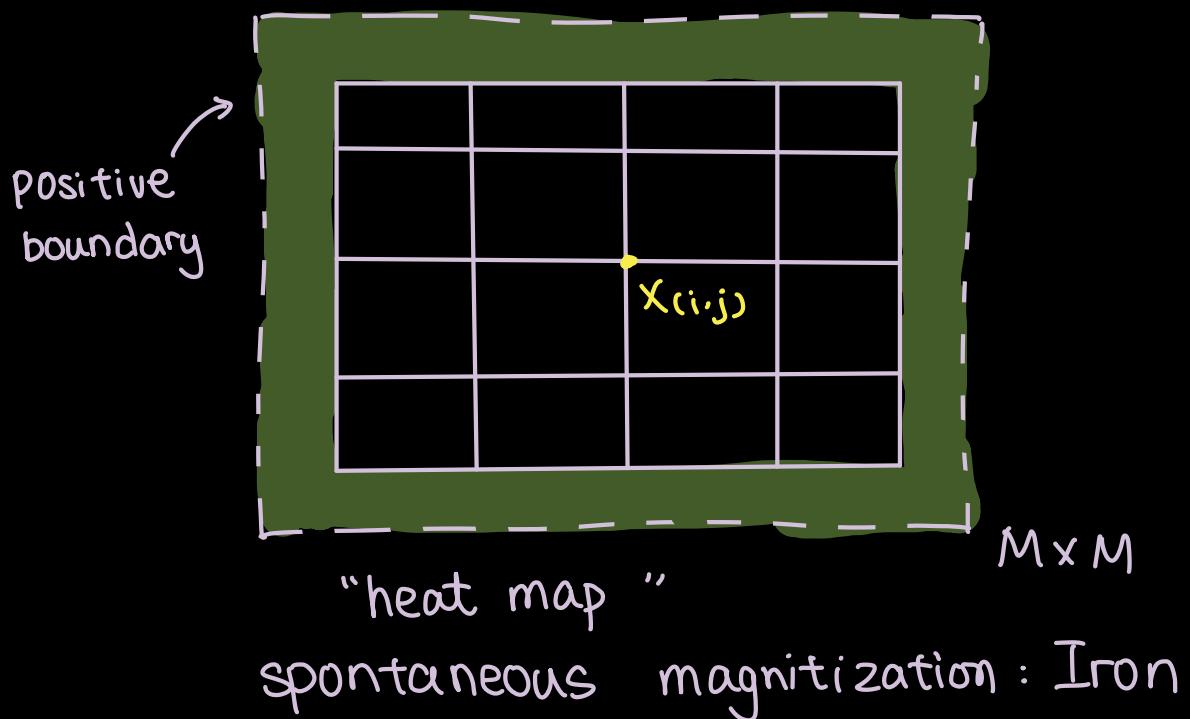
then:

$$= \frac{\cancel{\frac{1}{Z}} e^{\beta \sum_{i=1}^{m-1} X_i X_{i+1}}}{\cancel{\frac{1}{Z}} e^{\beta \sum_{k=1}^{m-1} \underbrace{X_k X_{k+1}}_{X_i = +1}} + \cancel{\frac{1}{Z}} e^{\beta \sum_{k=1}^{m-1} \underbrace{X_k X_{k+1}}_{X_i = -1}}}$$

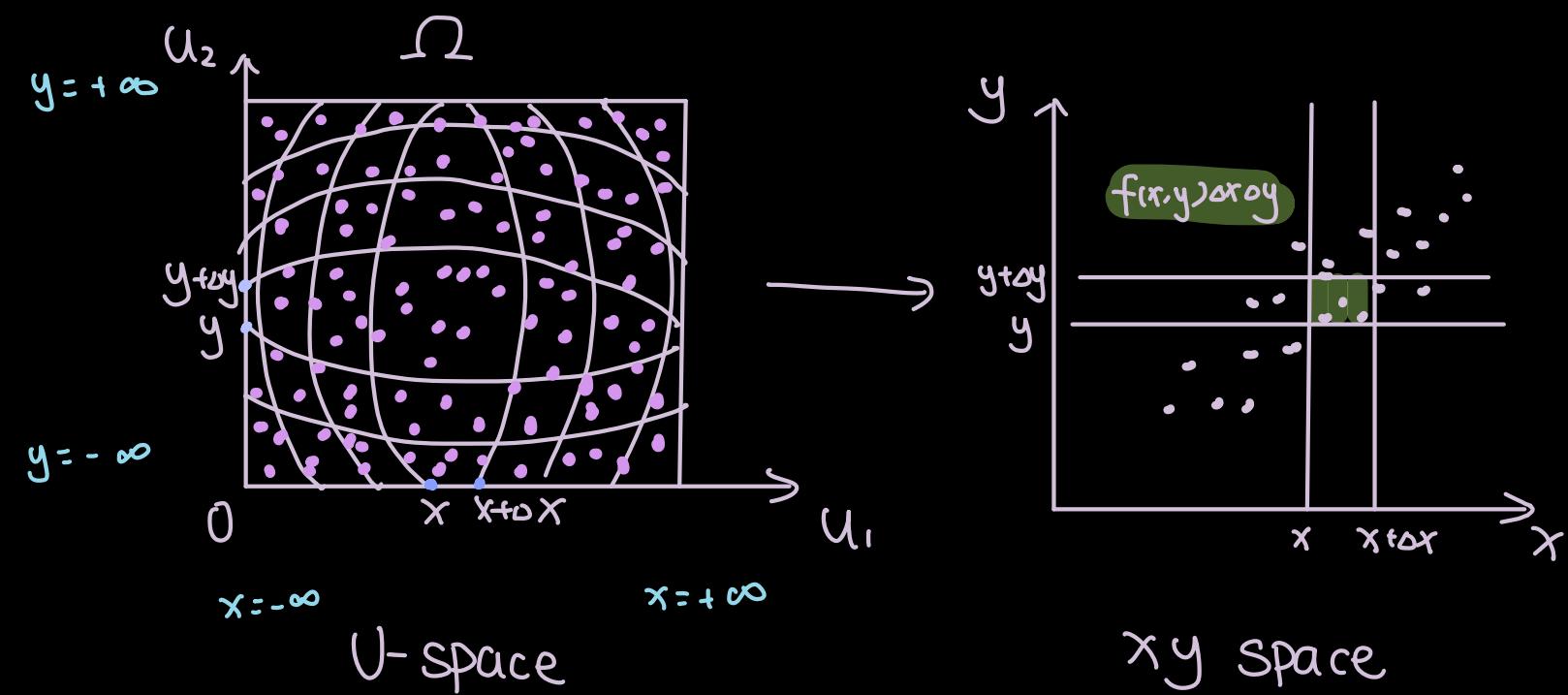
$$\begin{aligned} \pi(X_i = +1 | X_{-i}) &= \frac{e^{\beta(x_{i-1} + x_{i+1})}}{e^{\beta(x_{i-1} + x_{i+1})} + e^{-\beta(x_{i-1} + x_{i+1})}} \\ &= \pi(X_i = +1 | X_{i-1}, X_{i+1}) = \rho \end{aligned}$$

only depends on the spin of its  
two neighbors

# 2D Space



## Continuous Case



① Marginalization :  $f_x(x) \cancel{\Delta x} = \sum_y f(x, y) \cancel{\Delta x} \cancel{\Delta y}$

$$= \int f(x, y) dy$$

$$\cdot f_Y(y) = \int f(x, y) dx$$

② Conditioning :

$$f_{Y|X}(y|x) dy = \frac{f(x, y) dx dy}{f_X(x) dx}$$

$$P(Y \in (y, y+dy) \mid X \in (x, x+dx))$$

$$\cdot f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

③ Factorization :

$$\begin{aligned} f(x, y) &= f(x) f(y|x) \\ &= f(y) f(x|y) \end{aligned}$$

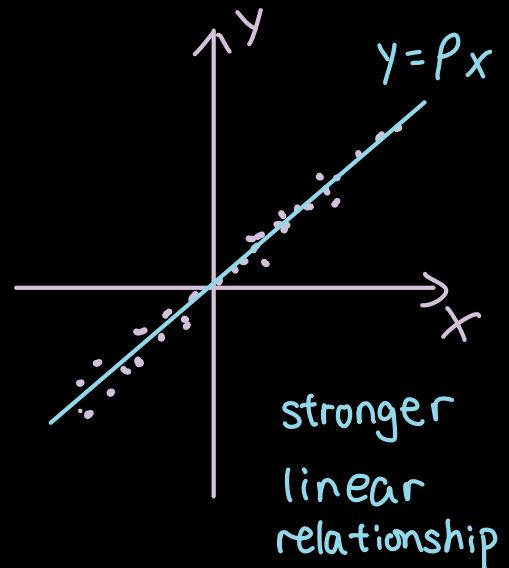
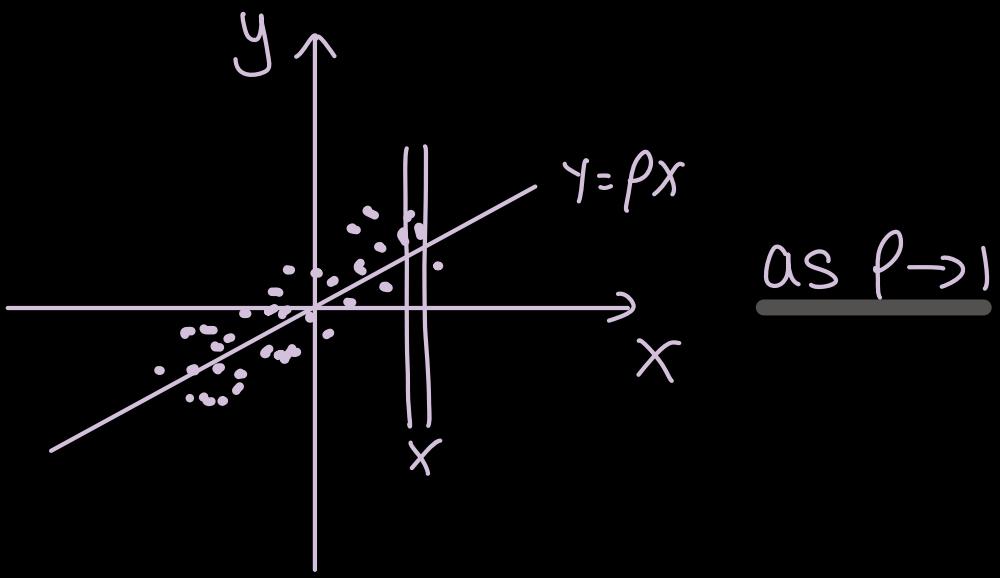
Bivariate Normal

$$X \sim N(0, 1) \quad f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

$$[Y \mid X=x] \sim N(\rho x, 1-\rho^2)$$
$$\downarrow \qquad \downarrow$$
$$\mu \qquad \sigma^2$$

$$f(y|x) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(y-\mu)^2}{2\sigma^2}}$$

$$= \frac{1}{\sqrt{2\pi(1-\rho^2)}} e^{-\frac{(y-\rho x)^2}{2(1-\rho^2)}}$$



$$f(x,y) = f(x) f(y|x)$$

$$= \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{1}{2} \left( x^2 + \frac{(y-\rho x)^2}{1-\rho^2} \right)}$$

$$\frac{x^2 - \cancel{\rho^2}x^2 + y^2 - 2\rho xy + \cancel{\rho^2}x^2}{1 - \rho^2}$$

$$= \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{x^2 + y^2 - 2\rho xy}{2(1-\rho^2)}}$$

Symmetry  $\rightarrow f(x|y) [x | Y=y] \sim N(\rho y, 1-\rho^2)$

# Using Gibbs Sampler

Initialize  $X_0, Y_0$

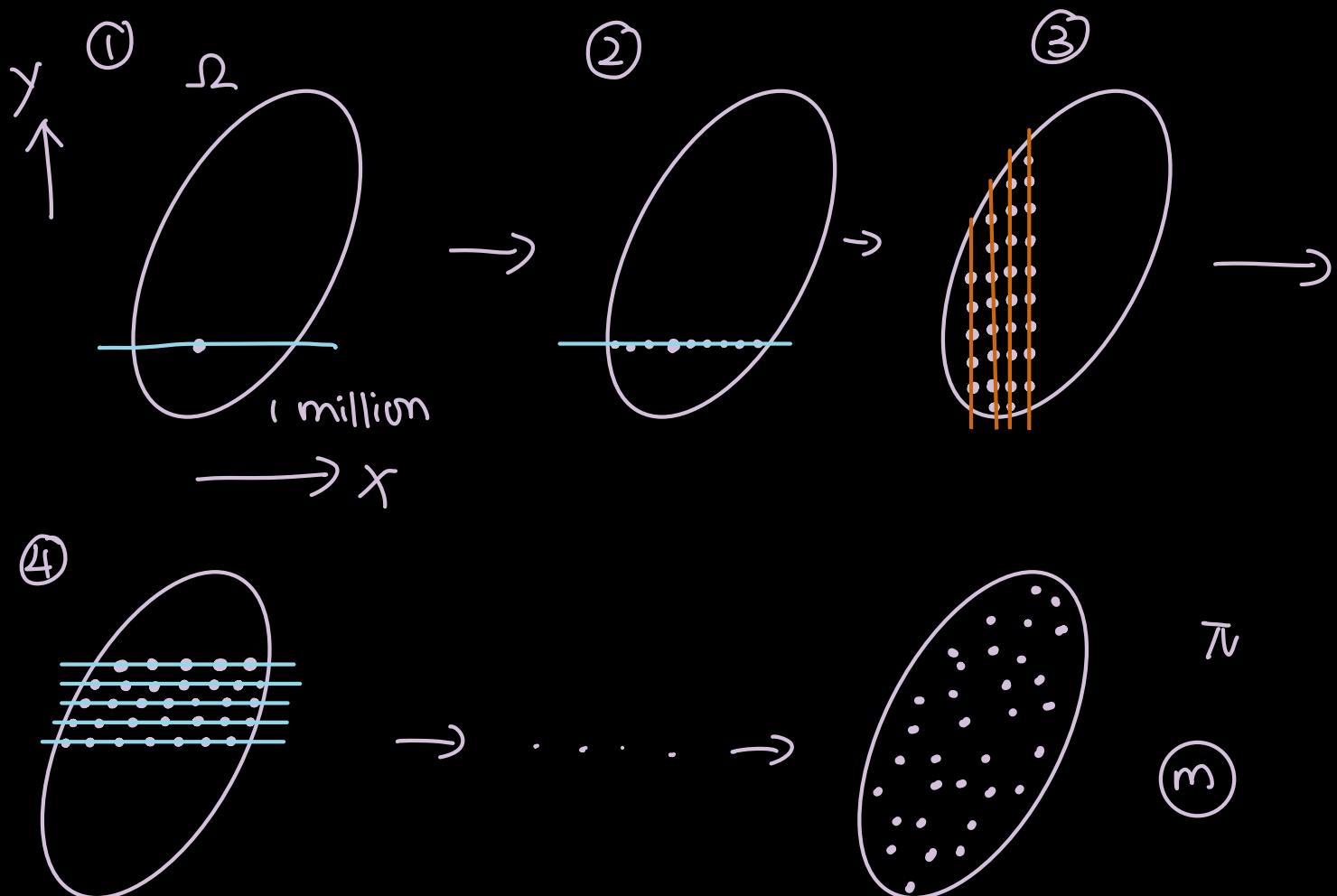
Iterate:

$$X_t \sim N(\rho Y_{t-1}, 1 - \rho^2) = \rho Y_{t-1} + \sqrt{1-\rho^2} N(0, 1)$$

$$Y_t \sim N(\rho X_t, 1 - \rho^2) = \rho X_t + \sqrt{1-\rho^2} N(0, 1)$$

## Intuitive Example

$$\pi(x, y) \sim \text{Unif}(\Omega)$$



# Monte-Carlo Integration

$$I = E_{\pi}(h(x))$$

MCMC  $\rightarrow$   $x_0 \sim \pi, x_1 \sim \pi, \dots, x_B \sim \pi, x_{B+1} \sim \dots, x_T \sim \pi$   
Burn-in period

$$\hat{I} = \frac{1}{T-B} \sum_{i=B+1}^T h(x_i)$$

$$E(\hat{I}) = I \quad \text{Var}(\hat{I})$$

# Bayesian Statistics

parameter  $\theta \sim P(\theta)$  prior distribution

$$[x|\theta] \sim P(x|\theta)$$

↓

data

$$\rightarrow \text{posterior } P(\theta|x) \xrightarrow{\text{conditioning}} \frac{P(\theta, x)}{P(x)} \xleftarrow[\text{marginalization}]{=} \underbrace{\frac{P(\theta)P(x|\theta)}{\int P(\theta)P(x|\theta) d\theta}}_{\text{factorization}}$$

$$= \frac{P(\theta)P(x|\theta)}{\int P(\theta)P(x|\theta) d\theta} \rightarrow z$$

Use MCMC  $\sim p(\theta | x)$

gives us plausible values of  $\theta$