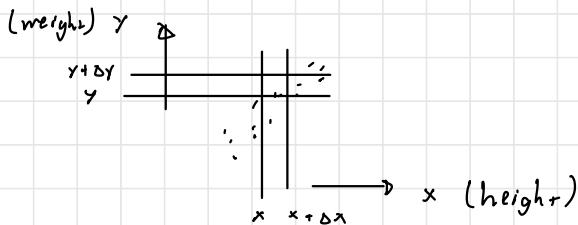


Lecture 11



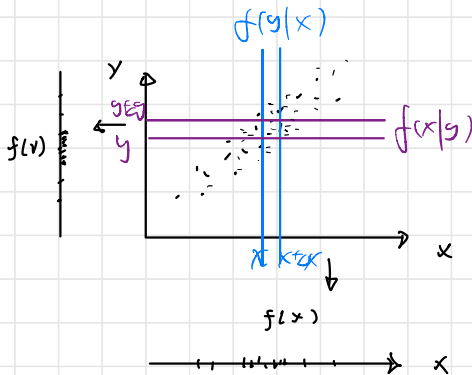
• Probability :

- $(x, y) \sim f(x, y)$
- imagine a population of $N (\rightarrow \infty)$
equally likely possibilities
- Population version scatterplot



• Counting :

- joint : $f(x, y)$
- marginal : $f(x), f(y)$
- conditional : $f(y|x), f(x|y)$

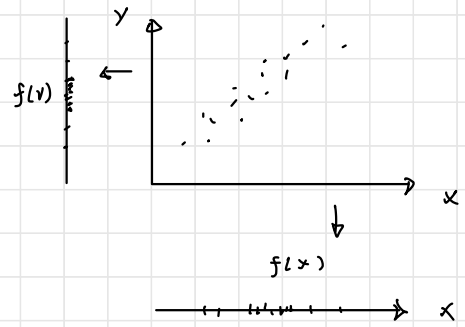


- 3 operations & 1 merge rule

(1) Marginalization:

$$f(x) = \int f(x, y) dy$$

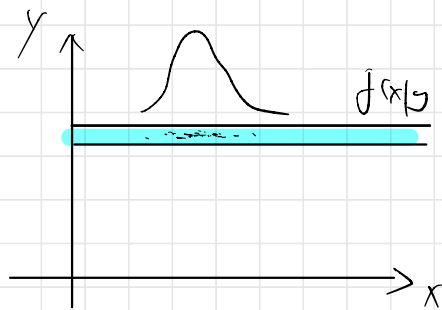
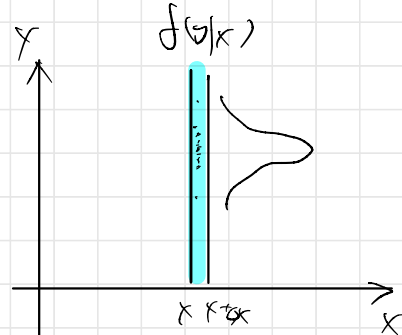
$$f(y) = \int f(x, y) dx$$



(2) Conditioning

$$f(y|x) = \frac{f(x, y)}{f(x)} = \frac{f(x, y)}{\underbrace{\int f(x, y) dy}_{\text{normalization}}}$$

$$f(x|y) = \frac{f(x, y)}{f(y)} = \frac{f(x, y)}{\int f(x, y) dx}$$



(3) Factorization:

$$f(x, y) = f(x) f(y|x) = f(y) f(x|y)$$

- Meta Rule: Insert the same condition
count the same subpopulation

- Example

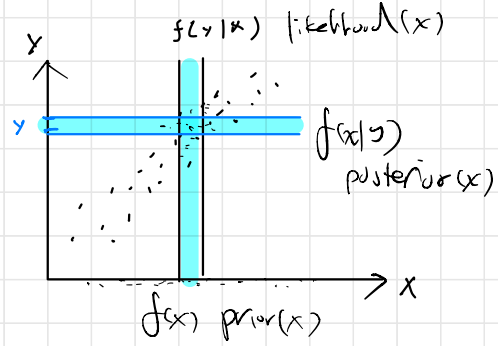
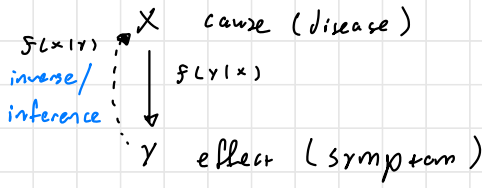
(1) $f(x|z) = \int f(x, y|z) dy$ marginalization

(2) $f(y|x, z) = \frac{f(x, y|z)}{f(x|z)}$ conditioning

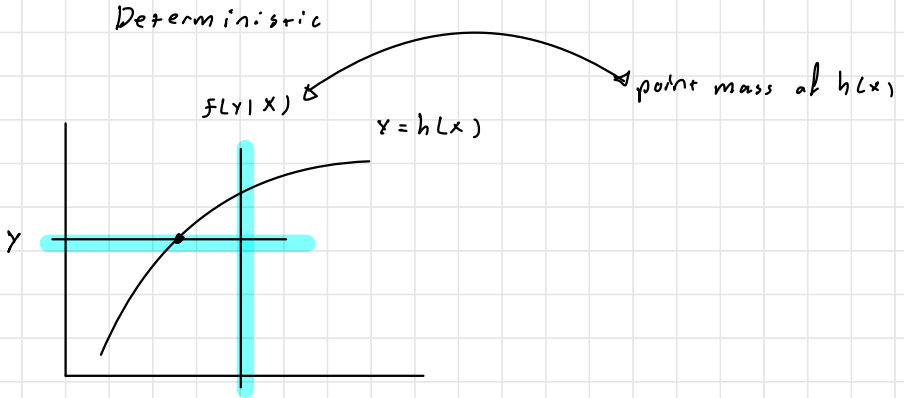
(3) $f(x, y|z) = f(x|z) f(y|x, z)$ factorization

- Discrete: $\int \mapsto \sum$

• Baye's Rule :



$$\begin{cases}
 x \sim f(x) & \text{prior}(x) \\
 [y|x] \sim f(y|x) & \text{likelihood}(x) \\
 \rightarrow [x|y] \sim f(x|y) & \text{posterior}(x)
 \end{cases}$$



Solving : $y = h(x)$
 Solution : $x = h^{-1}(y)$

• Bayes' Rule:

$$\begin{aligned} f(x|y) &= \frac{f(x,y)}{f(y)} && \text{conditioning} \\ &\quad \downarrow \\ &\quad \text{fixed} \\ &= \frac{f(x,y)}{\int f(x,y) dx} && \text{marginalization} \\ &= \frac{f(x) f(y|x)}{\int f(x) f(y|x) dx} && \text{factorization} \end{aligned}$$

• Note: $f(x|y) \propto f(x,y) = f(x) f(y|x)$

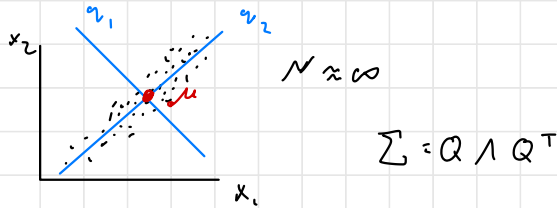
$\begin{array}{cccc} \uparrow & \uparrow & \downarrow & \downarrow \\ \text{fixed} & \text{fixed} & \text{prior} & \text{likelihood} \end{array}$

as a function of x

• Multivariate Gaussian:

$$X \sim \mathcal{N}(\mu, \Sigma)$$

$$f(x) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right)$$



$$\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \sim \mathcal{N}\left(0, \Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}\right)$$

Special case when $\Sigma_{21} = \Sigma_{12} = 0$.

Then,

$$f(x_1, x_2) = f(x_1) f(x_2), \text{ i.e. } x_1 \perp x_2$$

• In general, diagonalize

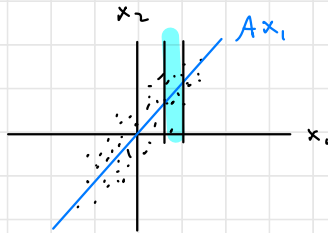
$$\begin{pmatrix} x_1 \\ z \end{pmatrix} = \begin{pmatrix} x_1 \\ x_2 - Ax_1 \end{pmatrix} = \begin{pmatrix} I & 0 \\ -A & I \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \sim \mathcal{N} \left(0, \begin{pmatrix} \Sigma_{11} & 0 \\ 0 & \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12} \end{pmatrix} \right)$$

$$A = \Sigma_{21} \Sigma_{11}^{-1}$$

$$x_2 = Ax_1 + \varepsilon, \quad \varepsilon \perp x_1$$

$$[x_2 | x_1] \sim \mathcal{N}(Ax_1, \Sigma_{22} - \dots)$$

↑
fixed



• Change Notation:

$$\begin{pmatrix} x \\ y \end{pmatrix} \sim \mathcal{N} \left(0, \begin{pmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{pmatrix} \right)$$

- population - version regression

| | | |
|---|---------|---------|
| 1 | | |
| ⋮ | | |
| i | x_i^T | y_i^T |
| ⋮ | | |
| N | X | Y |

$$\hat{\beta}_{LS} = (X^T X)^{-1} X^T Y$$

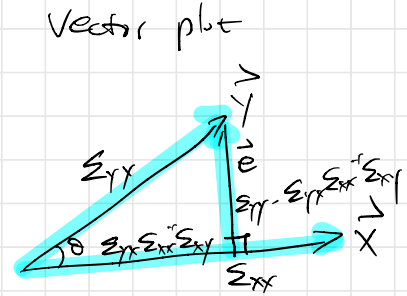
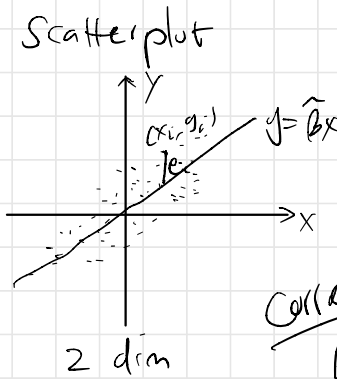
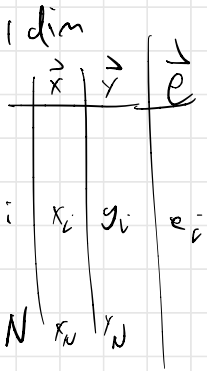
$$\frac{1}{N} X^T X = \begin{matrix} 1 & 2 & \dots & i & \dots & N \\ \boxed{x_i} & & & & & \end{matrix} \begin{matrix} \\ \\ \\ \\ \\ \end{matrix} x_i^T$$

$$= \frac{1}{N} \sum_{i=1}^N x_i x_i^T \stackrel{1 \text{ dim}}{=} \frac{1}{N} \sum x_i^2 = \frac{1}{N} |\bar{x}|^2$$

$$= \mathbb{E}[(x - \mu)(x - \mu)^T] = \text{Var}(x) = \sum_{xx}$$

Similarly $\frac{1}{N} X^T Y = \sum_{xy} \stackrel{1 \text{ dim}}{=} \frac{1}{N} \langle \bar{x}, \bar{y} \rangle$

So $\hat{\beta} = \sum_{xx}^{-1} \sum_{xy} = A^{-1}$



Correlation

$$\rho = \text{Corr}(x, y) = \frac{\text{Cov}(x, y)}{\sqrt{\text{Var}(x)} \sqrt{\text{Var}(y)}} = \text{Cov}\left(\frac{x - \mu_x}{\sigma_x}, \frac{y - \mu_y}{\sigma_y}\right)$$

$$= \frac{\frac{1}{N} \langle \vec{x}, \vec{y} \rangle}{\sqrt{\frac{1}{N} |\vec{x}|^2} \sqrt{\frac{1}{N} |\vec{y}|^2}} = \cos \theta$$

$$\frac{x - \mu_x}{\sigma_x} \rightarrow \begin{pmatrix} x \\ y \end{pmatrix} \sim N\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}\right)$$

$$\frac{y - \mu_y}{\sigma_y} \rightarrow$$

$$[y|x] \sim N(\rho x, 1 - \rho^2)$$

↓ ↓
Son height father height

< 1

Regression

$$1 - \rho^2 = \text{sm}^2 = \frac{|\vec{e}|^2}{|\vec{y}|^2} = \frac{\text{Var}(e)}{\text{Var}(y)}$$

Regression towards the mean
original meaning of "regression"

• Bayesian Regression:

| Data | p | |
|------|-----|-----|
| | X | Y |
| n | | |

$$[Y | X, \beta] \sim \mathcal{N}(X\beta, \sigma^2 I_n)$$

$p(\beta)$
 prior: $\beta \sim \mathcal{N}(0, \sigma^2 I_p)$
 (as if β is sampled from population of N possibilities)

posterior:

$$p(\beta | X, Y) \propto \underbrace{p(\beta)}_{\text{prior}(\beta)} \underbrace{p(Y | X, \beta)}_{\text{likelihood}(\beta)}$$

Note: Gauss assumed $\sigma^2 = \infty$,
 so we know nothing about β
 (non-informative prior).

• What about finite σ^2

$$p(\beta | X, Y) \propto \frac{1}{(2\pi)^{\frac{p}{2}} (\sigma^2)^{\frac{p}{2}}} \exp\left(-\frac{1}{2\sigma^2} \beta^T \beta\right)$$

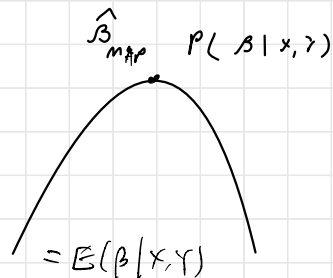
$$\underbrace{[Y | X, \beta]}_{\text{density}} \rightarrow \frac{1}{(2\pi)^{\frac{n}{2}} (\sigma^2)^{\frac{n}{2}}} \exp\left(-\frac{1}{2\sigma^2} (Y - X\beta)^T (Y - X\beta)\right)$$

$$\propto \exp\left(-\frac{1}{2} \left(\frac{1}{\sigma^2} |Y - X\beta|^2 + \frac{1}{\sigma^2} |\beta|^2 \right)\right)$$

Ridge - Regression

• Maximum A Posterior: (MAP)

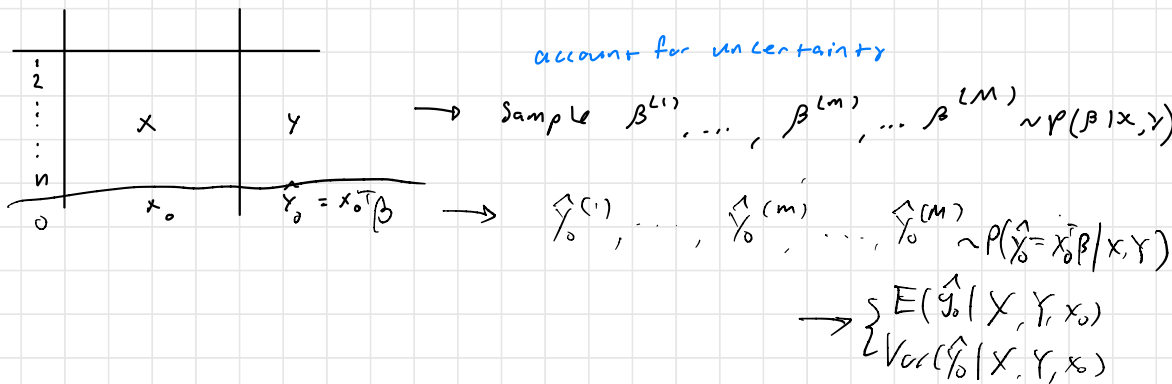
$$\begin{aligned} & \min_{\beta} \frac{1}{\sigma^2} |Y - X\beta|^2 + \frac{1}{\sigma^2} |\beta|^2 \\ & = \min_{\beta} |Y - X\beta|^2 + \frac{\sigma^2}{\sigma^2} |\beta|^2 \\ & \quad \uparrow \lambda > 0 \end{aligned}$$



- $[B | x, y] \sim p(B | x, y)$

↓
basis for inference / prediction

↓
Does not overfit



population of β → population of $\begin{pmatrix} y \\ \hat{y}_0 \end{pmatrix} \sim N(\theta, \quad)$ can be derived directly

- Bayesian / Frequentist controversy

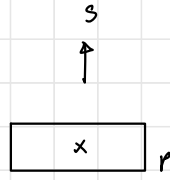
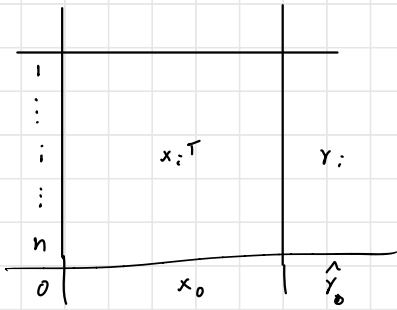
- example: prior, speed of light $c \sim \text{prior}(c)$

- $c \sim p(c)$

- feasible of repeated sampling?

$c_1, c_2, \dots, c_m \stackrel{iid}{\sim} p(c)$

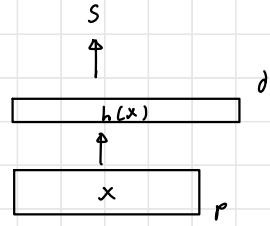
↓ frequency → prob



$$s = f(x) = x^T \beta$$

$$\beta \sim \mathcal{N}(0, \sigma^2 I_p)$$

(Ridge)



$$s = f(x) = h(x)^T \beta$$

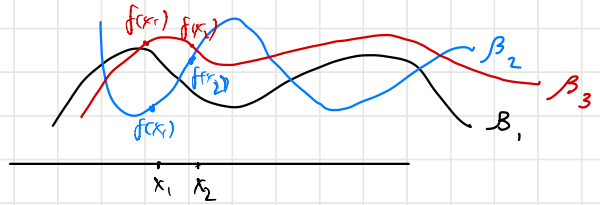
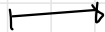
$$\beta \sim \mathcal{N}(0, \sigma^2 I_d)$$

(kernel)

• For simplicity assume x is 1d

population of $\beta \rightarrow$ population of $f(x) = h(x)^T \beta$

β pop



So we say $f(x) = h(x)^T \beta \sim$ Gaussian Process

$$\text{So } \text{Cov}(f(x_1), f(x_2)) = \text{Cov}(h(x_1)^T \beta, h(x_2)^T \beta)$$

Remember $\mu = 0$

$$\begin{aligned} &= \mathbb{E}(h(x_1)^T \beta \beta^T h(x_2)) \quad , \beta \text{ is random} \\ &\quad \quad \quad x_1 \neq x_2 \text{ is fixed} \\ &= h(x_1)^T \mathbb{E}(\beta \beta^T) h(x_2) \\ &= h(x_1)^T \sigma^2 I_d h(x_2) \\ &= \sigma^2 \langle h(x_1), h(x_2) \rangle \\ &= \sigma^2 k(x_1, x_2) \end{aligned}$$

$$\text{So } f(x) \sim \text{GP}(0, \sigma^2 k)$$

| | | | |
|----------|---------|----------|-----------------------------|
| 1 | | $f(x_1)$ | y_1 |
| \vdots | | | |
| i | x_i^T | $f(x_i)$ | $y_i = f(x_i) + \epsilon_i$ |
| \vdots | | | |
| n | | $f(x_n)$ | y_n |
| 0 | x_0 | $f(x_0)$ | |

$$\epsilon = \begin{pmatrix} \epsilon_1 \\ \vdots \\ \epsilon_i \\ \vdots \\ \epsilon_n \end{pmatrix} \sim N(0, \sigma^2 I)$$

$$\begin{pmatrix} f(x_1) \\ \vdots \\ f(x_i) \\ \vdots \\ f(x_n) \\ f(x_0) \end{pmatrix} \sim N \left(0, \begin{array}{c|c} & \begin{matrix} j & 0 \end{matrix} \\ \hline \begin{matrix} i & \end{matrix} & \begin{matrix} \sigma^2 k_{ij} & \sigma^2 k_{i0} \\ \sigma^2 k_{0j} & \sigma^2 k_{00} \end{matrix} \end{array} \right)$$

$$k_{ij} = k(x_i, x_j)$$

$$\begin{pmatrix} y_1 \\ \vdots \\ y_i \\ \vdots \\ y_n \\ \hat{y}_0 = f(x_0) \end{pmatrix} \sim \mathcal{N} \left(0, \begin{array}{c|c} & \sigma^2 k_0^T \\ \hline \sigma^2 K + \sigma^2 I_n & \\ \hline \sigma^2 k_0 & \sigma^2 k_0 \end{array} \right)$$

$$S_0 [\hat{y}_0 = f(x_0) | y, x] \sim \mathcal{N}(\sigma^2 k_0 (\sigma^2 K + \sigma^2 I_n)^{-1} y,$$

$$\sigma^2 k_0 - \sigma^2 k_0 (\sigma^2 K + \sigma^2 I_n)^{-1} \sigma^2 k_0^T)$$

Kernel Regression

$$k_0 (K + \frac{\sigma^2}{\tau^2} I_n)^{-1} y$$

$$= k_0 (K + \lambda I_n)^{-1} y$$

$$= k_0 c$$

$$= \begin{array}{|c|} \hline k_0 \\ \hline \end{array} \begin{array}{|c|} \hline c \\ \hline \end{array}$$

$$f(x_0) = \sum_{i=1}^n c_i k(x_i, x_0)$$